# **Independence and Counting**

Berlin Chen Department of Computer Science & Information Engineering National Taiwan Normal University

Reference:

- D. P. Bertsekas, J. N. Tsitsiklis, Introduction to Probability, Sections 1.5-1.6

# Independence (1/2)

- Recall that conditional probability  $\mathbf{P}(A|B)$  captures the partial information that event *B* provides about event *A*
- A special case arises when the occurrence of *B* provides no such information and does not alter the probability that *A* has occurred

$$\mathbf{P}(A|B) = \mathbf{P}(A)$$

- *A* is independent of *B* (*B* also is independent of *A*)  $\Rightarrow \mathbf{P}(A|B) = \frac{\mathbf{P}(A \cap B)}{\mathbf{P}(B)} = \mathbf{P}(A)$   $\Rightarrow \mathbf{P}(A \cap B) = \mathbf{P}(A)\mathbf{P}(B)$ 

# Independence (2/2)

- *A* and *B* are independent => *A* and *B* are disjoint (?)
  - No ! Why ?
    - A and B are disjoint then  $P(A \cap B) = 0$
    - However, if P(A) > 0 and P(B) > 0

 $\Rightarrow \mathbf{P}(A \cap B) \neq \mathbf{P}(A)\mathbf{P}(B)$ 



• Two disjoint events A and B with P(A) > 0 and P(B) > 0are never independent

# Independence: An Example (1/3)

• **Example 1.19.** Consider an experiment involving two successive rolls of a 4-sided die in which all 16 possible outcomes are equally likely and have probability 1/16

(a) Are the events,

 $A_i = \{1 \text{ st roll results in } i\},\$  $B = \{2 \text{ roll results in } i\}$  independent

 $B_j$  = {2nd roll results in *j* }, independent?



$$\mathbf{P}(A_i \cap B_j) = \frac{1}{16}$$

$$\mathbf{P}(A_i) = \frac{4}{16}, \ \mathbf{P}(B_j) = \frac{4}{16}$$

$$\Rightarrow \mathbf{P}(A_i \cap B_j) = \mathbf{P}(A_i)\mathbf{P}(B_j)$$

$$\Rightarrow A_i \text{ and } B_j \text{ are independent } !$$

Sample Space Pair of Rolls

Probability Law here

**Using Discrete Uniform** 

# Independence: An Example (2/3)

(b) Are the events,

A= {1st roll is a 1},

B= {sum of the two rolls is a 5}, independent?

 $\mathbf{P}(A) = \frac{4}{16} \quad (\text{ the results of two rolls are } (1,1),(1,2),(1,3),(1,4)))$   $\mathbf{P}(B) = \frac{4}{16} \quad (\text{ the results of two rolls are } (1,4),(2,3),(3,2),(4,1)))$   $\mathbf{P}(A \cap B) = \frac{1}{16} \quad (\text{ the only one result of two rolls is } (1,4))$   $\Rightarrow \mathbf{P}(A \cap B) = \mathbf{P}(A)\mathbf{P}(B)$   $\Rightarrow A \text{ and } B \text{ are independent } !$ 

# Independence: An Example (3/3)

(c) Are the events,

A= {maximum of the two rolls is 2},

*B*= {minimum of the two rolls is 2}, independent?

$$\mathbf{P}(A) = \frac{3}{16} \quad (\text{ the results of two rolls are } (1,2),(2,1),(2,2))$$

$$\mathbf{P}(B) = \frac{5}{16} \quad (\text{ the results of two rolls are } (2,2),(2,3),(2,4),(3,2),(4,2))$$

$$\mathbf{P}(A \cap B) = \frac{1}{16} \quad (\text{ the only one result of two rolls is } (2,2))$$

$$\Rightarrow \mathbf{P}(A \cap B) \neq \mathbf{P}(A)\mathbf{P}(B)$$

$$\Rightarrow A \text{ and } B \text{ are dependent!}$$

### Conditional Independence (1/2)

• Given an event *C*, the events *A* and *B* are called conditionally independent if

$$\mathbf{P}(A \cap B|C) = \mathbf{P}(A|C)\mathbf{P}(B|C)$$
<sup>1</sup>

- We also know that

$$\mathbf{P}(A \cap B | C) = \frac{\mathbf{P}(A \cap B \cap C)}{\mathbf{P}(C)} \qquad \text{multiplication rule}$$
$$= \frac{\mathbf{P}(C)\mathbf{P}(B | C)\mathbf{P}(A | B \cap C)}{\mathbf{P}(C)}^{2}$$

- If P(B|C) > 0, we have an alternative way to express conditional independence

$$\mathbf{P}(A|B\cap C) = \mathbf{P}(A|C)^{\mathbf{3}}$$

# Conditional Independence (2/2)

• Notice that independence of two events *A* and *B* with respect to the unconditionally probability law does not imply conditional independence, and vice versa

 $\mathbf{P}(A \cap B) = \mathbf{P}(A)\mathbf{P}(B) \quad \bigstar \quad \mathbf{P}(A \cap B|C) = \mathbf{P}(A|C)\mathbf{P}(B|C)$ 

- If A and B are independent, the same holds for
  (i) A and B<sup>c</sup>
  (ii) A<sup>c</sup> and B<sup>c</sup>
  - How can we verify it ? (See Problem 43)

Conditional Independence: Examples (1/2)

 Example 1.20. Consider two independent fair coin tosses, in which all four possible outcomes are equally likely. Let
 Using Discrete Uniform

Probability Law here

 $H_1 = \{1 \text{ st toss is a head}\}, (H,T), (H,H)$  $H_2 = \{2 \text{ nd toss is a head}\}, (T,H), (H,H)$  $D = \{\text{the two tosses have different results}\}. (T,H), (H,T)$ 

$$\mathbf{P}(H_1|D) = \frac{1}{2} \qquad (H,T)$$

$$\mathbf{P}(H_2|D) = \frac{1}{2} \qquad (T,H)$$

$$\mathbf{P}(H_1 \cap H_2|D) = \frac{\mathbf{P}(H_1 \cap H_2 \cap D)}{\mathbf{P}(D)} = 0 \neq \mathbf{P}(H_1|D)\mathbf{P}(H_2|D)$$

$$\Rightarrow H_1 \text{ and } H_2 \text{ are conditionally dependent !}$$

# Conditional Independence: Examples (2/2)

- **Example 1.21.** There are two coins, a blue and a red one
  - We choose one of the two at random, each being chosen with probability 1/2, and proceed with two independent tosses
  - The coins are biased: with the blue coin, the probability of heads in any given toss is 0.99, whereas for the red coin it is 0.01
  - Let *B* be the event that the blue coin was selected. Let also  $H_i$  be the event that the *i*-th toss resulted in heads al case:  $\mathbf{P}(H \cap H \mid R) - \mathbf{P}(H \mid R) \mathbf{P}(H \mid R)$  Given the choice of a coin, the

conditional case:

$$\mathbf{P}(H_{1} \cap H_{2}|B) = \mathbf{P}(H_{1}|B) + (H_{2}|B) = \mathbf{P}(H_{1}|B) + (H_{2}|B) \quad \text{events } H_{1} \text{ and } H_{2} \text{ are independent}$$
  
unconditional case: 
$$\mathbf{P}(H_{1} \cap H_{2}) \stackrel{?}{=} \mathbf{P}(H_{1})\mathbf{P}(H_{2})$$
  

$$\mathbf{P}(H_{1}) = \mathbf{P}(B)\mathbf{P}(H_{1}|B) + \mathbf{P}(B^{C})\mathbf{P}(H_{1}|B^{C}) = \frac{1}{2} \cdot 0.99 + \frac{1}{2} \cdot 0.01 = \frac{1}{2}$$
  

$$\mathbf{P}(H_{2}) = \mathbf{P}(B)\mathbf{P}(H_{2}|B) + \mathbf{P}(B^{C})\mathbf{P}(H_{2}|B^{C}) = \frac{1}{2} \cdot 0.99 + \frac{1}{2} \cdot 0.01 = \frac{1}{2}$$
  

$$\mathbf{P}(H_{1} \cap H_{2}) = \mathbf{P}(B)\mathbf{P}(H_{1} \cap H_{2}|B) + \mathbf{P}(B^{C})\mathbf{P}(H_{1} \cap H_{2}|B^{C})$$
  

$$= \frac{1}{2} \cdot 0.99 \cdot 0.99 + \frac{1}{2} \cdot 0.01 \cdot 0.01 \neq \frac{1}{4}$$
  
Probability-Berlin Chen 10

#### Independence of a Collection of Events

• We say that the events  $A_1, A_2, \dots, A_n$  are **independent** if

$$\mathbf{P}\left(\bigcap_{i\in S}A_i\right) = \prod_{i\in S}\mathbf{P}(A_i), \text{ for every subset } S \text{ of } \{1,2,\ldots,n\}$$

• For example, the independence of three events  $A_1, A_2, A_3$  amounts to satisfying the four conditions

$$\mathbf{P}(A_{1} \cap A_{2}) = \mathbf{P}(A_{1})\mathbf{P}(A_{2})$$
  

$$\mathbf{P}(A_{1} \cap A_{3}) = \mathbf{P}(A_{1})\mathbf{P}(A_{3})$$
  

$$\mathbf{P}(A_{2} \cap A_{3}) = \mathbf{P}(A_{2})\mathbf{P}(A_{3})$$
  

$$\mathbf{P}(A_{1} \cap A_{2} \cap A_{3}) = \mathbf{P}(A_{1})\mathbf{P}(A_{2})\mathbf{P}(A_{3})$$

Independence of a Collection of Events: Examples (1/4)

- Example 1.22. Pairwise independence does not imply independence.
  - Consider two independent fair coin tosses, and the following events:

 $H_1 = \{ \text{ 1st toss is a head } \}, (H, T), (H, H) \}$ 

 $H_2 = \{ \text{ 2nd toss is a head } \}, (T, H), (H, H) \}$ 

 $D = \{$  the two tosses have different results  $\}$ . (T, H), (H, T)

 $\mathbf{P}(H_1 \cap H_2) = \mathbf{P}(H_1)\mathbf{P}(H_2)$   $\mathbf{P}(H_1 \cap D) = \mathbf{P}(H_1)\mathbf{P}(D)$   $\mathbf{P}(H_2 \cap D) = \mathbf{P}(H_2)\mathbf{P}(D)$ However,  $\mathbf{P}(H_1 \cap H_2 \cap D) = 0 \neq \mathbf{P}(H_1)\mathbf{P}(H_2)\mathbf{P}(D)$  Independence of a Collection of Events: Examples (2/4)

• Example 1.23. The equality

$$\mathbf{P}(A_1 \cap A_2 \cap A_3) = \mathbf{P}(A_1)\mathbf{P}(A_2)\mathbf{P}(A_3)$$

#### is not enough for independence.

Consider two independent rolls of a fair six-sided die, and the following events:

 $A = \{ \text{ 1st roll is } 1, 2, \text{ or } 3 \},\$   $B = \{ \text{ 1st roll is } 3, 4, \text{ or } 5 \},\$   $C = \{ \text{ the sum of the two rolls is } 9 \}.$   $P(A \cap B \cap C) = \frac{1}{36} = \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{4}{36} = P(A)P(B)P(C)$ However,  $P(A \cap B) = \frac{1}{6} \neq \frac{1}{2} \cdot \frac{1}{2} = P(A)P(B)$   $P(A \cap C) = \frac{1}{36} \neq \frac{1}{2} \cdot \frac{4}{36} = P(A)P(C)$   $P(B \cap C) = \frac{1}{12} \neq \frac{1}{2} \cdot \frac{4}{36} = P(B)P(C)$ 

Independence of a Collection of Events: Examples (3/4)

- Example 1.24. Network connectivity. A computer network connects two nodes A and B through intermediate nodes C, D, E, F (See next slide)
  - For every pair of directly connected nodes, say *i* and *j*, there is a given probability p<sub>ij</sub> that the link from *i* to *j* is up. We assume that link failures are independent of each other
  - What is the probability that there is a path connecting A and B in which all links are up?

**P**(series subsystem succeeds) =  $p_1 p_2 \cdots p_n$ 



 $\mathbf{P}(\text{parallel subsystem succeeds}) = 1 - \mathbf{P}(\text{parallel subsystem fails}) = 1 - (1 - p_i)(1 - p_2) \cdots (1 - p_n)$ 

Parallel Connection

Independence of a Collection of Events: Examples (4/4)

• Example 1.24. (cont.)



# Recall: Counting in Probability Calculation

- Two applications of the discrete uniform probability law
  - When the sample space  $\Omega$  has a finite number of equally likely outcomes, the probability of any event A is given by

 $\mathbf{P}(A) = \frac{\text{number of elements of } A}{\text{number of elements of } \Omega}$ 

– When we want to calculate the probability of an event A with a finite number of equally likely outcomes, each of which has an already known probability p. Then the probability of A is given by

 $\mathbf{P}(A) = p \cdot (\text{number of elements of } A)$ 

• E.g., the calculation of *k* heads in *n* coin tosses

# The Counting Principle

• Consider a process that consists of *r* stages. Suppose that:

(a) There are  $n_1$  possible results for the first stage

- (b) For every possible result of the first stage, there are  $n_2$  possible results at the second stage
- (c) More generally, for all possible results of the first *i* -1 stages, there are  $n_i$  possible results at the *i*-th stage
- Then, the total number of possible results of the *r*-stage process is



# Common Types of Counting

• Permutations of *n* objects

$$n!=n\cdot(n-1)\cdot(n-2)\cdots 2\cdot 1$$

• *k*-permutations of *n* objects

$$\frac{n!}{(n-k)!}$$

• Combinations of *k* out of *n* objects

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

 Partitions of *n* objects into *r* groups with the *i*-th group having n<sub>i</sub> objects

$$\binom{n}{n_1, n_2, \dots, n_r} = \frac{n!}{n_1! n_2! \cdots n_r!}$$

# Summary of Chapter 1 (1/2)

- A probability problem can usually be broken down into a few basic steps:
  - 1. The description of the sample space, i.e., the set of possible outcomes of a given experiment
  - 2. The (possibly indirect) specification of the probability law (the probability of each event)
  - 3. The calculation of probabilities and conditional probabilities of various events of interest

# Summary of Chapter 1 (2/2)

- Three common methods for calculating probabilities
  - The counting method: if the number of outcome is finite and all outcome are equally likely

$$\mathbf{P}(A) = \frac{\text{number of elements of A}}{\text{number of elements of }\Omega}$$

- The sequential method: the use of the multiplication (chain) rule

$$\mathbf{P}\left(\bigcap_{i=1}^{n}A_{i}\right) = \mathbf{P}\left(A_{1}\right)\mathbf{P}\left(A_{2}|A_{1}\right)\mathbf{P}\left(A_{3}|A_{1}\cap A_{2}\right)\cdots\mathbf{P}\left(A_{n}|\bigcap_{i=1}^{n-1}A_{i}\right)$$

 The divide-and-conquer method: the probability of an event is obtained based on a set of conditional probabilities

$$\mathbf{P}(B) = \mathbf{P}(A_1 \cap B) + \dots + \mathbf{P}(A_n \cap B)$$
$$= \mathbf{P}(A_1)\mathbf{P}(B|A_1) + \dots + \mathbf{P}(A_n)\mathbf{P}(B|A_n)$$

•  $A_1, \dots, A_n$  are disjoint events that form a partition of the sample space

# Recitation

- SECTION 1.5 Independence
  - Problems 37, 38, 39, 40, 42