# Further Topics on Random Variables: Conditional Expectation and Variance Revisited 

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## Revisit: Conditional Expectation and Variance

- Goal: To introduce two useful probability laws
- Law of Iterated Expectations

$$
\mathbf{E}[\mathbf{E}[X \mid Y]]=\mathbf{E}[X]
$$

- Law of Total Variance

$$
\operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])
$$

## More on Conditional Expectation

- Recall that the conditional expectation $\mathbf{E}[x \mid Y=y]$ is defined by

$$
\mathbf{E}[X \mid Y=y]=\sum_{x} x \cdot p_{X \mid Y}(x \mid y), \quad(\text { If } X \text { is discrete })
$$

and

$$
\mathbf{E}[X \mid Y=y]=\int_{-\infty}^{\infty} x \cdot f_{X \mid Y}(x \mid y) d x . \text { (If } X \text { is continuous) }
$$

- $\mathbf{E}[X \mid Y=y]$ in fact can be viewed as a function of $Y$, because its value depends on the value $y$ of $Y$
- Is $\mathbf{E}[X \mid Y]$ a random variable?
- What is the expected value of $\mathbf{E}[X \mid Y]$ ?
- Note also that the expectation of a function $g(Y)$ of $Y$

$$
\mathbf{E}[g(Y)]= \begin{cases}\sum_{y} g(y) p_{Y}(y), & \text { if } Y \text { is discrete } \\ \int_{-\infty}^{\infty} g(y) f_{Y}(y) d y, & \text { if } Y \text { is continuous }\end{cases}
$$

## An Illustrative Example (1/2)

- Example. Let the random variables $X$ and $Y$ have a joint PDF which is equal to 2 for $(x, y)$ belonging to the triangle indicated below and zero evervwhere else.

- What's the value of $\mathbf{E}[X \mid Y=y]$ ?

$$
\begin{array}{rlrl}
f_{Y}(y) & =\int_{-\infty}^{\infty} f_{X, Y}(x, y) d x & & f_{X \mid Y}(x \mid y)=\frac{f_{X, Y}(x, y)}{f_{Y}(y)}=\frac{1}{1-y} \\
& =\int_{0}^{1-y} f_{X, Y}(x, y) d x(\because X+Y \leq 1) & & \text { for } 0 \leq x \leq 1-y \text { where } 0 \leq y \leq 1 \\
& =\int_{0}^{1-y} 2 d x=2(1-y), \quad 0 \leq y \leq 1 & & \therefore \mathbf{E}[X \mid Y=y]=\int_{0}^{1-y} x \cdot \frac{1}{1-y} d x \\
& =\left.\frac{1}{2(1-y)} \cdot x^{2}\right|_{0} ^{1-y}=\frac{1-y}{2}
\end{array}
$$

## An Illustrative Example (2/2)

- We saw that $\quad \mathbf{E}[X \mid Y=y]=(1-y) / 2$. Hence, $\mathbf{E}[X \mid Y]$ is the random variable $(1-Y) / 2$

$$
\mathbf{E}[X \mid Y]=\frac{(1-Y)}{2}
$$

- The expectation of $\mathbf{E}[X \mid Y]$

$$
\mathbf{E}[\mathbf{E}[X \mid Y]]=\int_{-\infty}^{\infty} \mathbf{E}[X \mid Y=y] \underbrace{f_{Y}(y) d y=\mathbf{E}[X]}_{\text {Total Expectation Theorem }}
$$

For this problem, we thus have

$$
\begin{aligned}
\mathbf{E}[X] & =\mathbf{E}[\mathbf{E}[X \mid Y]=\mathbf{E}[(1-Y) / 2]=(1-\mathbf{E}[Y]) / 2 \\
\mathbf{E}[Y] & =\int_{0}^{1} y \cdot f_{Y}(y) d y \quad \therefore \mathbf{E}[X]=(1-\mathbf{E}[Y]) / 2=1 / 3 \\
& =\int_{0}^{1} y \cdot 2(1-y) d y \\
& =y^{2}-\left.(2 / 3) y^{3}\right|_{0} ^{1} \\
& =1 / 3
\end{aligned}
$$

## Law of Iterated Expectations

$$
\begin{gathered}
\mathbf{E}[\mathbf{E}[X \mid Y]=\mathbf{E}[X] \\
\mathbf{E}\left[\mathbf{E}[X \mid Y]= \begin{cases}\sum_{y} \mathbf{E}[X \mid Y=y] p_{Y}(y), & \text { (If } Y \text { is discrete) } \\
\int_{-\infty}^{\infty} \mathbf{E}[X \mid Y=y] f_{Y}(y) d y . & \text { (If } Y \text { is continuous) }\end{cases} \right.
\end{gathered}
$$

## An Illustrative Example (1/2)

- Example 4.17. We start with a stick of length $l$. We break it at a point which is chosen randomly and uniformly over its length, and keep the piece that contains the left end of the stick. We then repeat the same process on the stick that we were left with.
- What is the expected length of the stick that we are left with, after breaking twice?

Let $Y$ be the length of the stick after we break for the first time.
Let $X$ be the length after the second time.

$$
f_{Y}(y)=\left\{\begin{array}{ll}
\frac{1}{l}, & \text { for } 0 \leq y \leq l \\
0, & \text { otherwise }
\end{array} \text { and } f_{X \mid Y}(x \mid y)= \begin{cases}\frac{1}{y}, & \text { for } 0 \leq x \leq y \\
0, & \text { otherwise }\end{cases}\right.
$$

uniformly distributed
uniformly distributed

## An Illustrative Example (2/2)

- By the Law of Iterated Expectations, we have

$$
\begin{aligned}
\mathbf{E}[X] & =\mathbf{E}[\mathbf{E}[X \mid Y]] \\
& =\int_{0}^{l} \mathbf{E}[X \mid Y=y] f_{y}(y) d y=\int_{0}^{l}\left[\int_{0}^{y} x f_{X \mid Y}(x \mid y) d x\right] f_{y}(y) d y \\
& =\int_{0}^{l}\left[\int_{0}^{y} x \frac{1}{y} d x\right] \frac{1}{l} d y=\int_{0}^{l}\left[\left.\frac{1}{y} \cdot \frac{x^{2}}{2} \right\rvert\, \frac{y}{2}\right] \frac{1}{l} d y \\
& =\frac{1}{l} \cdot \int_{0}^{l} \frac{y}{2} d y=\left.\frac{1}{l} \cdot \frac{y^{2}}{4}\right|_{0} ^{l} \quad \text { Note that } \mathbf{E}[X \mid Y=y]=\frac{y}{2} \\
& =\frac{l}{4} \quad \begin{array}{l:l:l}
l
\end{array}
\end{aligned}
$$

## Averaging by Section (1/3)

- Averaging by section can be viewed as a special case of the law of iterated expectations
- Example 4.18. Averaging Quiz Scores by Section.
- A class has $n$ students and the quiz score of student $i$ is $x_{i}$. The average quiz score is

$$
m=\frac{1}{n} \sum_{i=1}^{n} x_{i}
$$

- If students are divided into $k$ disjoint subsets $A_{1}, A_{2, \ldots}, A_{k}$, the average score in section $s$ is

$$
m_{s}=\frac{1}{n_{s}} \sum_{x_{i} \in A_{s}} x_{i}
$$

## Averaging by Section (2/3)

- Example 4.18. (cont.)
- The average score of over the whole class can be computed by taking a weighted average of the average score $m_{s}$ of each class $s$, while the weight given to section $s$ is proportional to the number of students in that section

$$
\begin{aligned}
\sum_{s=1}^{k} \frac{n_{s}}{n} m_{s} & =\sum_{s=1}^{k} \frac{n_{s}}{n} \cdot \frac{1}{n_{S}} \sum_{x_{i} \in A_{s}} x_{i} \\
& =\frac{1}{n} \sum_{s=1}^{k} \sum_{x_{i} \in A_{s}}^{k} x_{i} \\
& =\frac{1}{n} \sum_{i=1}^{n} x_{i} \\
& =m
\end{aligned}
$$

## Averaging by Section (3/3)

- Example 4.18. (cont.)
- Its relationship with the law of iterated expectations
- Two random variable defined
- $X$ : quiz score of a student (or outcome)
» Each student (or outcome) is uniformly distributed
- $Y:$ section of a student $Y \in\{1, \ldots, k\}$
$\Rightarrow \mathbf{E}[X]=m$

$$
\begin{equation*}
\mathbf{E}[X \mid Y=s]=\frac{1}{n_{s}} \sum_{i \in A_{s}} x_{i}=m_{s} \tag{?}
\end{equation*}
$$

$\because P(Y=s)=\frac{n_{s}}{n}$

$$
\begin{align*}
\therefore m & =\mathbf{E}[X]=\mathbf{E}[\mathbf{E}[X \mid Y]]=\sum_{s=1}^{k} \mathbf{E}[X \mid Y=s] P(Y=s)  \tag{?}\\
& =\sum_{s=1}^{k} m_{s} \cdot \frac{n_{s}}{n}
\end{align*}
$$

## More on Conditional Variance

- Recall that the conditional variance of $X$, given $Y=y$, is defined by

$$
\operatorname{var}(X \mid Y=y)=\mathbf{E}\left[(X-\mathbf{E}[X \mid Y=y])^{2} \mid Y=y\right]
$$

- $\operatorname{var}(X \mid Y)$ in fact can be viewed as a function of $Y$, because its value $\operatorname{var}(X \mid Y=y)$ depends on the value $y$ of $Y$
- Is $\operatorname{var}(X \mid Y)$ a random variable?
- What is the expected value of $\operatorname{var}(X \mid Y)$ ?

Note that $\mathbf{E}[\operatorname{var}(X \mid Y)] \neq \operatorname{var}(X)$

## Law of Total Variance

- The expectation of the conditional variance $\operatorname{var}(X \mid Y)$ is related to the unconditional variance $\operatorname{var}(X)$

$$
\operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])
$$

$$
\begin{aligned}
& \operatorname{var}(X)=\mathbf{E}\left[X^{2}\right]-(\mathbf{E}[X])^{2} \\
&\left.=\mathbf{E}\left[\mathbf{E}\left[X^{2} \mid Y\right]\right]-(\mathbf{E}[\mathbf{E}[X \mid Y]])^{2}\right) \\
& \begin{aligned}
\operatorname{var}(Z)=\mathbf{E}\left[Z^{2}\right]-\left(\mathbf{E}[Z)^{2}\right. \\
\Rightarrow \mathbf{E}\left[Z^{2}\right]=\operatorname{var}(Z)+\left(\mathbf{E}[Z)^{2}\right.
\end{aligned}=\mathbf{E}\left[\operatorname{var}(X \mid Y)+(\mathbf{E}[X \mid Y])^{2}\right]-(\mathbf{E}[\mathbf{E}[X \mid Y]])^{2} \\
& \mathbf{E}[X]=\mathbf{E}[\mathbf{E}[X \mid Y]]
\end{aligned}
$$

## Illustrative Examples (1/4)

- Example 4.17. (continued) Consider again the problem where we break twice a stick of length , at randomly chosen points, with $Y$ being the length of the stick after the first break and $X$ being the length after the second break
- Calculate $\operatorname{var}(X)$ using the law of total variance

$$
\operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])
$$


We also know that if a random variable $Z \quad \Rightarrow \operatorname{var}(X \mid Y=y)=\frac{y^{2}}{12}$
is uniformly distribute d in $[a, b]$, then its $\Rightarrow$ "random variable" $\operatorname{var}(X \mid Y)$ can be variance is
$\operatorname{var}(Z)=\frac{(b-a)^{2}}{12}$
expressed as $\frac{Y^{2}}{12}\left(\right.$ That is $\left.\operatorname{var}(X \mid Y)=\frac{Y^{2}}{12}\right)$
(a function of $Y$ )

## Illustrative Examples (2/4)

$$
\begin{aligned}
\mathbf{E}[\operatorname{var}(X \mid Y)] & =\int_{0}^{l} \operatorname{var}(X \mid Y=y) f_{Y}(y) d y \\
& =\int_{0}^{l} \frac{y^{2}}{12} \frac{1}{l} d y \\
& =\left.\frac{y^{3}}{36 \cdot l}\right|_{0} ^{l}=\frac{l^{2}}{36}
\end{aligned}
$$

$$
\square \quad \therefore \operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])
$$

Note that $\mathbf{E}[X \mid Y=y]=\frac{y}{2} \quad$ cf. p. 14

$$
=\frac{l^{2}}{36}+\frac{l^{2}}{48}=\frac{7 \cdot l^{2}}{144}
$$

$\Rightarrow \mathbf{E}[X \mid Y]=\frac{Y}{2} \quad$ (a function of $Y$ )
$\Rightarrow \operatorname{var}(\mathbf{E}[X \mid Y])$
$=\operatorname{var}\left(\frac{Y}{2}\right)=\frac{1}{4} \operatorname{var}(Y)$
$=\frac{1}{4} \cdot \frac{l^{2}}{12}=\frac{l^{2}}{48} \quad(Y$ is uniformly distributed $)$

## Illustrative Examples (3/4)

- Example 4.21. Computing Variances by Conditioning.
- Consider a continuous random variable $X$ with the PDF given in the following figure. We define an auxiliary (discrete) random variable $Y$ as follows:


$$
\operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])
$$

We know that if a random variable $Z$ is uniformly distribute d in $[a, b]$, then its variance is

$$
\operatorname{var}(Z)=\frac{(b-a)^{2}}{12}
$$

(1)

$$
\begin{aligned}
\Rightarrow & p_{Y}(1)=\int_{0}^{1} 1 / 2 d x=1 / 2 \\
& p_{Y}(2)=\int_{1}^{3} 1 / 4 d x=1 / 2 \\
\Rightarrow & f_{X \mid Y}(x \mid Y=1)=\left\{\begin{array}{l}
1, \text { for } 0 \leq x \leq 1 \\
0, \text { otherwise }
\end{array}\right. \\
& f_{X \mid Y}(x \mid Y=2)=\left\{\begin{array}{l}
1 / 2, \text { for } 1 \leq x \leq 3 \\
0, \text { otherwise }
\end{array}\right.
\end{aligned}
$$

(2) $\operatorname{var}(X \mid Y=1)=(1-0)^{2} / 12=1 / 12$ $\operatorname{var}(X \mid Y=2)=(3-1)^{2} / 12=1 / 3$
$\Rightarrow \mathbf{E}[\operatorname{var}(X \mid Y)]=\operatorname{var}(X \mid Y=1) p_{Y}(1)+\operatorname{var}(X \mid Y=2) p_{Y}(2)$

$$
=1 / 12 \cdot 1 / 2+1 / 3 \cdot 1 / 2=5 / 24
$$

Note that $f_{X \mid A}(x)= \begin{cases}\frac{f_{X}(x)}{P(X \in A)}, & \text { if } x \in A \\ 0, & \text { otherwise }\end{cases}$

## Illustrative Examples (4/4)

We know that if a random variable $Z$
(4) $\therefore \operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])$

$$
=9 / 16+5 / 24=37 / 48
$$

mean is

$$
E[Z]=\frac{a+b}{2}
$$

(3)

$$
\begin{aligned}
& \Rightarrow \mathbf{E}[X \mid Y=1]=(0+1) / 2=1 / 2 \\
& \quad \mathbf{E}[X \mid Y=2]=(1+3) / 2=2 \\
& \mathbf{E}[\mathbf{E}[X \mid Y]=\mathbf{E}[X] \\
& \quad=\mathbf{E}[X \mid Y=1] p_{r_{Y}}(1)+\mathbf{E}[X \mid Y=2] p_{Y}(2) \\
& \quad=1 / 2 \cdot 1 / 2+2 \cdot 1 / 2=5 / 4 \\
& \operatorname{var}(\mathbf{E}[X \mid Y])= \\
& \left(\mathbf{E}[X \mid Y=1]-\mathbf{E}\left[\mathbf{E}[X|Y|)^{2} p_{Y}(1)\right.\right. \\
& \quad+\left(\mathbf{E}[X \mid Y=2]-\mathbf{E}[\mathbf{E}[X \mid Y])^{2} p_{Y}(2)\right. \\
& =(1 / 2-5 / 4)^{2} \cdot 1 / 2+(2-5 / 4)^{2} \cdot 1 / 2 \\
& =9 / 16
\end{aligned}
$$

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Justification
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$$
\begin{aligned}
\mathbf{E}[X] & =\int_{-\infty}^{\infty} x \cdot f_{X}(x) d x \\
& =\int_{0}^{1} x \cdot 1 / 2 d x+\int_{1}^{3} x \cdot 1 / 4 d x \\
& =\left.\frac{1}{4} x^{2}\right|_{0} ^{1}+\left.\frac{1}{8} x^{2}\right|_{1} ^{3} \\
& =5 / 4 \\
\operatorname{var}(x) & =\int_{-\infty}^{\infty}(x-\mathbf{E}[X])^{2} \cdot f_{X}(x) d x \\
& =\int_{0}^{1}(x-5 / 4)^{2} \cdot 1 / 2 d x+\int_{1}^{3}(x-5 / 4)^{2} \cdot 1 / 4 d x \\
& =\left.\frac{1}{6}(x-5 / 4)^{3}\right|_{0} ^{1}+\left.\frac{1}{12}(x-5 / 4)^{3}\right|_{1} ^{3} \\
& =\frac{1}{6}\left((-1 / 4)^{3}-(-5 / 4)^{3}\right)+\frac{1}{12}\left((7 / 4)^{3}-(-1 / 4)^{3}\right) \\
& =\frac{1}{6} \cdot \frac{124}{64}+\frac{1}{12} \cdot \frac{344}{64}=\frac{37}{48}
\end{aligned}
$$

Note that for discrete random variable $Z$
$\operatorname{var}(Z)=\mathbf{E}\left[(Z-\mathbf{E}(Z))^{2}\right]=\sum_{z}(z-\mathbf{E}(Z))^{2} p_{Z}(z)$

## Averaging by Section

- For a two-section (or two-cluster) problem

$\bigcirc: x_{i} \in$ section 1
$\triangle: x_{i} \in \operatorname{section} 2$
$\mathbf{E}[X]=\mathbf{E}\left[\mathbf{E}[X \mid Y]=\sum \mathbf{E}[X \mid Y=s] p(Y=s)\right.$
$\operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])$
These two measures have been widely used for linear discriminant analysis (LDA)
variability of $\mathbf{E}[X \mid Y]$ (the outcome means of individual sections)

Also called "within cluster" variation

## Properties of Conditional Expectation and Variance

- $\mathbf{E}[X \mid Y=y]$ is a number, whose value depends on $y$.
- $\mathbf{E}[X \mid Y]$ is a function of the random variable $Y$, hence a random variable. Its experimental value is $\mathbf{E}[X \mid Y=y]$ whenever the experimental value of $Y$ is $y$.
- $\mathbf{E}[\mathbf{E}[X \mid Y]]=\mathbf{E}[X]$ (law of iterated expectations).
- $\operatorname{var}(X \mid Y)$ is a random variable whose experimental value is $\operatorname{var}(X \mid Y=$ $y$ ), whenever the experimental value of $Y$ is $y$.
- $\operatorname{var}(X)=\mathbf{E}[\operatorname{var}(X \mid Y)]+\operatorname{var}(\mathbf{E}[X \mid Y])$.

