

Hidden Markov Models for Speech Recognition



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References:

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- 5. Young. *HMMs and Related Speech Recognition Technologies*. Chapter 27, Springer Handbook of Speech Processing, Springer, 2007
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Hidden Markov Model (HMM): A Brief Overview

<u>History</u>

- Published in papers of Baum in late 1960s and early 1970s
- Introduced to speech processing by Baker (CMU) and Jelinek (IBM) in the 1970s (discrete HMMs)
- Then extended to continuous HMMs by Bell Labs

Assumptions

- Speech signal can be characterized as a parametric random (stochastic) process
- Parameters can be estimated in a precise, well-defined manner

Three fundamental problems

- Evaluation of probability (likelihood) of a sequence of observations given a specific HMM
- Adjustment of model parameters so as to best account for observed signals
- Determination of a best sequence of model states

Stochastic Process

- A stochastic process is a mathematical model of a probabilistic experiment that evolves in time and generate s a sequence of numeric values
 - Each numeric value in the sequence is modeled by a random variable
 - A stochastic process is just a (finite/infinite) sequence of random variables
- Examples
 - (a) The sequence of recorded values of a speech utterance
 - (b) The sequence of daily prices of a stock
 - (c) The sequence of hourly traffic loads at a node of a communication network
 - (d) The sequence of radar measurements of the position of an airplane

Observable Markov Model

- Observable Markov Model (Markov Chain)
 - First-order Markov chain of N states is a triple (S, A, π)
 - **S** is a set of *N* states
 - **A** is the $N \times N$ matrix of transition probabilities between states $P(s_t=j|s_{t-1}=i, s_{t-2}=k,)=P(s_t=j|s_{t-1}=i)=A_{ij}$ First-order and time-invariant assumptions
 - π is the vector of initial state probabilities $\pi_i = P(s_1 = j)$
 - The output of the process is the set of states at each instant of time, when each state corresponds to an observable event
 - The output in any given state is not random (*deterministic!*)
 - Too simple to describe the speech signal characteristics

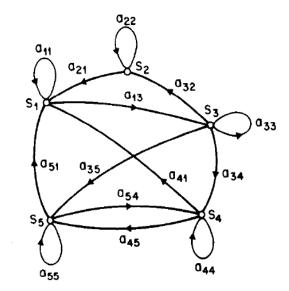
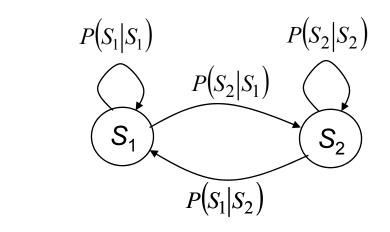
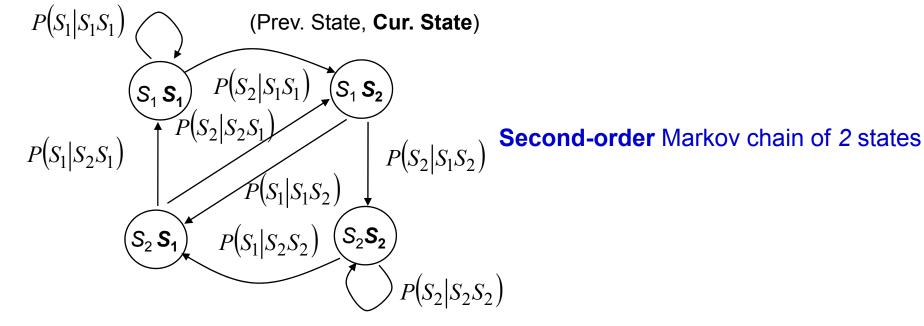


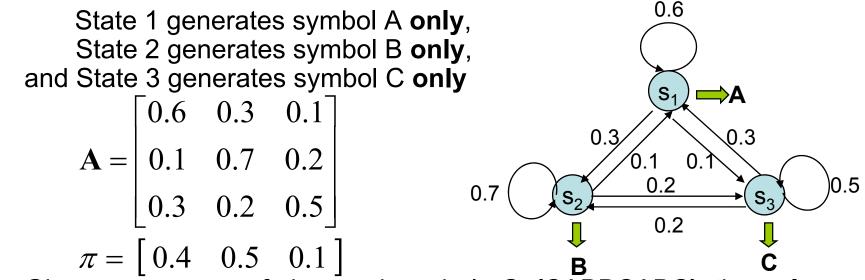
Fig. 1. A Markov chain with 5 states (labeled S_1 to S_5) with selected state transitions.



First-order Markov chain of 2 states







- Given a sequence of observed symbols $O = \{CABBCABC\}$, the only one corresponding state sequence is $\{S_3S_1S_2S_2S_3S_1S_2S_3\}$, and the corresponding probability is

 $P(\mathbf{O}|\lambda) = P(S_3)P(S_1|S_3)P(S_2|S_1)P(S_2|S_2)P(S_3|S_2)P(S_1|S_3)P(S_2|S_1)P(S_3|S_2) = 0.1 \times 0.3 \times 0.3 \times 0.7 \times 0.2 \times 0.3 \times 0.3 \times 0.2 = 0.00002268$

• Example 2: A three-state Markov chain for the *Dow Jones Industrial average*

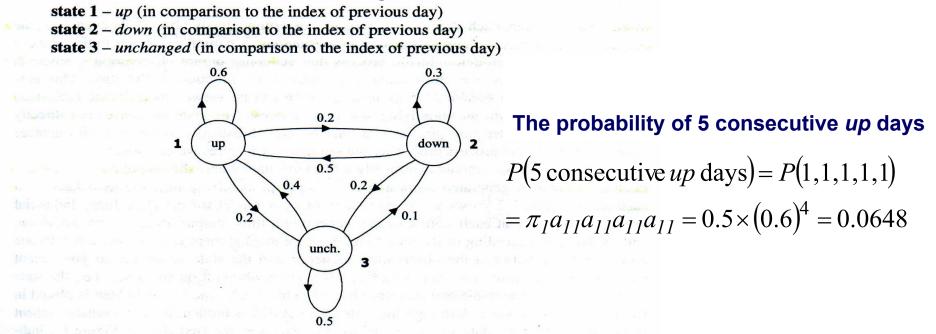


Figure 8.1 A Markov chain for the Dow Jones Industrial average. Three states represent up, down, and unchanged, respectively.

The parameter for this Dow Jones Markov chain may include a state-transition probability matrix $\begin{bmatrix} 0 & - \end{bmatrix}$

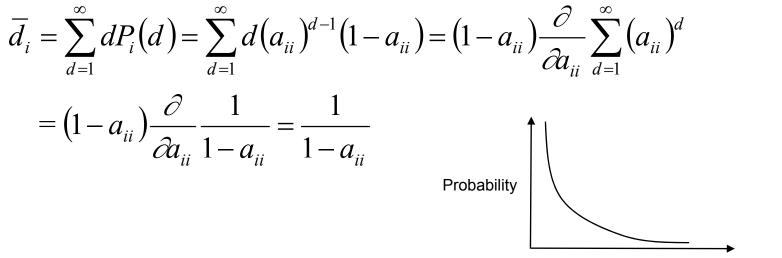
Name of the second	0.6	0.2	0.2]	de transmissione alger	0.5	
$A = \left\{a_{ij}\right\} =$	0.5	0.3	0.2	$\boldsymbol{\pi} = (\boldsymbol{\pi}_i)^t =$	0.2	
	0.4	0.1	0.5		0.3	1

and an initial state probability matrix

• Example 3: Given a Markov model, what is the mean occupancy duration of each state *i*

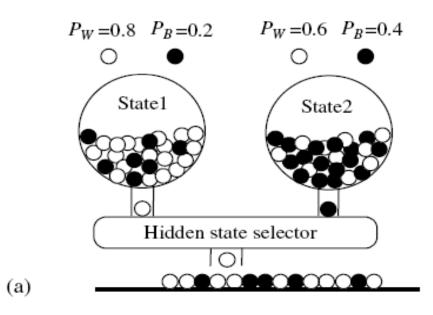
 $P_i(d)$ = probability mass function of duration d in state i= $(a_{ii})^{d-1}(1-a_{ii})$

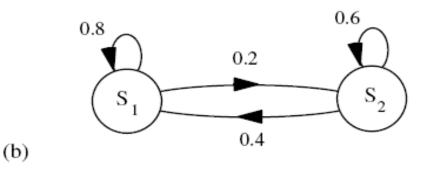
Expected number of duration in a state



Time (Duration)

Hidden Markov Model





(a) Illustration of a two-layered random process. (b) An HMM model of the process in (a).

- HMM, an extended version of Observable Markov Model
 - The observation is turned to be a probabilistic function (discrete or continuous) of a state instead of an one-to-one correspondence of a state
 - The model is a doubly embedded stochastic process with an underlying stochastic process that is not directly observable (hidden)
 - What is hidden? **The State Sequence!** According to the observation sequence, we are not sure which state sequence generates it!
- Elements of an HMM (the State-Output HMM) $\lambda = \{S, A, B, \pi\}$
 - **S** is a set of *N* states
 - **A** is the N×N matrix of transition probabilities between states
 - B is a set of N probability functions, each describing the observation probability with respect to a state
 - $-\pi$ is the vector of initial state probabilities

- Two major assumptions
 - First order (Markov) assumption
 - The state transition depends only on the origin and destination
 - Time-invariant

$$P(s_t = j | s_{t-1} = i) = P(s_\tau = j | s_{\tau-1} = i) = P(j | i) = A_{i,j}$$

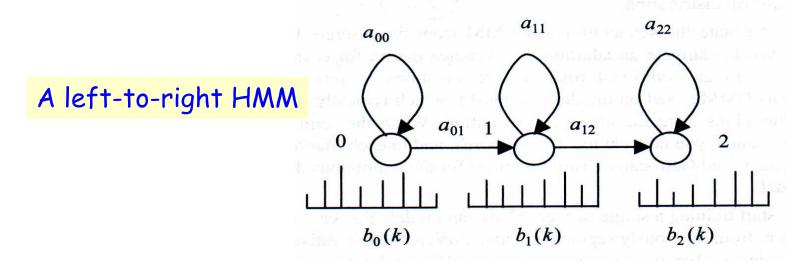
- Output-independent assumption
 - All observations are dependent on the state that generated them, not on neighboring observations

$$P(\mathbf{o}_t | s_t, \dots, \mathbf{o}_{t-2}, \mathbf{o}_{t-1}, \mathbf{o}_{t+1}, \mathbf{o}_{t+2} \dots) = P(\mathbf{o}_t | s_t)$$

- Two major types of HMMs according to the observations
 - Discrete and finite observations:
 - The observations that all distinct states generate are finite in number

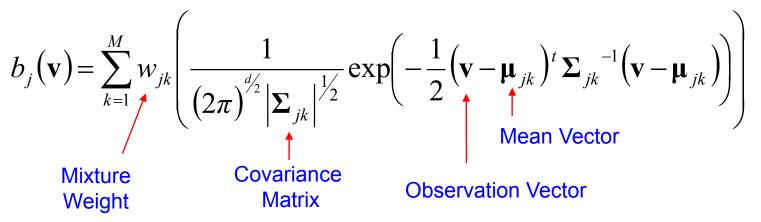
 $\boldsymbol{V}=\{\boldsymbol{v}_1, \ \boldsymbol{v}_2, \ \boldsymbol{v}_3, \ \ldots, \ \boldsymbol{v}_M\}, \ \boldsymbol{v}_k \in \boldsymbol{R}^L$

In this case, the set of observation probability distributions B={b_j(v_k)}, is defined as b_j(v_k)=P(o_t=v_k|s_t=j), 1≤k≤M, 1≤j≤N o_t: observation at time t, s_t: state at time t
 ⇒ for state j, b_i(v_k) consists of only M probability values



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- Two major types of HMMs according to the observations
 - Continuous and infinite observations:
 - The observations that all distinct states generate are infinite and continuous, that is, V={v | v∈R^d}
 - In this case, the set of observation probability distributions B={b_j(v)}, is defined as b_j(v)=f_{O|S}(o_t=v|s_t=j), 1≤j≤N
 ⇒ b_j(v) is a continuous probability density function (pdf) and is often a mixture of Multivariate Gaussian (Normal) Distributions



- Multivariate Gaussian Distributions
 - When X=(X₁, X₂,..., X_d) is a *d*-dimensional random vector, the multivariate Gaussian pdf has the form:

$$f(\mathbf{X} = \mathbf{x} | \mathbf{\mu}, \mathbf{\Sigma}) = N(\mathbf{x}; \mathbf{\mu}, \mathbf{\Sigma}) = \frac{1}{(2\pi)^{\frac{d}{2}} |\mathbf{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(\mathbf{x} - \mathbf{\mu})^{t} \mathbf{\Sigma}^{-1}(\mathbf{x} - \mathbf{\mu})\right)$$

where $\mathbf{\mu}$ is the *L* - dimensional mean vector, $\mathbf{\mu} = E[\mathbf{x}]$
 $\mathbf{\Sigma}$ is the coverance matrix, $\mathbf{\Sigma} = E[(\mathbf{x} - \mathbf{\mu})(\mathbf{x} - \mathbf{\mu})^{t}] = E[\mathbf{x}\mathbf{x}^{t}] - \mathbf{\mu}\mathbf{\mu}^{t}$
and $|\mathbf{\Sigma}|$ is the the determinant of $\mathbf{\Sigma}$
The *i*-*j*th elevment σ_{ij} of $\mathbf{\Sigma}, \sigma_{ij} = E[(x_{i} - \mu_{i})(x_{j} - \mu_{j})] = E[x_{i}x_{j}] - \mu_{i}\mu_{j}$

- If $X_1, X_2, ..., X_d$ are independent, the covariance matrix is reduced to diagonal covariance
 - The distribution as *d* independent scalar Gaussian distributions
 - Model complexity is significantly reduced

• Multivariate Gaussian Distributions

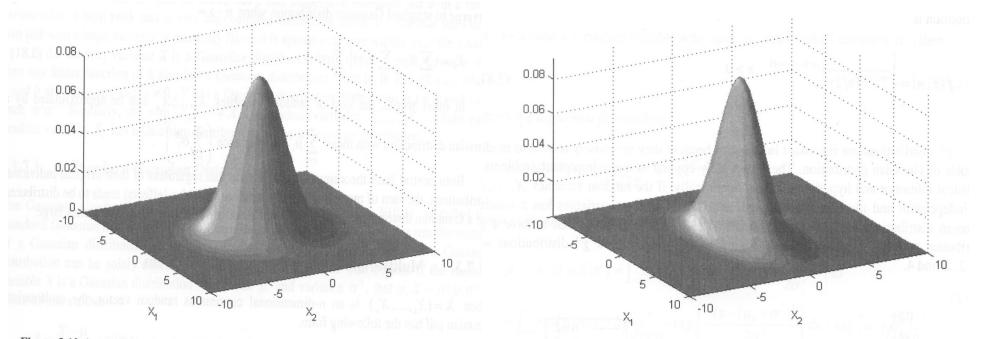
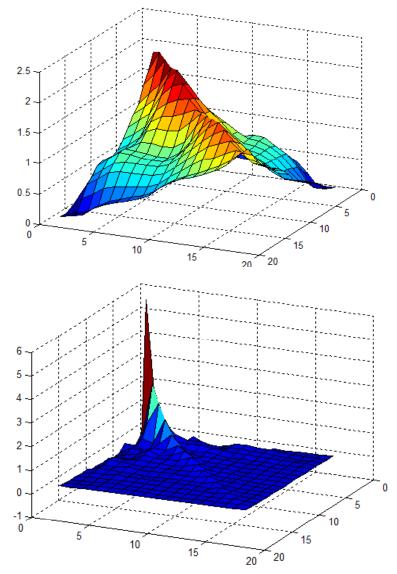


Figure 3.12 A two-dimensional multivariate Gaussian distribution with independent random variables x_1 and x_2 that have the same variance.

Figure 3.13 Another two-dimensional multivariate Gaussian distribution with independent random variable x_1 and x_2 which have different variances.

 Covariance matrix of the correlated feature vectors (Mel-frequency filter bank outputs)

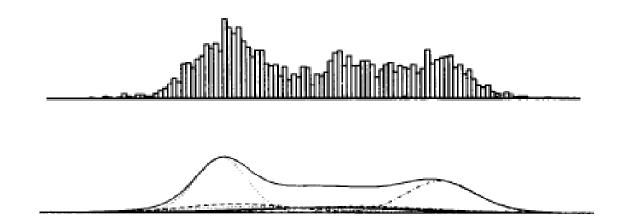
- Covariance matrix of the partially de-correlated feature vectors (MFCC without C₀)
 - MFCC: Mel-frequency cepstral coefficients



- Multivariate Mixture Gaussian Distributions (cont.)
 - More complex distributions with multiple local maxima can be approximated by Gaussian (a unimodal distribution) mixture

$$f(\mathbf{x}) = \sum_{k=1}^{M} w_k N_k(\mathbf{x}; \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \qquad \sum_{k=1}^{M} w_k = 1$$

 Gaussian mixtures with enough mixture components can approximate any distribution



- Ergodic HMM • Example 4: a 3-state discrete HMM λ 0.6 0.6 0.3 0.1 $\mathbf{A} = \begin{vmatrix} 0.1 & 0.7 & 0.2 \\ 0.3 & 0.2 & 0.5 \end{vmatrix}$ **S**₁) →{A:.3,B:.2,C:.5} 0.3 0.3 $b_1(\mathbf{A}) = 0.3, b_1(\mathbf{B}) = 0.2, b_1(\mathbf{C}) = 0.5$ 01 0.2 0.5 0.7 $b_2(\mathbf{A}) = 0.7, b_2(\mathbf{B}) = 0.1, b_2(\mathbf{C}) = 0.2$ S₃ $| S_2$ 0.2 $b_3(\mathbf{A}) = 0.3, b_3(\mathbf{B}) = 0.6, b_3(\mathbf{C}) = 0.1$ $\pi = \begin{bmatrix} 0.4 & 0.5 & 0.1 \end{bmatrix}$ **{A:**.7,**B:**.1,**C:**.2**} {A:**.3,**B:**.6,**C:**.1**}**
 - Given a sequence of observations O={ABC}, there are 27
 possible corresponding state sequences, and therefore the corresponding probability is

$$P(\mathbf{O}|\lambda) = \sum_{i=1}^{27} P(\mathbf{O}, \mathbf{S}_i | \lambda) = \sum_{i=1}^{27} P(\mathbf{O}|\mathbf{S}_i, \lambda) P(\mathbf{S}_i | \lambda), \quad \mathbf{S}_i : \text{state sequence}$$

E.g. when $\mathbf{S}_i = \{s_2 s_2 s_3\}, P(\mathbf{O}|\mathbf{S}_i, \lambda) = P(\mathbf{A}|s_2) P(\mathbf{B}|s_2) P(\mathbf{C}|s_3) = 0.7 * 0.1 * 0.1 = 0.007$

$$P(\mathbf{S}_i | \lambda) = P(s_2) P(s_2|s_2) P(s_3|s_2) = 0.5 * 0.7 * 0.2 = 0.07$$

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- Notation :
 - $O=\{o_1o_2o_3....o_7\}$: the observation (feature) sequence
 - **S**= { $s_1s_2s_3...,s_7$ }: the state sequence
 - λ : model, for HMM, $\lambda = \{A, B, \pi\}$
 - $P(\mathbf{O}|\lambda)$: The probability of observing **O** given the model λ
 - $P(O|S,\lambda)$: The probability of observing **O** given λ and a state sequence **S** of λ
 - $P(\mathbf{O}, \mathbf{S} | \lambda)$: The probability of observing **O** and **S** given λ
 - $P(S|O,\lambda)$: The probability of observing S given O and λ
- Useful formula

- Bayes' Rule :

$$P(A|B) = \frac{P(A,B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)} \implies \frac{P(A|B,\lambda) = \frac{P(A,B|\lambda)}{P(B|\lambda)} = \frac{P(B|A,\lambda)P(A|\lambda)}{P(B|\lambda)}}{P(B|\lambda)}$$

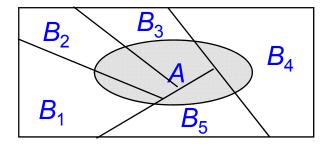
$$\lambda : \text{model describing the probability}$$

P(A,B) = P(B|A)P(A) = P(A|B)P(B) chain rule

- Useful formula (Cont.):
 - Total Probability Theorem

marginal probability $P(A) = \begin{cases} \sum_{all \ B} P(A, B) = \sum_{all \ B} P(A|B)P(B), & \text{if } B \text{ is disrete and disjoint} \\ \int_{B} f(A, B)dB = \int_{B} f(A|B)f(B)dB, & \text{if } B \text{ is continuous} \end{cases}$

if
$$x_1, x_2, \dots, x_n$$
 are independent,
 $\Rightarrow P(x_1, x_2, \dots, x_n) = P(x_1)P(x_2)\dots P(x_n)$



Venn Diagram

$$E_{z}[q(z)] = \begin{cases} \sum_{k} P(z=k)q(k), & z: \text{discrete} \\ \int_{z}^{k} f_{z}(z)q(z)dz, & z: \text{continuous} \end{cases}$$

Expectation

Three Basic Problems for HMM

- Given an observation sequence $O=(o_1, o_2, \dots, o_T)$, and an HMM $\lambda = (S, A, B, \pi)$
 - Problem 1:

How to *efficiently* compute $P(\mathbf{O}|\lambda)$?

Evaluation problem

– Problem 2:

How to choose an optimal state sequence $S=(s_1, s_2, ..., s_T)$?

Decoding Problem

– Problem 3:

How to adjust the model parameter $\lambda = (A, B, \pi)$ to maximize $P(O|\lambda)$? \Rightarrow *Learning / Training Problem*

Given **O** and λ , find $P(\mathbf{O}|\lambda) = \text{Prob}[\text{observing O given } \lambda]$

- Direct Evaluation
 - Evaluating all possible state sequences of length *T* that generating observation sequence *O*

$$P(\boldsymbol{O} \mid \lambda) = \sum_{all \mid \boldsymbol{S}} P(\boldsymbol{O}, \boldsymbol{S} \mid \lambda) = \sum_{all \mid \boldsymbol{S}} P(\boldsymbol{O} \mid \boldsymbol{S}, \lambda) P(\boldsymbol{S} \mid \lambda)$$

- $P(\mathbf{S} | \lambda)$: The probability of each path **S**
 - By Markov assumption (First-order HMM)

$$P\left(\boldsymbol{S} \mid \boldsymbol{\lambda}\right) = P\left(s_{1} \mid \boldsymbol{\lambda}\right) \prod_{t=2}^{T} P\left(s_{t} \mid s_{1}^{t-1}, \boldsymbol{\lambda}\right)$$
By chain rule

$$\approx P\left(s_{1} \mid \boldsymbol{\lambda}\right) \prod_{t=2}^{T} P\left(s_{t} \mid s_{t-1}, \boldsymbol{\lambda}\right)$$
By Markov assumption

$$= \pi_{s_{1}} a_{s_{1}s_{2}} a_{s_{2}s_{3}} \dots a_{s_{T-1}s_{T}}$$

- Direct Evaluation (cont.)
 - $P(O|S, \lambda)$: The joint output probability along the path S
 - By output-independent assumption
 - The probability that a particular observation symbol/vector is emitted at time *t* depends only on the state s_t and is conditionally independent of the past observations

$$P(\boldsymbol{O} | \boldsymbol{S}, \boldsymbol{\lambda}) = P(\boldsymbol{o}_{1}^{T} | \boldsymbol{s}_{1}^{T}, \boldsymbol{\lambda})$$

$$= P(\boldsymbol{o}_{1} | \boldsymbol{s}_{1}^{T}, \boldsymbol{\lambda}) \prod_{t=2}^{T} P(\boldsymbol{o}_{t} | \boldsymbol{o}_{1}^{t-1}, \boldsymbol{s}_{1}^{T}, \boldsymbol{\lambda})$$

$$\approx \prod_{t=1}^{T} P(\boldsymbol{o}_{t} | \boldsymbol{s}_{t}, \boldsymbol{\lambda}) \quad \text{By output-independent assumption}$$

$$= \prod_{t=1}^{T} b_{s_{t}}(\boldsymbol{o}_{t})$$

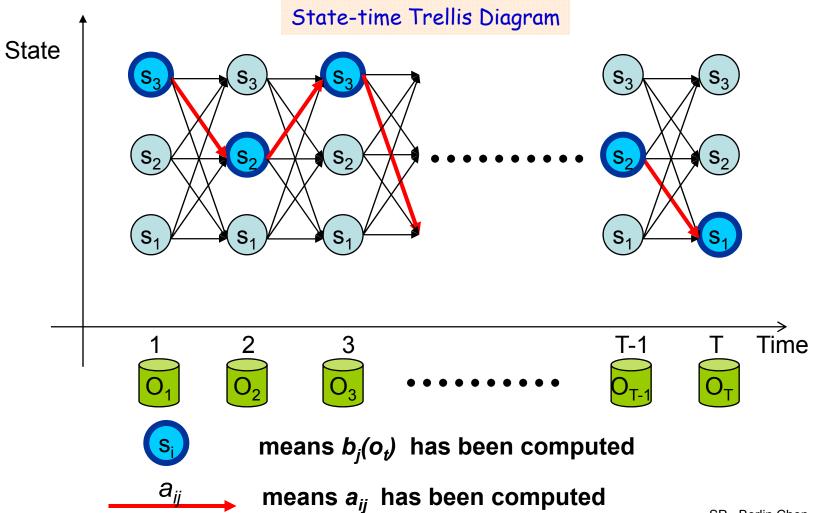
- Direct Evaluation (Cont.) $P(\boldsymbol{o}_{l}|s_{t},\lambda) = b_{s_{t}}(\boldsymbol{o}_{t})$ $P(\boldsymbol{o}|\lambda) = \sum_{all \ \boldsymbol{s}} P(\boldsymbol{S}|\lambda) P(\boldsymbol{O}|\boldsymbol{S},\lambda)$ $= \sum_{all \ \boldsymbol{s}} \left(\left[\pi_{s_{1}}a_{s_{1}s_{2}}a_{s_{2}s_{3}}\dots a_{s_{T-1}s_{T}} \right] \left[b_{s_{1}}(\boldsymbol{o}_{1})b_{s_{2}}(\boldsymbol{o}_{2})\dots b_{s_{T}}(\boldsymbol{o}_{T}) \right] \right)$ $= \sum_{s_{1},s_{2}\dots,s_{T}} \pi_{s_{1}}b_{s_{1}}(\boldsymbol{o}_{1})a_{s_{1}s_{2}}b_{s_{2}}(\boldsymbol{o}_{2})\dots a_{s_{T-1}s_{T}}b_{s_{T}}(\boldsymbol{o}_{T})$
 - Huge Computation Requirements: $O(N^{T})$
 - Exponential computational complexity

Complexity : $(2T-1)N^T MUL \approx 2TN^T$, N^T-1 ADD

• A more efficient algorithms can be used to evaluate $P(\mathbf{O}|\lambda)$

- Forward/Backward Procedure/Algorithm

• Direct Evaluation (Cont.)



- The Forward Procedure

- Base on the HMM assumptions, the calculation of $P(s_t | s_{t-1}, \lambda)$ and $P(o_t | s_t, \lambda)$ involves only s_{t-1} , s_t and o_t , so it is possible to compute the likelihood with recursion on t
- Forward variable : $\alpha_{t}(i) = P(o_{1}o_{2}...o_{t}, s_{t} = i|\lambda)$
 - The probability that the HMM is in state *i* at time *t* having generating partial observation $o_1 o_2 \dots o_t$

Basic Problem 1 of HMM - The Forward Procedure (cont.)

Algorithm

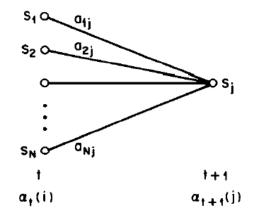
1. Initialization $\alpha_1(i) = \pi_i b_i(o_1), \ 1 \le i \le N$

2. Induction
$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right] b_j(\boldsymbol{o}_{t+1}), \quad 1 \le t \le T-1, 1 \le j \le N$$

3.Termination
$$P(\boldsymbol{O}|\lambda) = \sum_{i=1}^{N} \alpha_T(i)$$

- Complexity: $O(N^2T)$

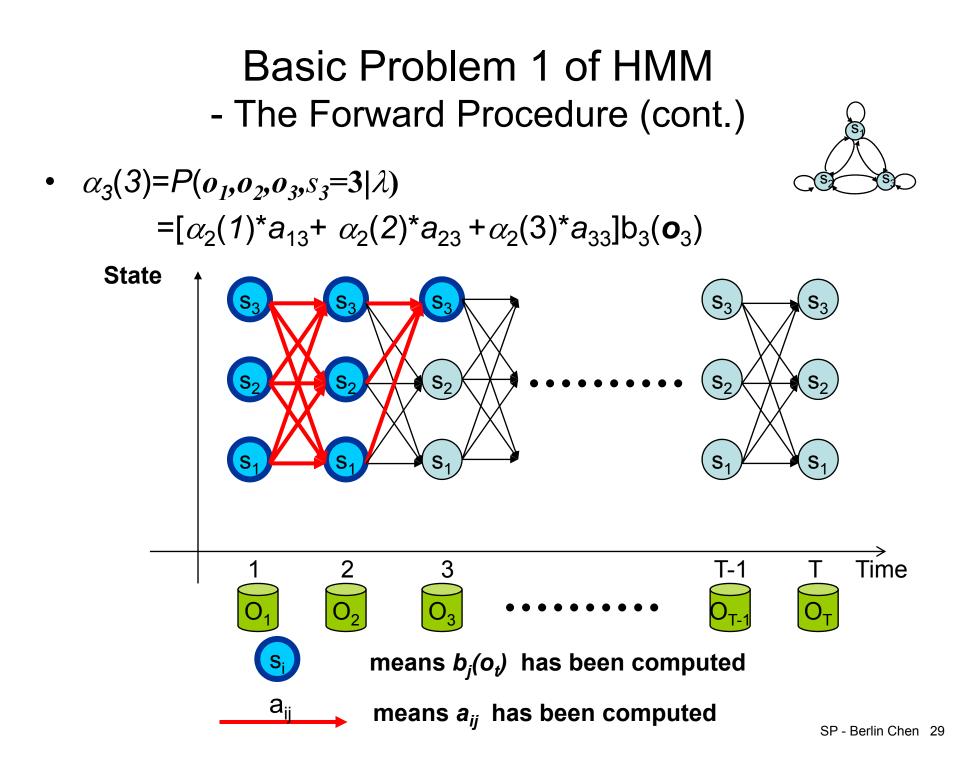
MUL : $N(N+1)(T-1)+N \approx N^2 T$ ADD : $(N-1)N(T-1)+(N-1)\approx N^2 T$



- Based on the lattice (trellis) structure
 - Computed in a *time-synchronous* fashion from *left-to-right*, where each cell for time *t* is completely computed before proceeding to time *t*+1
- All state sequences, regardless how long previously, merge to *N* nodes (states) at each time instance *t*

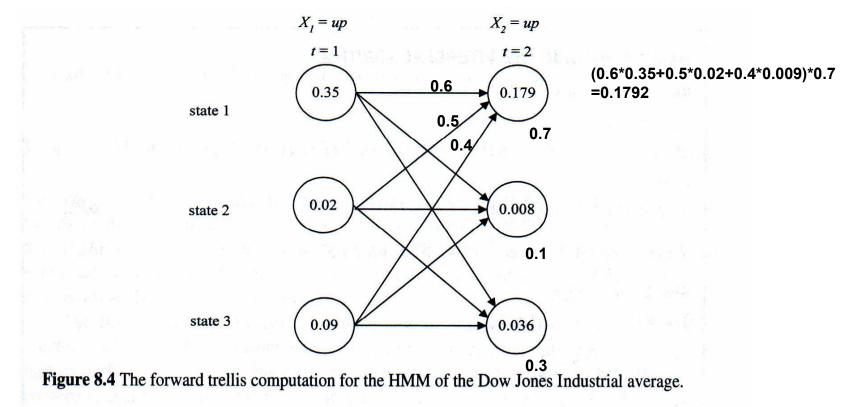
- The Forward Procedure (cont.)

$$\begin{aligned} \alpha_{t}(j) &= P(o_{1}o_{2}...o_{t}, s_{t} = j|\lambda) \\ &= P(o_{1}o_{2}...o_{t}|s_{t} = j,\lambda)P(s_{t} = j|\lambda) \end{aligned} \xrightarrow{P(A,B) = P(B|A)P(A)} \begin{array}{c} \text{output} \\ \text{independent} \\ \text{assumption} \end{aligned} \\ &= P(o_{1}o_{2}...o_{t-1}|s_{t} = j,\lambda)P(o_{t}|s_{t} = j,\lambda)P(s_{t} = j|\lambda) \\ &= P(o_{1}o_{2}...o_{t-1}, s_{t} = j|\lambda)P(o_{t}|s_{t} = j,\lambda) \\ &= P(o_{1}o_{2}...o_{t-1}, s_{t} = j|\lambda)P(o_{t}|s_{t} = j,\lambda) \\ &= P(o_{1}o_{2}...o_{t-1}, s_{t} = j|\lambda)b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i, s_{t} = j|\lambda)\right]b_{j}(o_{t}) \xrightarrow{P(A,B)} \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i, s_{t} = j|\lambda)\right]b_{j}(o_{t}) \xrightarrow{P(A,B)} \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i, s_{t} = j|\lambda)P(s_{t} = j|o_{1}o_{2}...o_{t-1}, s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i|\lambda)P(s_{t} = j|s_{t-1} = i, \lambda)\right]b_{j}(o_{t}) \\ &= \left[\sum_$$



Basic Problem 1 of HMM - The Forward Procedure (cont.)

• A three-state Hidden Markov Model for the *Dow Jones* Industrial average



- The Backward Procedure

• Backward variable : $\beta_t(i) = P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, \dots, \mathbf{o}_T | s_t = i, \lambda)$

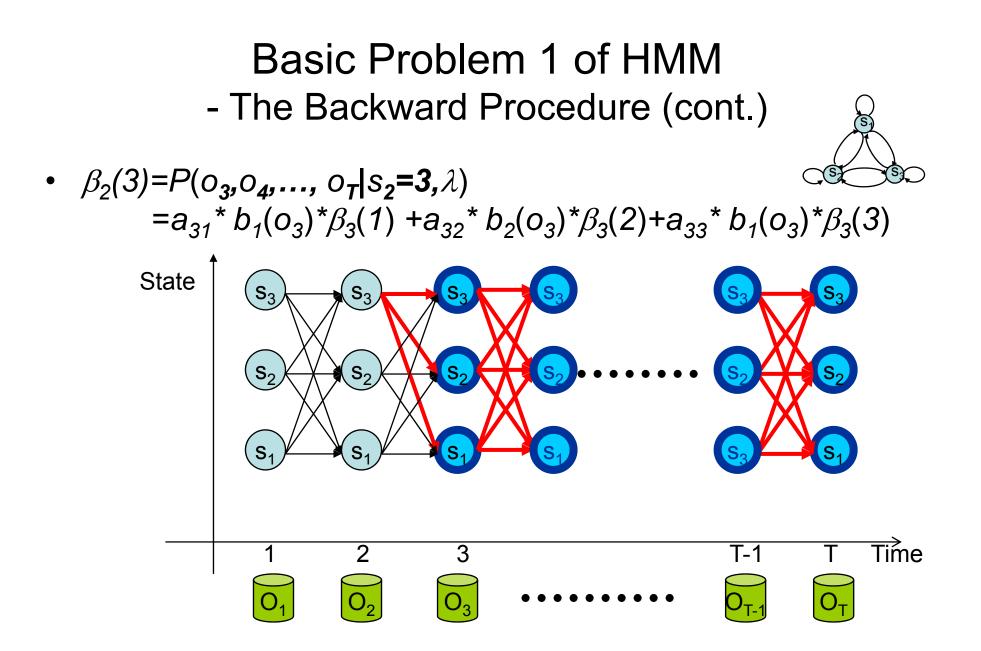
1. Initialization : $\beta_{T}(i) = 1$, $1 \le i \le N$ 2. Induction : $\beta_{t}(i) = \sum_{j=1}^{N} a_{ij} b_{j}(\boldsymbol{o}_{t+1}) \beta_{t+1}(j)$, $1 \le t \le T-1, 1 \le i \le N$ 3. Termination : $P(\boldsymbol{O}|\lambda) = \sum_{j=1}^{N} \pi_{j} b_{j}(\boldsymbol{o}_{1}) \beta_{1}(j)$ Complexity MUL: $2N^{2}(T-1) + 2N \approx N^{2}T$; ADD: $(N-1)N(T-1) + N \approx N^{2}T$

Basic Problem 1 of HMM - Backward Procedure (cont.)

• Why
$$P(\boldsymbol{O}, s_t = i | \lambda) = \alpha_t(i) \beta_t(i)$$
 ?

$$\begin{aligned} \alpha_{t}(i) \beta_{t}(i) \\ &= P(\boldsymbol{o}_{1}, \boldsymbol{o}_{2}, ..., \boldsymbol{o}_{t}, s_{t} = i | \lambda) \cdot P(\boldsymbol{o}_{t+1}, \boldsymbol{o}_{t+2}, ..., \boldsymbol{o}_{T} | s_{t} = i, \lambda) \\ &= P(\boldsymbol{o}_{1}, \boldsymbol{o}_{2}, ..., \boldsymbol{o}_{t} | s_{t} = i, \lambda) P(s_{t} = i | \lambda) P(\boldsymbol{o}_{t+1}, \boldsymbol{o}_{t+2}, ..., \boldsymbol{o}_{T} | s_{t} = i, \lambda) \\ &= P(\boldsymbol{o}_{1}, ..., \boldsymbol{o}_{t}, ..., \boldsymbol{o}_{T} | s_{t} = i, \lambda) P(s_{t} = i | \lambda) \\ &= P(\boldsymbol{o}_{1}, ..., \boldsymbol{o}_{t}, ..., \boldsymbol{o}_{T}, s_{t} = i | \lambda) \\ &= P(\boldsymbol{O}, s_{t} = i | \lambda) \end{aligned}$$
state

$$P(\boldsymbol{O} | \lambda) = \sum_{i=1}^{N} P(\boldsymbol{O}, s_{t} = i | \lambda) = \sum_{i=1}^{N} \alpha_{t}(i) \beta_{t}(i)$$



Basic Problem 2 of HMM

How to choose an optimal state sequence $S=(s_1, s_2, \dots, s_T)$?

 The first optimal criterion: Choose the states s_t are individually most likely at each time t

Define a posteriori probability variable $\gamma_t(i) = P(s_t = i | \boldsymbol{O}, \boldsymbol{\lambda})$

$$\gamma_{t}(i) = \frac{P(s_{t} = i, \boldsymbol{O} | \boldsymbol{\lambda})}{P(\boldsymbol{O} | \boldsymbol{\lambda})} = \frac{P(s_{t} = i, \boldsymbol{O} | \boldsymbol{\lambda})}{\sum_{m=1}^{N} P(s_{t} = m, \boldsymbol{O} | \boldsymbol{\lambda})} = \frac{\alpha_{t}(i) \beta_{t}(i)}{\sum_{m=1}^{N} \alpha_{t}(m) \beta_{t}(m)}$$

state occupation probability (count) – a soft alignment of HMM state to the observation (feature)

- Solution : $s_t^* = \arg_i \max[\gamma_t(i)], 1 \le t \le T$
 - Problem: maximizing the probability at each time *t* individually $S^* = s_1^* s_2^* \dots s_T^*$ may not be a valid sequence (e.g. $a_{s_t^*s_{t+1}^*} = 0$)

Basic Problem 2 of HMM (cont.) • $P(s_3 = 3, \mathbf{O} \mid \lambda) = \alpha_3(3)^* \beta_3(3)$ β₃(3) α₃(3) State a₂₃= S₂ Sa S time 2 3 T-1 Т 1 O_3 O_T 0 \cap

Basic Problem 2 of HMM - The Viterbi Algorithm

- The second optimal criterion: The Viterbi algorithm can be regarded as the dynamic programming algorithm applied to the HMM or as a modified forward algorithm
 - Instead of summing up probabilities from different paths coming to the same destination state, the Viterbi algorithm picks and remembers the best path
 - Find a single optimal state sequence $S=(s_1, s_2, \dots, s_T)$
 - How to find the second, third, etc., optimal state sequences (difficult ?)
 - The Viterbi algorithm also can be illustrated in a trellis framework similar to the one for the forward algorithm
 - State-time trellis diagram

- The Viterbi Algorithm (cont.)

Algorithm

Find a best state sequence $S = (s_1, s_2, ..., s_T)$ for a given observation $O = (o_1, o_2, ..., o_T)$?

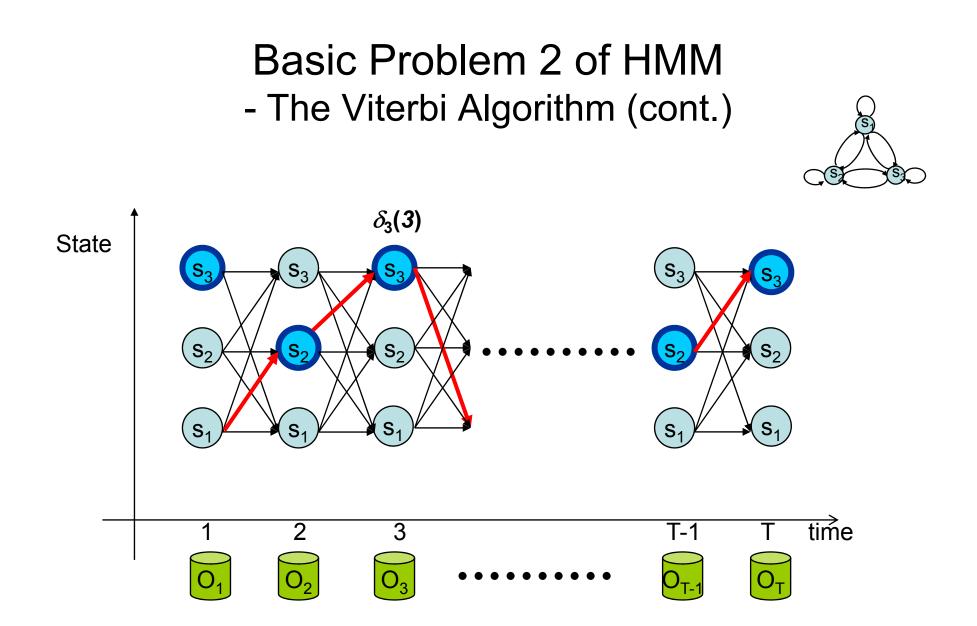
Define a new variable

$$\delta_t(i) = \max_{s_1, s_2, ..., s_{t-1}} P[s_1, s_2, ..., s_{t-1}, s_t = i, o_1, o_2, ..., o_t | \lambda]$$

= the best score along a single path at time t, which accounts for the first t observation and ends in state i

By induction $\therefore \delta_{t+1}(j) = \left[\max_{1 \le i \le N} \delta_t(i) a_{ij}\right] b_j(o_{t+1})$ $\psi_{t+1}(j) = \arg \max_{1 \le i \le N} \delta_t(i) a_{ij}$ For backtracing We can backtrace from $s_T^* = \arg \max_{1 \le i \le N} \delta_T(i)$

- Complexity: $O(N^2T)$



Basic Problem 2 of HMM - The Viterbi Algorithm (cont.)

• A three-state Hidden Markov Model for the *Dow Jones Industrial average*

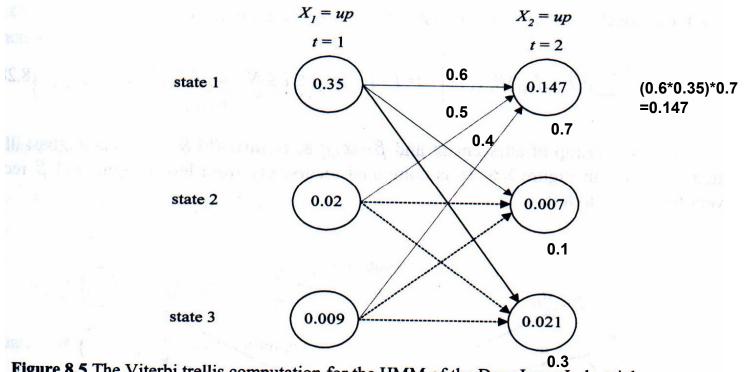


Figure 8.5 The Viterbi trellis computation for the HMM of the Dow Jones Industrial average.

Basic Problem 2 of HMM - The Viterbi Algorithm (cont.)

• Algorithm in the logarithmic form

Find a best state sequence $S = (s_1, s_2, ..., s_T)$ for a given observation $O = (o_1, o_2, ..., o_T)$?

Define a new variable

$$\delta_t(i) = \max_{s_1, s_2, ..., s_{t-1}} \log P[s_1, s_2, ..., s_{t-1}, s_t = i, o_1, o_2, ..., o_t | \lambda]$$

= the best score along a single path at time *t*, which accounts for the first *t* observation and ends in state *i*

By induction
$$\therefore \delta_{t+1}(j) = \left[\max_{1 \le i \le N} \left(\delta_t(i) + \log a_{ij}\right)\right] + \log b_j(o_{t+1})$$

 $\psi_{t+1}(j) = \arg \max_{1 \le i \le N} \left(\delta_t(i) + \log a_{ij}\right) \dots$ For backtracing
We can backtrace from $s_T^* = \arg \max_{1 \le i \le N} \delta_T(i)$

Homework-1

• A three-state Hidden Markov Model for the *Dow Jones* Industrial average

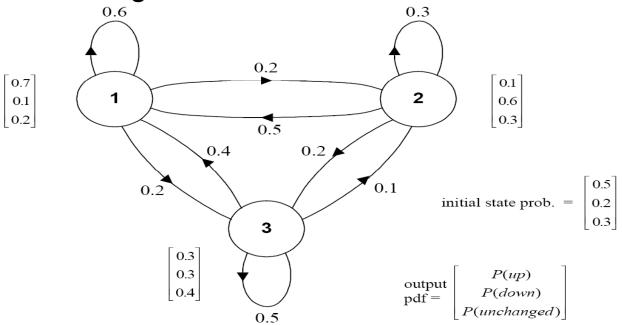


Figure 8.2 A hidden Markov model for the Dow Jones Industrial average. The three states no longer have deterministic meanings as the Markov chain illustrated in Figure 8.1.

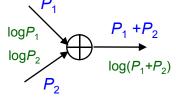
- Find the probability:

 $P(up, up, unchanged, down, unchanged, down, up|\lambda)$

 Fnd the optimal state sequence of the model which generates the observation sequence: (up, up, unchanged, down, unchanged, down, up)

Probability Addition in F-B Algorithm

 In Forward-backward algorithm, operations usually implemented in logarithmic domain



• Assume that we want to add P_1 and P_2

if
$$P_1 \ge P_2$$

 $\log_b (P_1 + P_2) = \log_b P_1 + \log_b (1 + b^{\log_b P_2 - \log_b P_1})$
else
 $\log_b (P_1 + P_2) = \log_b P_2 + \log_b (1 + b^{\log_b P_1 - \log_b P_2})$

The values of $\log_{b}(1+b^{x})$ can be saved in in a table to speedup the operations

Probability Addition in F-B Algorithm (cont.)

• An example code

```
#define LZERO (-1.0E10) // ~log(0)
#define LSMALL (-0.5E10) // log values < LSMALL are set to LZERO
#define minLogExp -log(-LZERO) // ~=-23
double LogAdd(double x, double y)
double temp,diff,z;
 if (x<y)
   temp = x; x = y; y = temp;
 diff = y-x; //notice that diff <= 0
 if (diff<minLogExp) // if y' is far smaller than x'
   return (x<LSMALL)? LZERO:x;
  else
   z = \exp(diff);
   return x + \log(1.0 + z);
```

Basic Problem 3 of HMM Intuitive View

- How to adjust (re-estimate) the model parameter λ=(A, B, π) to maximize P(O₁,..., O_L|λ) or logP(O₁,..., O_L|λ)?
 - Belonging to a typical problem of "inferential statistics"
 - The most difficult of the three problems, because there is no known analytical method that maximizes the joint probability of the training data in a close form

$$\log P(\mathbf{O}_{1}, \mathbf{O}_{2}, ..., \mathbf{O}_{L} | \lambda) = \log \prod_{l=1}^{L} P(\mathbf{O}_{l} | \lambda)$$

$$= \sum_{l=1}^{L} \log P(\mathbf{O}_{l} | \lambda) = \sum_{l=1}^{R} \log \sum_{all \in \mathbf{S}} P(\mathbf{S} | \lambda) P(\mathbf{O}_{l} | \mathbf{S}, \lambda)$$

The "log of sum" form is difficult to deal with

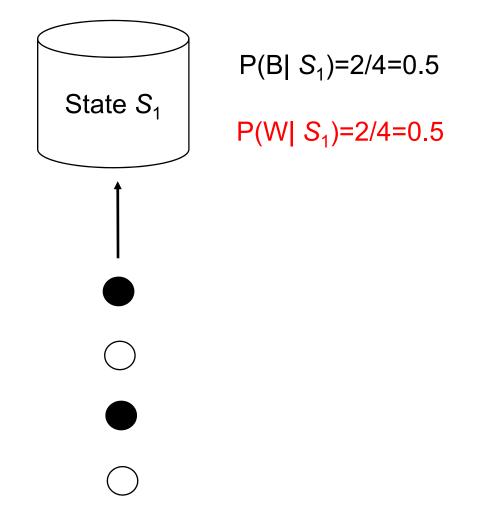
- Suppose that we have L training utterances for the HMM

- \mathbf{S} : a possible state sequence of the HMM

- The data is incomplete because of the hidden state sequences
- Well-solved by the *Baum-Welch* (known as *forward-backward*) algorithm and *EM* (*Expectation-Maximization*) algorithm
 - Iterative update and improvement
 - Based on Maximum Likelihood (ML) criterion

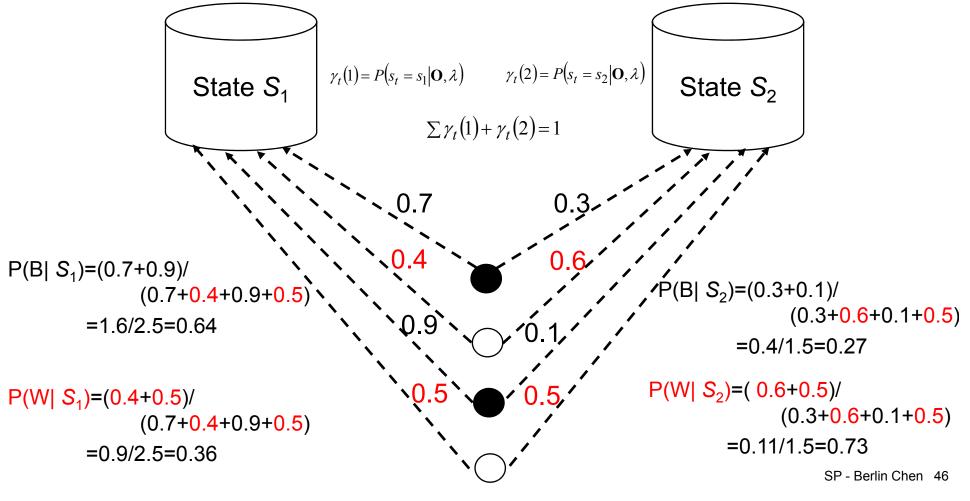
Maximum Likelihood (ML) Estimation: A Schematic Depiction (1/2)

- Hard Assignment
 - Given the data follow a multinomial distribution

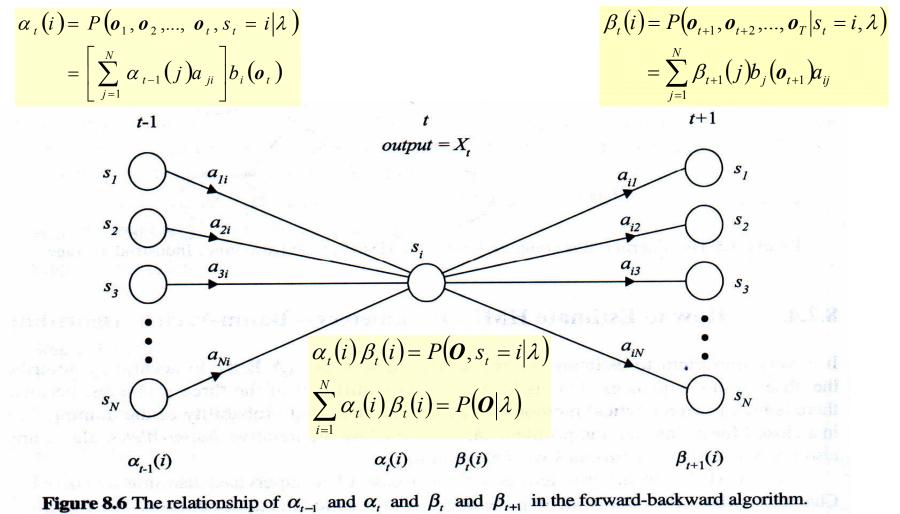


Maximum Likelihood (ML) Estimation: A Schematic Depiction (1/2)

- Soft Assignment
 - Given the data follow a multinomial distribution
 - Maximize the likelihood of the data given the alignment



• Relationship between the forward and backward variables



• Define a new variable:

$$\boldsymbol{\xi}_{t}(i,j) = P(\boldsymbol{s}_{t}=i,\boldsymbol{s}_{t+1}=j|\boldsymbol{O},\boldsymbol{\lambda})$$

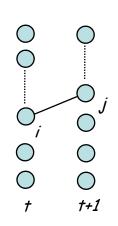
- Probability being at state *i* at time *t* and at state *j* at time *t*+1

$$\begin{aligned} \xi_t(i,j) &= \frac{P(s_t = i, s_{t+1} = j, \mathbf{O} | \boldsymbol{\lambda})}{P(\mathbf{O} | \boldsymbol{\lambda})} \\ &= \frac{\alpha_t(i) a_{ij} b_j(\mathbf{o}_{t+1}) \beta_{t+1}(j)}{P(\mathbf{O} | \boldsymbol{\lambda})} = \frac{\alpha_t(i) a_{ij} b_j(\mathbf{o}_{t+1}) \beta_{t+1}(j)}{\sum_{m=1}^N \sum_{n=1}^N \alpha_t(m) a_{mn} b_n(\mathbf{o}_{t+1}) \beta_{t+1}(n)} \end{aligned}$$

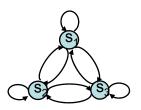
• Recall the posteriori probability variable:

$$\gamma_{t}(i) = P\left(s_{t} = i | \boldsymbol{O}, \lambda\right)$$
Note : $\gamma_{t}(i)$ also can be represented as
$$\frac{\alpha_{t}(r) \rho_{t}(r)}{\sum_{m=1}^{N} \alpha_{t}(m) \beta_{t}(m)}$$

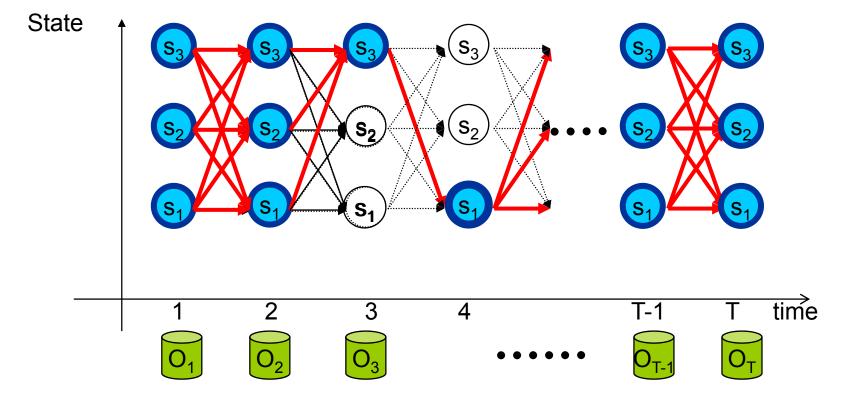
$$\gamma_{t}(i) = \sum_{j=1}^{N} P\left(s_{t} = i, s_{t+1} = j | \boldsymbol{O}, \lambda\right) = \sum_{j=1}^{N} \xi_{t}(i, j) \quad (\text{for } t < T)$$
SP - Berlin Chen 48



 $\alpha(i) B(i)$



• $P(s_3 = 3, s_4 = 1, \mathbf{O} \mid \lambda) = \alpha_3(3)^* a_{31}^* b_1(o_4)^* \beta_1(4)$



• $\xi_t(i, j) = P(s_t = i, s_{t+1} = j | \boldsymbol{O}, \boldsymbol{\lambda})$ $\sum_{t=1}^{T-1} \xi_t(i, j) = \text{expected number of transitions from state } i \text{ to state } j \text{ in } \boldsymbol{O}$

•
$$\gamma_t(i) = P(s_t = i | \boldsymbol{O}, \lambda)$$

 $\sum_{t=1}^{T-1} \gamma_t(i) = \sum_{t=1}^{T-1} \sum_{j=1}^{N} \xi_t(i, j) = \text{expected number of transitions from state } i \text{ in } \boldsymbol{O}$

• A set of reasonable re-estimation formula for $\{A, \pi\}$ is

 $\overline{\pi}_i$ = expected frequency (number of times) in state *i* at time *t* = 1 = $\gamma_1(i)$

 $\overline{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state } i} = \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \gamma_t(i)}$

Formulae for Single Training Utterance

- A set of reasonable re-estimation formula for {B} is
 - For discrete and finite observation $b_j(\mathbf{v}_k) = P(\mathbf{o}_t = \mathbf{v}_k | \mathbf{s}_t = j)$

 $\overline{b}_{j}(\mathbf{v}_{k}) = \overline{P}(\mathbf{o} = \mathbf{v}_{k}|s = j) = \frac{\text{expected number of times in state } j \text{ and observing symbol } \mathbf{v}_{k}}{\text{expected number of times in state } j} = \frac{\sum_{t=1}^{T} \gamma_{t}(j)}{\sum_{t=1}^{T} \gamma_{t}(j)}$

- For continuous and infinite observation $b_j(\mathbf{v}) = f_{\mathbf{o}|\mathbf{s}}(\mathbf{o}_t = \mathbf{v}|\mathbf{s}_t = j)$,

$$\overline{b}_{j}(\mathbf{v}) = \sum_{k=1}^{M} \overline{c}_{jk} N(\mathbf{v}; \overline{\boldsymbol{\mu}}_{jk}, \overline{\boldsymbol{\Sigma}}_{jk}) = \sum_{k=1}^{M} \overline{c}_{jk} \left(\frac{1}{(\sqrt{2\pi})^{L} |\overline{\boldsymbol{\Sigma}}_{jk}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{v} - \overline{\boldsymbol{\mu}}_{jk})^{T} \overline{\boldsymbol{\Sigma}}_{jk}^{-1} (\mathbf{v} - \overline{\boldsymbol{\mu}}_{jk})\right) \right)$$

Modeled as a mixture of multivariate Gaussian distributions

$$p(A|B) = \frac{p(A,B)}{P(B)}$$

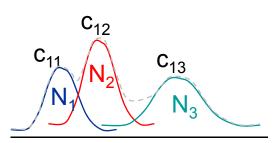
- For continuous and infinite observation (Cont.)
 - Define a new variable $\gamma_t(j,k)$
 - $\gamma_t(j,k)$ is the probability of being in state *j* at time *t* with the *k*-th mixture component accounting for **o**_t

$$\begin{split} \gamma_t(j,k) &= P(s_t = j, m_t = k | \mathbf{O}, \boldsymbol{\lambda}) \\ &= P(s_t = j | \mathbf{O}, \boldsymbol{\lambda}) P(m_t = k | s_t = j, \mathbf{O}, \boldsymbol{\lambda}) \\ &= \gamma_t(j) P(m_t = k | s_t = j, \mathbf{O}, \boldsymbol{\lambda}) \\ &= \gamma_t(j) \frac{p(m_t = k, \mathbf{O} | s_t = j, \boldsymbol{\lambda})}{p(\mathbf{O} | s_t = j, \boldsymbol{\lambda})} \\ &= \gamma_t(j) \frac{P(m_t = k | s_t = j, \boldsymbol{\lambda}) p(\mathbf{O} | s_t = j, m_t = k, \boldsymbol{\lambda})}{p(\mathbf{O} | s_t = j, \boldsymbol{\lambda})} \end{split}$$

 $= \gamma_t(j) \frac{P(m_t = k | s_t = j, \lambda) p(\mathbf{o}_t | s_t = j, m_t = k, \lambda)}{p(\mathbf{o}_t | s_t = j, \lambda)}$

 $= \left[\frac{\alpha_t(j)\beta_t(j)}{\sum_{i=1}^{N}\alpha_t(s)\beta_t(s)}\right] \left[\frac{c_{jk}N(\mathbf{o}_t;\mathbf{\mu}_{jk},\mathbf{\Sigma}_{jk})}{\sum_{i=1}^{M}c_{jm}N(\mathbf{o}_t;\mathbf{\mu}_{jm},\mathbf{\Sigma}_{jm})}\right]$

(observation - independent assumption is applied)



Distribution for State 1

Note:
$$\gamma_t(j) = \sum_{m=1}^M \gamma_t(j,m)$$

- For continuous and infinite observation (Cont.)

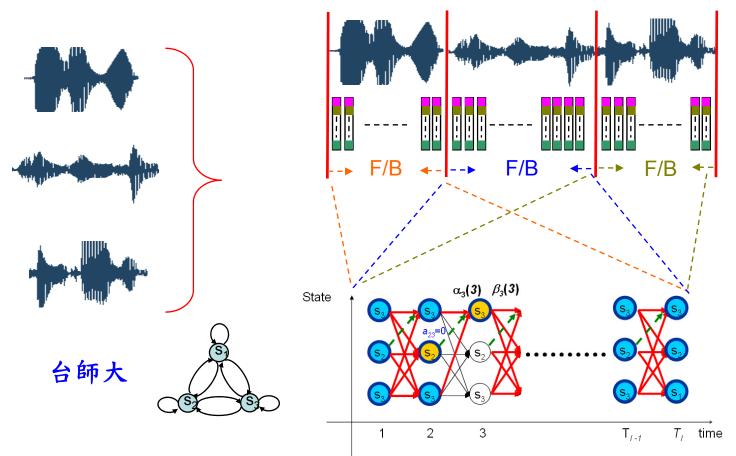
$$\overline{c}_{jk} = \frac{\text{expected number of times in state } j \text{ and mixture } k}{\text{expected number of times in state } j} = \frac{\sum_{t=1}^{T} \gamma_t(j,k)}{\sum_{t=1}^{T} \sum_{m=1}^{M} \gamma_t(j,m)}$$

 $\overline{\boldsymbol{\mu}}_{jk} = \text{weighted average (mean) of observations at state } j \text{ and mixture } k = \frac{\sum_{t=1}^{T} \gamma_t(j,k) \cdot \boldsymbol{o}_t}{\sum_{t=1}^{T} \gamma_t(j,k)}$

$$\overline{\boldsymbol{\Sigma}}_{jk} = \text{weighted covariance of observations at state } j \text{ and mixture } k$$
$$= \frac{\sum_{t=1}^{T} \gamma_t(j,k) \cdot (\boldsymbol{o}_t - \overline{\boldsymbol{\mu}}_{jk}) (\boldsymbol{o}_t - \overline{\boldsymbol{\mu}}_{jk})^t}{\sum_{t=1}^{T} \gamma_t(j,k)}$$

Formulae for Single Training Utterance

• Multiple Training Utterances



- For continuous and infinite observation (Cont.)

 $\overline{\pi}_i$ = expected frequency (number of times) in state *i* at time $(t = 1) = \frac{1}{L} \sum_{l=1}^{L} \gamma_1^l(i)$



$$\overline{c}_{jk} = \frac{\text{expected number of times in state } j \text{ and mixture } k}{\text{expected number of times in state } j} = \frac{\sum_{l=1}^{L} \sum_{t=1}^{T_l} \gamma_t^l(j,k)}{\sum_{l=1}^{L} \sum_{t=1}^{T_l} M \gamma_t^l(j,m)}$$

$$\overline{\mu}_{jk} = \text{weighted average (mean) of observations at state } j \text{ and mixture } k = \frac{\sum_{l=1}^{L} \sum_{t=1}^{T_l} \gamma_t^l(j,k) \cdot \mathbf{o}_t}{\sum_{l=1}^{L} \sum_{t=1}^{T_l} \gamma_t^l(j,k)}$$

 $\overline{\Sigma}_{jk} = \text{weighted covariance of observations at state } j \text{ and mixture } k$ $= \frac{\sum_{l=1}^{L} \sum_{t=1}^{T_l} \gamma_t^l(j,k) \cdot \left(\mathbf{o}_t - \overline{\mathbf{\mu}}_{jk}\right) \left(\mathbf{o}_t - \overline{\mathbf{\mu}}_{jk}\right)^t}{\sum_{l=1}^{L} \sum_{t=1}^{T_l} \gamma_t^l(j,k)}$

Formulae for Multiple (L) Training Utterances

- For discrete and finite observation (cont.)

 $\overline{\pi_i}$ = expected frequency (number of times) in state *i* at time $(t = 1) = \frac{1}{L} \sum_{l=1}^{L} \gamma_1^l(i)$

$$\overline{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state } i} = \frac{\sum_{l=1}^{L} \sum_{t=1}^{T_l-1} \zeta_t^l(i,j)}{\sum_{l=1}^{L} \sum_{t=1}^{T_l-1} \gamma_t^l(i)}$$

$$\overline{b}_{j}(\mathbf{v}_{k}) = \overline{P}(\mathbf{o} = \mathbf{v}_{k}|s = j) = \frac{\text{expected number of times in state } j \text{ and observing symbol } \mathbf{v}_{k}}{\text{expected number of times in state } j} = \frac{\sum_{l=1}^{L} \sum_{t=1}^{T_{l}} \gamma_{t}^{l}(j)}{\sum_{l=1}^{L} \sum_{t=1}^{T_{l}} \gamma_{t}^{l}(j)}$$

Formulae for Multiple (L) Training Utterances

Semicontinuous HMMs

- The HMM state mixture density functions are tied together across all the models to form a set of shared kernels
 - The semicontinuous or tied-mixture HMM

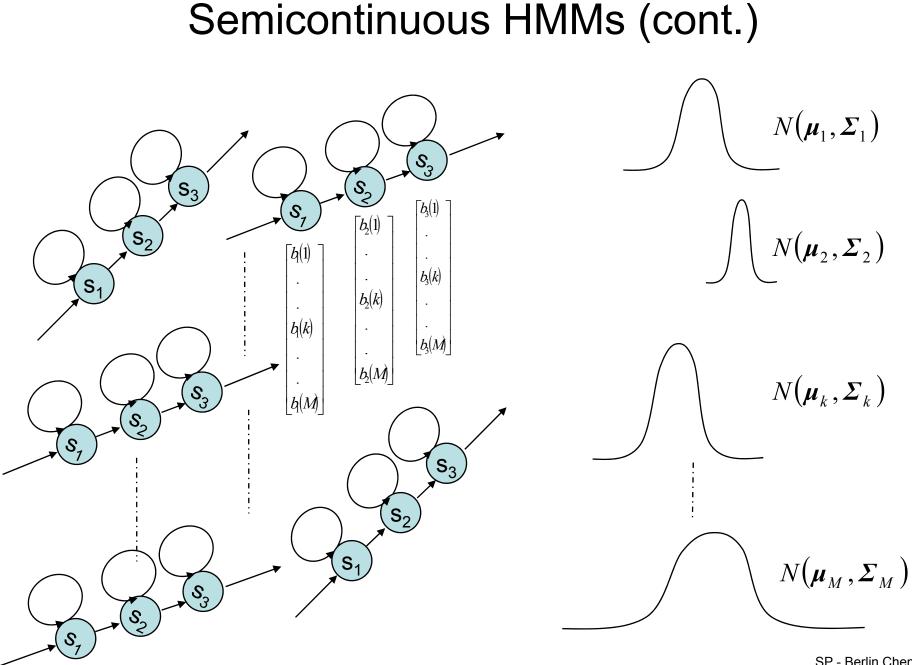
$$b_{j}(\boldsymbol{o}) = \sum_{k=1}^{M} b_{j}(k) f(\boldsymbol{o}|v_{k}) = \sum_{k=1}^{M} b_{j}(k) N(\boldsymbol{o}, \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

state output Probability of state *j*

k-th mixture weight *k*-th mixture density function or *k*-th codeword t of state *j* (shared across HMMs, *M* is very large) (discrete, model-dependent)

A combination of the discrete HMM and the continuous HMM

- A combination of *discrete* model-dependent weight coefficients and *continuous* model-independent codebook probability density functions
- Because *M* is large, we can simply use the *L* most significant values $f(o|v_k)$
 - Experience showed that *L* is $1 \sim 3\%$ of *M* is adequate
- Partial tying of $f(\boldsymbol{o}|\boldsymbol{v}_k)$ for different phonetic class



HMM Topology

- Speech is time-evolving non-stationary signal
 - Each HMM state has the ability to capture some quasi-stationary segment in the non-stationary speech signal
 - A *left-to-right* topology is a natural candidate to model the speech signal (also called the "beads-on-a-string" model)

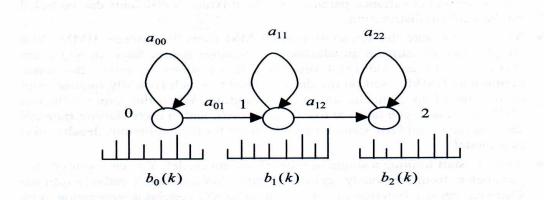


Figure 8.8 A typical hidden Markov model used to model phonemes. There are three states (0-2) and each state has an associated output probability distribution.

 It is general to represent a phone using 3~5 states (English) and a syllable using 6~8 states (Mandarin Chinese)

Initialization of HMM

- A good initialization of HMM training : <u>Segmental K-Means Segmentation into States</u>
 - Assume that we have a training set of observations and an initial estimate of all model parameters
 - Step 1 : The set of training observation sequences is segmented into states, based on the initial model (finding the optimal state sequence by *Viterbi* Algorithm)
 - Step 2 :
 - For discrete density HMM (using M-codeword codebook)

 $\overline{b}_{j}(k) = \frac{\text{the number of vectors with codebook index } k \text{ in state } j}{\text{the number of vectors in state } j}$

• For continuous density HMM (M Gaussian mixtures per state)

 \Rightarrow cluster the observation vectors within each state *j* into a set of *M* clusters

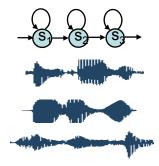
 \overline{w}_{im} = number of vectors classified in cluster *m* of state *j*

divided by the number of vectors in state *j*

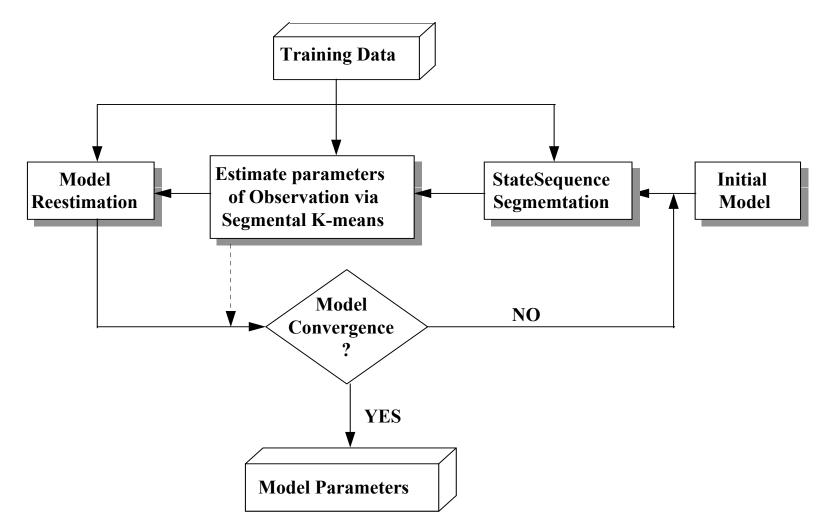
 $\overline{\mu}_{jm}$ = sample mean of the vectors classified in cluster *m* of state *j*

 $\overline{\Sigma}_{jm}$ = sample covariance matrix of the vectors classified in cluster *m* of state *j*

Step 3: Evaluate the model score
 If the difference between the previous and current model scores is greater than a threshold, go back to Step 1, otherwise stop, the initial model is generated

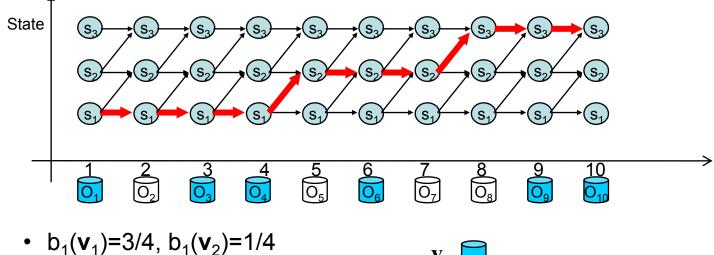


Initialization of HMM (cont.)



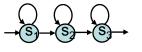
Initialization of HMM (cont.)

- An example for discrete HMM
 - 3 states and 2 codeword



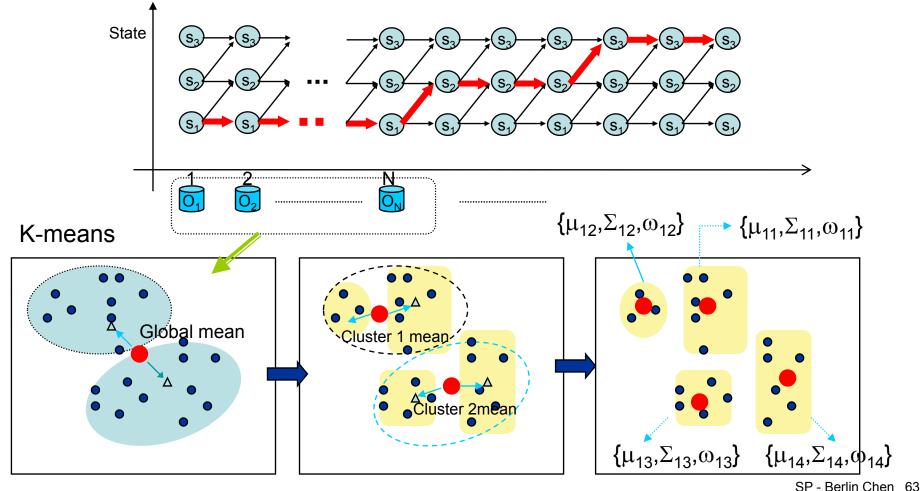
- $b_2(\mathbf{v}_1)=1/3$, $b_2(\mathbf{v}_2)=2/3$
- $b_3(\mathbf{v}_1)=2/3$, $b_3(\mathbf{v}_2)=1/3$

 $\mathbf{v}_1 \square$ $\mathbf{v}_2 \square$



Initialization of HMM (cont.)

- An example for Continuous HMM
 - 3 states and 4 Gaussian mixtures per state



Known Limitations of HMMs (1/3)

- The assumptions of conventional HMMs in Speech Processing
 - The state duration follows an exponential distribution
 - Don't provide adequate representation of the temporal structure of speech

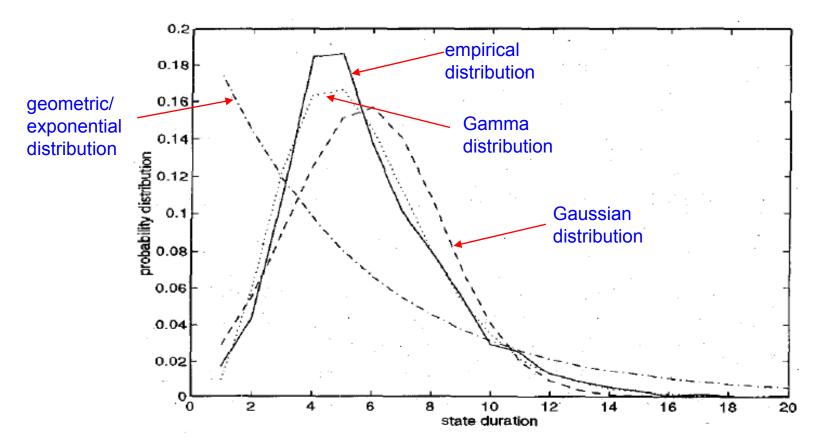
 $d_i(t) = a_{ii}^{t-1}(1 - a_{ii})$

- First-order (Markov) assumption: the state transition depends only on the origin and destination
- Output-independent assumption: all observation frames are dependent on the state that generated them, not on neighboring observation frames

Researchers have proposed a number of techniques to address these limitations, albeit these solution have not significantly improved speech recognition accuracy for practical applications.

Known Limitations of HMMs (2/3)

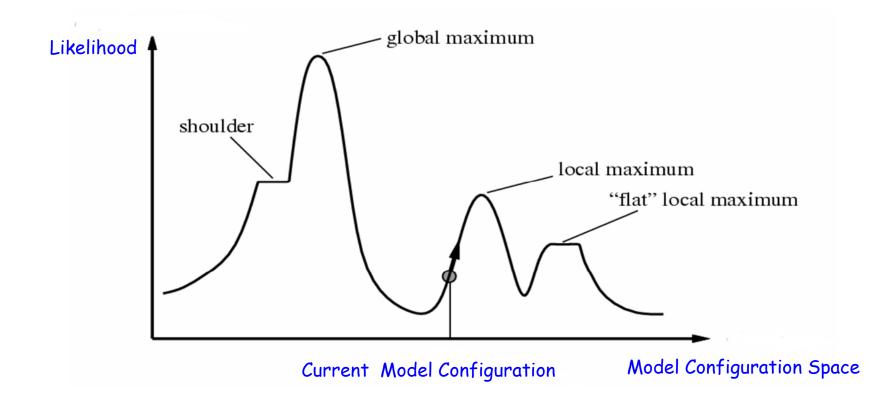
• Duration modeling



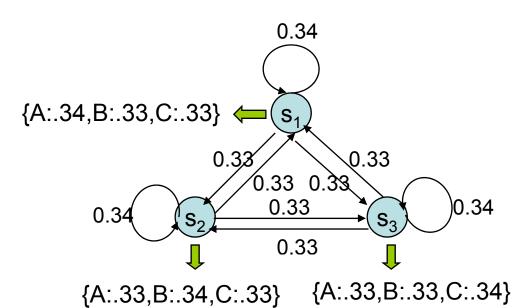
Duration distributions for the seventh state of the word "seven:" empirical distribution (solid line); Gauss fit (dashed line); gamma fit (dotted line); and (d) geometric fit (dash-dot line).

Known Limitations of HMMs (3/3)

• The HMM parameters trained by the *Baum-Welch* algorithm and *EM* algorithm were only locally optimized



Homework-2 (1/2)



TrainSet 1:

- 1. ABBCABCAABC
- 2. ABCABC
- 3. ABCA ABC
- 4. BBABCAB
- 5. BCAABCCAB
- 6. CACCABCA
- 7. CABCABCA
- 8. CABCA
- 9. CABCA

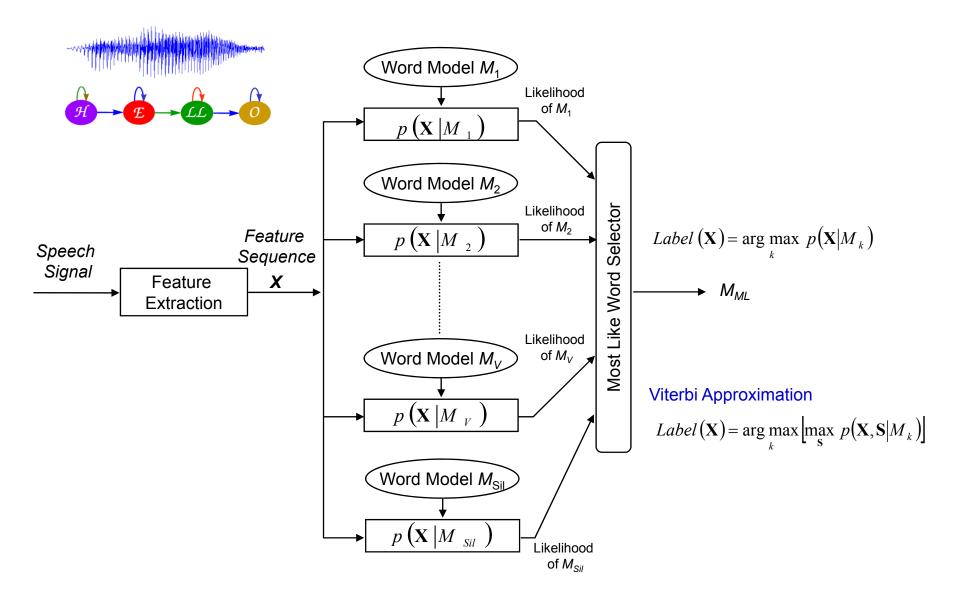
TrainSet 2:

- 1. BBBCCBC
- 2. CCBABB
- 3. AACCBBB
- 4. BBABBAC
- 5. CCA ABBAB
- 6. BBBCCBAA
- 7. ABBBBABA
- 8. CCCCC
- 9. BBAAA

Homework-2 (2/2)

- P1. Please specify the model parameters after the first and 50th iterations of Baum-Welch training
- P2. Please show the recognition results by using the above training sequences as the testing data (The so-called inside testing).
 *You have to perform the recognition task with the HMMs trained from the first and 50th iterations of Baum-Welch training, respectively
- P3. Which class do the following testing sequences belong to? ABCABCCAB AABABCCCCBBB
- P4. What are the results if Observable Markov Models were instead used in P1, P2 and P3?

Isolated Word Recognition

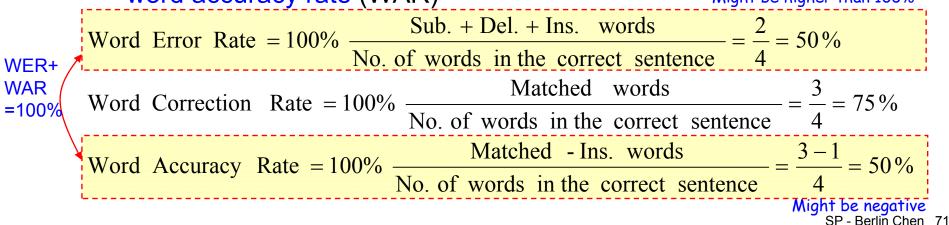


Measures of ASR Performance (1/8)

- Evaluating the performance of automatic speech recognition (ASR) systems is critical, and the Word Recognition Error Rate (WER) is one of the most important measures
- There are typically three types of word recognition errors
 - Substitution
 - An incorrect word was substituted for the correct word
 - Deletion
 - A correct word was omitted in the recognized sentence
 - Insertion
 - An extra word was added in the recognized sentence
- How to determine the minimum error rate?

Measures of ASR Performance (2/8)

- Calculate the WER by aligning the correct word string against the recognized word string
 - A maximum substring matching problem
 - Can be handled by dynamic programming deleted
- Example: Correct : "the effect is clear" Recognized: "effect is not clear"
 - matched inserted matched
 - Error analysis: one deletion and one insertion
 - Measures: word error rate (WER), word correction rate (WCR), word accuracy rate (WAR)
 Might be higher than 100%

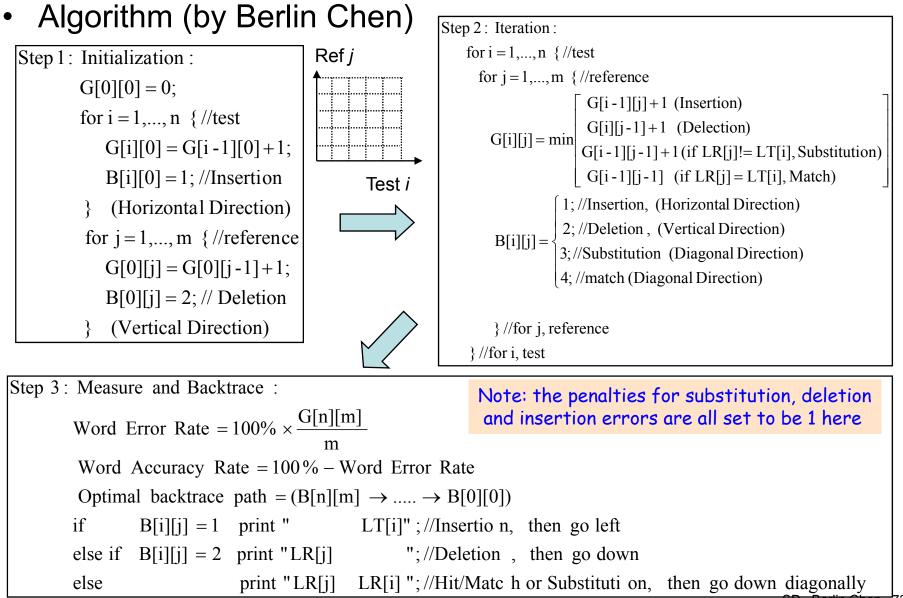


Measures of ASR Performance (3/8)

• A Dynamic Programming Algorithm (Textbook)

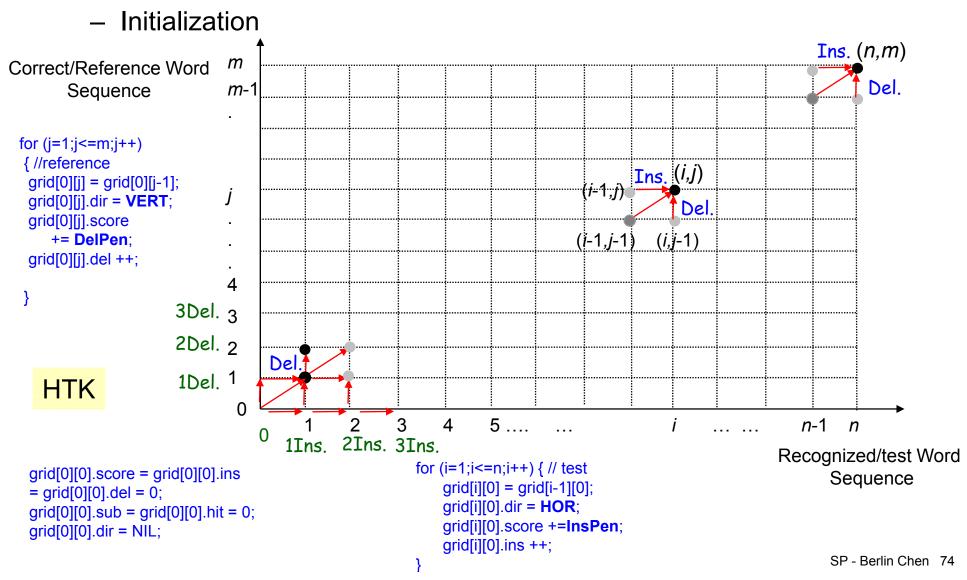
ALGORITHM 9.1: ALGORITHM TO MEASURE THE WORD ERROR RATE **Step 1:** Initialization R[0,0] = 0 $R[i, j] = \infty$ if (i < 0) or (j < 0) B[0,0] = 0Step 2: Iteration for i = 1, ..., n { //denotes for the word length of the correct/reference sentence for j = 1, ..., m { //denotes for the word length of the recognized/test sentence R[i-1, j]+1 (deletion) minimum word R[i-1, j-1] (match)/hit error alignment R[i-1, j-1]+1 (substitution) at the a grid [*i*,*j*] $R[i, j] = \min$ R[i, j-1]+1 (insertion) 1 if deletion Test i kinds of 2 if insertion
 3 if match /hit $B[i,j] = \left\{ \right.$ alignment if substitution Ref i Step 3: Backtracking and termination word error rate = $100\% \times \frac{R(n,m)}{100\%}$ optimal backward path = $(s_1, s_2, \dots, 0)$ where $s_1 = B[n,m]$, $s_t = \begin{bmatrix} B[i-1,j] \text{ if } s_{t-1} = 1 \\ B[i,j-1] \text{ if } s_{t-1} = 2 \\ B[i-1,j-1] \text{ if } s_{t-1} = 3 \text{ or } 4 \end{bmatrix}$ for $t = 2, \dots$ until $s_t = 0$

Measures of ASR Performance (4/8)

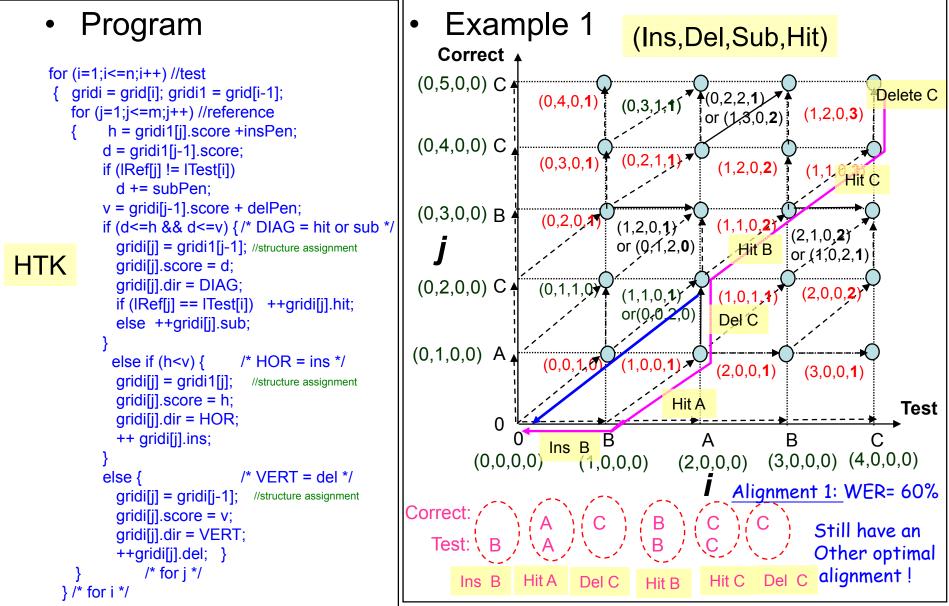


Measures of ASR Performance (5/8)

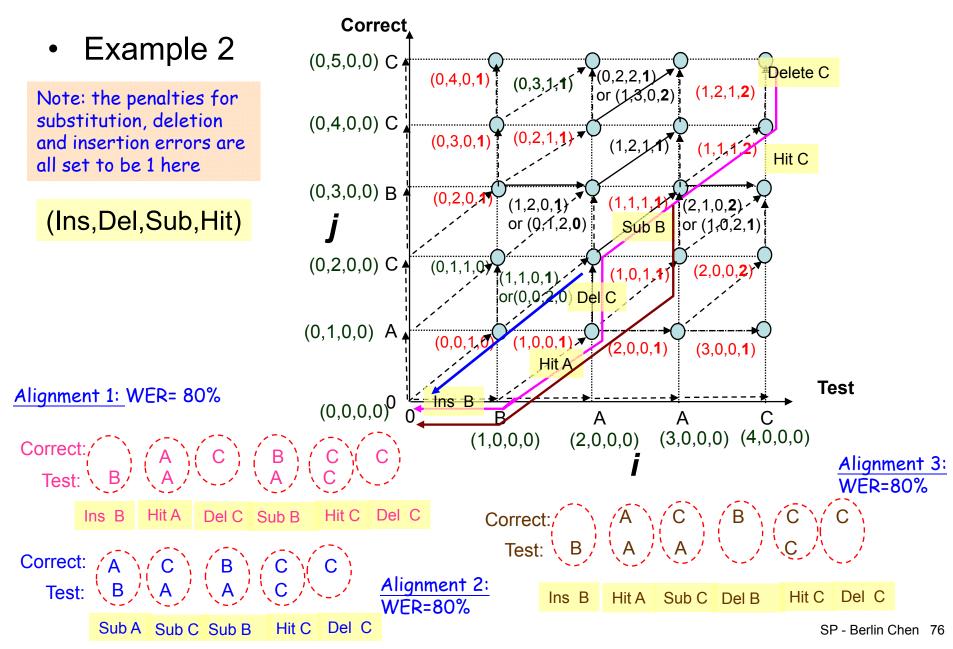
A Dynamic Programming Algorithm



Measures of ASR Performance (6/8)



Measures of ASR Performance (7/8)



Measures of ASR Performance (8/8)

• Two common settings of different penalties for substitution, deletion, and insertion errors

```
/* HTK error penalties */
subPen = 10;
delPen = 7;
insPen = 7;
/* NIST error penalties*/
subPenNIST = 4;
delPenNIST = 3;
insPenNIST = 3;
```

Self-Exercise (1/2)

Measures of ASR Performance

Reference

.

ASR Output

.

Self-Exercise (2/2)

- 506 BN stories of ASR outputs
 - Report the CER (character error rate) of the first one, 100, 200, and 506 stories
 - The result should show the number of substitution, deletion and insertion errors

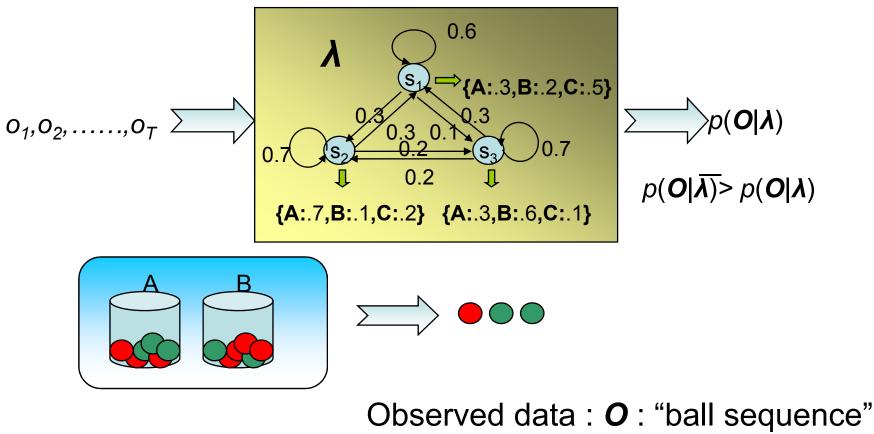
Overall Results
SENT: %Correct=0.00 [H=0, S=1, N=1] WORD: %Corr=81.52, Acc=81.52 [H=75, D=4, S=13, I=0, N=92] ===================================
Overall Results
SENT: %Correct=0.00 [H=0, S=100, N=100] WORD: %Corr=87.66, Acc=86.83 [H=10832, D=177, S=1348, I=102, N=12357] ====================================
Overall Results
SENT: %Correct=0.00 [H=0, S=200, N=200] WORD: %Corr=87.91, Acc=87.18 [H=22657, D=293, S=2824, I=186, N=25774] ===================================
Overall Results
SENT: %Correct=0.00 [H=0, S=506, N=506] WORD: %Corr=86.83, Acc=86.06 [H=57144, D=829, S=7839, I=504, N=65812] ====================================

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Symbols for Mathematical Operations

Αα	alpha	Ιιiota	Pρ	rho
Ββ	beta	К к карра	Σσ	sigma
	gamma	$\Lambda_{-}\lambda_{-}$ lambda	Ττ	tau
Εε	epsilon	Μμ mu	Υυ	upsilon
Δδ	delta	N v nu	Φφ	pĥi
Zζ	zeta	Ξξxi	Xχ	chi
Нη	eta	O o omicron	Ψψ	psi
Θθ	theta	Плрі	ωΩ	omega

The EM Algorithm (1/7)



Latent data : **S** : "bottle sequence"

Parameters to be estimated to maximize $\log P(O|\lambda)$ $\lambda = \{P(A), P(B), P(B|A), P(A|B), P(R|A), P(G|A), P(R|B), P(G|B)\}$

The EM Algorithm (2/7)

- Introduction of EM (Expectation Maximization):
 - Why EM?
 - Simple optimization algorithms for likelihood function relies on the intermediate variables, called latent data In our case here, *the state sequence* is the latent data
 - Direct access to the data necessary to estimate the parameters is impossible or difficult
 In our case here, it is almost impossible to estimate {*A*,*B*, π} without consideration of the *state sequence*
 - Two Major Steps :
 - *E* : expectation with respect to the latent data using the current estimate of the parameters and conditioned on the $E\left[\bullet\right]_{s\mid\lambda,o}$ observations
 - M: provides a new estimation of the parameters according to Maximum likelihood (ML) or Maximum A Posterior (MAP) Criteria

The EM Algorithm (3/7)

ML and MAP

• Estimation principle based on observations:

$$\mathbf{x} = (\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n) \iff \mathbf{X} = \{X_1, X_2, \dots, X_n\}$$

- The Maximum Likelihood (ML) Principle find the model parameter Φ so that the likelihood $p(x|\Phi)$ is maximum

for example, if $\Phi = \{\mu, \Sigma\}$ is the parameters of a multivariate normal distribution, and **X** is *i.i.d.* (independent, identically distributed), then the ML estimate of $\Phi = \{\mu, \Sigma\}$ is

$$\boldsymbol{\mu}_{ML} = \frac{1}{n} \sum_{i=1}^{n} \boldsymbol{x}_{i} , \boldsymbol{\Sigma}_{ML} = \frac{1}{n} \sum_{i=1}^{n} (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{ML}) (\boldsymbol{x}_{i} - \boldsymbol{\mu}_{ML})^{t}$$

- The Maximum A Posteriori (MAP) Principle find the model parameter Φ so that the likelihood $p(\Phi|x)$ is maximum SP-Berlin Chen 83

The EM Algorithm (4/7)

- The EM Algorithm is important to HMMs and other learning techniques
 - Discover new model parameters to maximize the log-likelihood of incomplete data $\log P(O|\lambda)$ by iteratively maximizing the expectation of log-likelihood from complete data $\log P(O, S|\lambda)$
- Firstly, using scalar (discrete) random variables to introduce the EM algorithm
 - The observable training data $oldsymbol{O}$
 - We want to maximize $P(\boldsymbol{O}|\boldsymbol{\lambda})$, $\boldsymbol{\lambda}$ is a parameter vector
 - The hidden (unobservable) data $m{S}$
 - E.g. the component probabilities (or densities) of observable data *O*, or the underlying state sequence in HMMs

The EM Algorithm (5/7)

- Assume we have λ and estimate the probability that each s occurred in the generation of *O*
- Pretend we had in fact observed a complete data pair (O, S) with frequency proportional to the probability $P(O, S | \lambda)$, to computed a new $\overline{\lambda}$, the maximum likelihood estimate of λ
- Does the process converge?
- Algorithm unknown model setting

 $P(\mathbf{O}, \mathbf{S} | \overline{\lambda}) = P(\mathbf{S} | \mathbf{O}, \overline{\lambda}) P(\mathbf{O} | \overline{\lambda})$ Bayes' rule complete data likelihood incomplete data likelihood

 Log-likelihood expression and expectation taken over S $\log P(\boldsymbol{O}|\boldsymbol{\overline{\lambda}}) = \log P(\boldsymbol{O}, \boldsymbol{S}|\boldsymbol{\overline{\lambda}}) - \log P(\boldsymbol{S}|\boldsymbol{O}, \boldsymbol{\overline{\lambda}})$ $\log P(\boldsymbol{O}|\boldsymbol{\overline{\lambda}}) = \sum \left[P(\boldsymbol{S}|\boldsymbol{O}, \boldsymbol{\lambda}) \log P(\boldsymbol{O}|\boldsymbol{\overline{\lambda}}) \right]$ take expectation over **S** $=\sum_{\boldsymbol{\alpha}} \left[P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda}) \log P(\boldsymbol{O},\boldsymbol{S}|\boldsymbol{\overline{\lambda}}) \right] - \sum_{\boldsymbol{\alpha}} \left[P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda}) \log P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\overline{\lambda}}) \right]$

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The EM Algorithm (6/7)

- Algorithm (Cont.)
 - We can thus express $\log P(\mathbf{O}|\overline{\lambda})$ as follows $\log P(\mathbf{O}|\overline{\lambda})$ $= \sum_{s} \left[P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{O}, \mathbf{S}|\overline{\lambda}) \right] - \sum_{s} \left[P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{S}|\mathbf{O}, \overline{\lambda}) \right]$ $= Q(\lambda, \overline{\lambda}) - H(\lambda, \overline{\lambda})$

where

$$Q(\lambda, \overline{\lambda}) = \sum_{\mathbf{S}} \left[P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{O}, \mathbf{S}|\overline{\lambda}) \right]$$
$$H(\lambda, \overline{\lambda}) = \sum_{\mathbf{S}} \left[P(\mathbf{S}|\mathbf{O}, \lambda) \log P(\mathbf{S}|\mathbf{O}, \overline{\lambda}) \right]$$

• We want $\log P(\mathbf{O}|\overline{\lambda}) \ge \log P(\mathbf{O}|\lambda)$ $\log P(\mathbf{O}|\overline{\lambda}) - \log P(\mathbf{O}|\lambda)$ $= [Q(\lambda,\overline{\lambda}) - H(\lambda,\overline{\lambda})] - [Q(\lambda,\lambda) - H(\lambda,\lambda)]$ $= Q(\lambda,\overline{\lambda}) - Q(\lambda,\lambda) - H(\lambda,\overline{\lambda}) + H(\lambda,\lambda)$

The EM Algorithm (7/7)

• $-H(\lambda,\overline{\lambda}) + H(\lambda,\lambda)$ has the following property $-H(\lambda,\overline{\lambda})+H(\lambda,\lambda)$ Kullbuack-Leibler (KL) distance $= -\sum_{\boldsymbol{S}} \left[P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda}) \log \frac{P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\overline{\lambda}})}{P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda})} \right]$ $\geq \sum_{\boldsymbol{S}} \left| P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda}) \left(1 - \frac{P(\boldsymbol{S}|\boldsymbol{O},\bar{\boldsymbol{\lambda}})}{P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda})} \right) \right| \quad (\because \log x \leq x - 1) \quad \text{Jensen's inequality}$ $=\sum_{\bar{\boldsymbol{\lambda}}}\left[P(\boldsymbol{S}|\boldsymbol{O},\boldsymbol{\lambda})-P(\boldsymbol{S}|\boldsymbol{O},\bar{\boldsymbol{\lambda}})\right]$ = 0 $\therefore -H(\lambda,\overline{\lambda}) + H(\lambda,\lambda) \geq 0$ - Therefore, for maximizing $\log P(\boldsymbol{O}|\boldsymbol{\overline{\lambda}})$, we only need to maximize the Q-function (auxiliary function) Expectation of the complete $Q(\lambda, \overline{\lambda}) = \sum_{\alpha} \left[P(S|O, \lambda) \log P(O, S|\overline{\lambda}) \right]$ data log likelihood with respect to the latent state sequences

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EM Applied to Discrete HMM Training (1/5)

- Apply EM algorithm to iteratively refine the HMM parameter vector $\lambda = (A, B, \pi)$
 - By maximizing the auxiliary function

$$Q(\lambda, \overline{\lambda}) = \sum_{S} \left[P(S|O, \lambda) \log P(O, S|\overline{\lambda}) \right]$$
$$= \sum_{S} \left[\frac{P(O, S|\lambda)}{P(O|\lambda)} \log P(O, S|\overline{\lambda}) \right]$$

- Where $P(\boldsymbol{O}, \boldsymbol{S} | \boldsymbol{\lambda})$ and $P(\boldsymbol{O}, \boldsymbol{S} | \boldsymbol{\lambda})$ can be expressed as $P(\boldsymbol{O}, \boldsymbol{S} | \boldsymbol{\lambda}) = \pi_{s_1} \left[\prod_{t=1}^{T-1} a_{s_t s_{t+1}} \right] \left[\prod_{t=1}^{T} b_{s_t}(\boldsymbol{o}_t) \right]$ $\log P(\boldsymbol{O}, \boldsymbol{S} | \boldsymbol{\lambda}) = \log \pi_{s_1} + \sum_{t=1}^{T-1} \log a_{s_t s_{t+1}} + \sum_{t=1}^{T} \log b_{s_t}(\boldsymbol{o}_t)$ $\log P(\boldsymbol{O}, \boldsymbol{S} | \boldsymbol{\lambda}) = \log \pi_{s_1} + \sum_{t=1}^{T-1} \log a_{s_t s_{t+1}} + \sum_{t=1}^{T} \log b_{s_t}(\boldsymbol{o}_t)$

EM Applied to Discrete HMM Training (2/5)

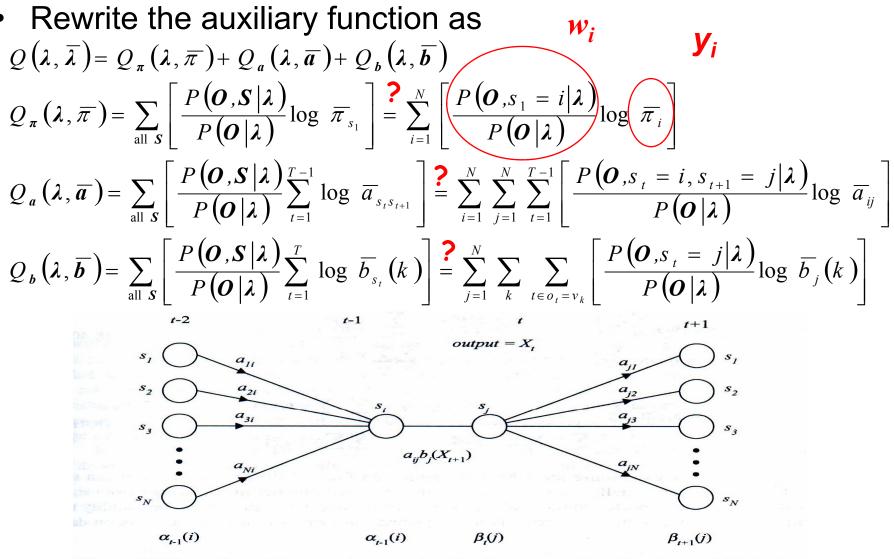


Figure 8.7 Illustration of the operations required for the computation of $\gamma_t(i, j)$, which is the probability of taking the transition from state *i* to state *j* at time *t*.

EM Applied to Discrete HMM Training (3/5)

- The auxiliary function contains three independent terms, π_i , a_{ij} and $b_j(k)$
 - Can be maximized individually
 - All of the same form

$$F(\mathbf{y}) = g(y_1, y_2, ..., y_N) = \sum_{j=1}^N w_j \log y_j, \text{ where } \sum_{j=1}^N y_j = 1, \text{ and } y_j \ge 0$$

$$F(\mathbf{y}) \text{ has maximum value when : } y_j = \frac{w_j}{\sum_{j=1}^N w_j}$$

EM Applied to Discrete HMM Training (4/5)

• **Proof**: Apply Lagrange Multiplier

By applying Lagrange Multiplier
$$\ell$$

Suppose that $F = \sum_{j=1}^{N} w_j \log y_j = \sum_{j=1}^{N} w_j \log y_j + \ell \left(\sum_{j=1}^{N} y_j - 1 \right)$
 $\frac{\partial F}{\partial y_j} = \frac{w_j}{y_j} + \ell = 0 \Rightarrow \ell = -\frac{w_j}{y_j} \forall j$
Constraint
 $\ell \sum_{j=1}^{N} y_j = -\sum_{j=1}^{N} w_j \Rightarrow \ell = -\sum_{j=1}^{N} w_j$
 $\therefore y_j = \frac{w_j}{\sum_{j=1}^{N} w_j}$

Lagrange Multiplier: http://www.slimy.com/~steuard/teaching/tutorials/Lagrange.html

EM Applied to Discrete HMM Training (5/5)

• The new model parameter set $\overline{\lambda} = (\overline{\pi}, \overline{A}, \overline{B})$ can be expressed as:

$$\overline{\pi}_{i} = \frac{P\left(\boldsymbol{0}, s_{1} = i | \lambda\right)}{P\left(\boldsymbol{0} | \lambda\right)} = \gamma_{1}\left(i\right)$$

$$\overline{a}_{ij} = \frac{\sum_{t=1}^{T-1} P\left(\boldsymbol{0}, s_{t} = i, s_{t+1} = j | \lambda\right)}{\sum_{t=1}^{T-1} P\left(\boldsymbol{0}, s_{t} = i | \lambda\right)} = \frac{\sum_{t=1}^{T-1} \xi_{t}\left(i, j\right)}{\sum_{t=1}^{T-1} \gamma_{t}\left(i\right)}$$

$$\overline{b}_{i}\left(k\right) = \frac{\sum_{t=1}^{T} P\left(\boldsymbol{0}, s_{t} = i | \lambda\right)}{\sum_{t=1}^{T} P\left(\boldsymbol{0}, s_{t} = i | \lambda\right)} = \frac{\sum_{t=1}^{T} \gamma_{t}\left(i\right)}{\sum_{t=1}^{T} \gamma_{t}\left(i\right)}$$

EM Applied to Continuous HMM Training (1/7)

- Continuous HMM: the state observation does not come \bullet from a finite set, but from a continuous space
 - The difference between the discrete and continuous HMM lies. in a different form of state output probability
 - Discrete HMM requires the quantization procedure to map observation vectors from the continuous space to the discrete space
- Continuous Mixture HMM
 - The state observation distribution of HMM is modeled by multivariate Gaussian mixture density functions (*M* mixtures)

$$b_{j}(\boldsymbol{o}) = \sum_{k=1}^{M} c_{jk} b_{jk}(\boldsymbol{o})$$

$$= \sum_{k=1}^{M} c_{jk} N(\boldsymbol{o}; \boldsymbol{\mu}_{jk}, \boldsymbol{\Sigma}_{jk}) = \sum_{k=1}^{M} c_{jk} \left(\frac{1}{(\sqrt{2\pi})^{L} |\boldsymbol{\Sigma}_{jk}|^{1/2}} \exp\left(-\frac{1}{2}(\boldsymbol{o}-\boldsymbol{\mu}_{jk})^{L} \boldsymbol{\Sigma}_{jk}(\boldsymbol{o}-\boldsymbol{\mu}_{jk})\right) \right)$$

$$\underbrace{\sum_{k=1}^{M} c_{jk}}_{k=l} c_{jk} = l$$

$$\underbrace{\sum_{k=1}^{M} c_{jk}}_{SP - Berlin Chen}$$

$$W_{i2}$$
 W_{i3} N_1 N_2 N_3

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EM Applied to Continuous HMM Training (2/7)

• Express $b_j(o)$ with respect to each single mixture component $b_{jk}(o)$ Note: $\prod_{i=1}^{T} (\sum_{k=1}^{M} a_{ik})$

$$p(\mathbf{O}, \mathbf{S} | \boldsymbol{\lambda}) = \pi_{s_1} \left\{ \prod_{t=1}^{T-1} a_{s_t s_{t+1}} \right\} \left\{ \prod_{t=1}^{T} b_{s_t}(\mathbf{o}_t) \right\} \xrightarrow{= (a_{11} + a_{12} + \dots + a_{1M})(a_{21} + a_{22} + \dots + a_{2M})\dots(a_{T1} + a_{T2} + \dots + a_{TM})}{= \sum_{k_t=1}^{M} \sum_{k_t=1}^$$

K : one of the possible mixture component sequence along with the state sequence **S**

$$p(\mathbf{O}|\boldsymbol{\lambda}) = \sum_{\mathbf{S}} \sum_{\mathbf{K}} p(\mathbf{O}, \mathbf{S}, \mathbf{K}|\boldsymbol{\lambda})$$

EM Applied to Continuous HMM Training (3/7)

• Therefore, an auxiliary function for the EM algorithm can be written as:

$$Q(\lambda, \overline{\lambda}) = \sum_{\mathbf{S}} \sum_{\mathbf{K}} \left[P(\mathbf{S}, \mathbf{K} | \mathbf{O}, \lambda) \log p(\mathbf{O}, \mathbf{S}, \mathbf{K} | \overline{\lambda}) \right]$$

$$= \sum_{\mathbf{S}} \sum_{\mathbf{K}} \left[\frac{p(\mathbf{O}, \mathbf{S}, \mathbf{K} | \lambda)}{p(\mathbf{O} | \lambda)} \log p(\mathbf{O}, \mathbf{S}, \mathbf{K} | \overline{\lambda}) \right]$$

$$\log p(\mathbf{O}, \mathbf{S}, \mathbf{K} | \overline{\lambda}) = \log \overline{\pi}_{s_{1}} + \sum_{t=1}^{T-1} \log \overline{a}_{s_{t}s_{t+1}} + \sum_{t=1}^{T} \log \overline{b}_{s_{t}k_{t}}(\mathbf{0}_{t}) + \sum_{t=1}^{T} \log \overline{c}_{s_{t}k_{t}}$$

$$\bigcup$$

$$Q(\lambda, \overline{\lambda}) = Q_{\pi}(\lambda, \overline{\pi}) + Q_{a}(\lambda, \overline{a}) + Q_{b}(\lambda, \overline{b}) + Q_{c}(\lambda, \overline{c})$$

initial state transition Gaussian mixture probabilities probabilities functions

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EM Applied to Continuous HMM Training (4/7)

• The only difference we have when compared with Discrete HMM training

$$\begin{aligned}
& \gamma_t(j,k) \\
Q_b(\lambda, \overline{b}) = \sum_{t=1}^T \left\{ \left[\sum_{j=1}^N \sum_{k=1}^M \left| P(s_t = j, k_t = k | \boldsymbol{O}, \lambda) \right| \log \overline{b}_{jk}(\boldsymbol{o}_t) \right\} \\
& Q_c(\lambda, \overline{c}) = \sum_{t=1}^T \left\{ \left[\sum_{j=1}^N \sum_{k=1}^M P(s_t = j, k_t = k | \boldsymbol{O}, \lambda) \right] \log \overline{c}_{jk}(\boldsymbol{o}_t) \right\}
\end{aligned}$$

EM Applied to Continuous HMM Training (5/7) Let $\gamma_{t}(j,k) = \sum_{k=1}^{M} P(s_{t} = j, k_{t} = k | \boldsymbol{0}, \lambda)$ $\overline{b}_{jk}(\boldsymbol{o}_{t}) = N(\boldsymbol{o}_{t}; \overline{\boldsymbol{\mu}}_{jk}, \overline{\boldsymbol{\Sigma}}_{jk}) = \frac{1}{(2\pi)^{L/2} |\overline{\boldsymbol{\Sigma}}_{jk}|^{1/2}} \exp\left(-\frac{1}{2}(\boldsymbol{o}_{t} - \overline{\boldsymbol{\mu}}_{jk})\overline{\boldsymbol{\Sigma}}_{jk}^{-1}(\boldsymbol{o}_{t} - \overline{\boldsymbol{\mu}}_{jk})\right)$

$$\log \overline{b}_{jk}(o_{t}) = -\frac{L}{2} \cdot \log (2\pi) + \frac{1}{2} \cdot \log |\overline{\Sigma}_{jk}^{-1}| - \left(\frac{1}{2}(o_{t} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}(o_{t} - \overline{\mu}_{jk})\right)$$

$$\frac{\partial \log \overline{b}_{jk}(o_{t})}{\partial \overline{\mu}_{jk}} = \overline{\Sigma}_{jk}^{-1}(o_{t} - \overline{\mu}_{jk})$$

$$\frac{\partial Q_{b}(\lambda, \overline{b})}{\partial \overline{\mu}_{jk}} = \frac{\partial \sum_{t=1}^{T} \left\{ \left[\sum_{j=1}^{N} \sum_{k=1}^{M} \gamma_{t}(j,k) \log \overline{b}_{jk}(o_{t})\right] \right\}}{\partial \overline{\mu}_{jk}}$$

$$\Rightarrow \sum_{t=1}^{T} \left\{ \gamma_{t}(j,k) \overline{\Sigma}_{jk}^{-1}(o_{t} - \overline{\mu}_{jk}) \right\} = 0$$

$$\Rightarrow \overline{\mu}_{jk} = \frac{\sum_{t=1}^{T} \left[\gamma_{t}(j,k) \cdot o_{t} \right]}{\sum_{t=1}^{T} \gamma_{t}(j,k)}$$
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$$\log \overline{b}_{jk}(o_{\tau}) = -\frac{L}{2} \cdot \log (2\pi) - \frac{1}{2} \cdot \log |\overline{\Sigma}_{jk}| - \left(\frac{1}{2}(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})\right)$$

$$\frac{\partial \log \overline{b}_{jk}(o_{\tau})}{\partial(\overline{\Sigma}_{jk})} = -\left[\frac{1}{2} \cdot \left[\overline{\Sigma}_{jk}^{-1}\right]^{-1} + \left[\overline{\Sigma}_{jk}\right] \cdot \overline{\Sigma}_{jk}^{-1} - \left(\overline{\Sigma}_{jk}^{-1}\right] \frac{1}{2}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}\right)\right]$$

$$= -\frac{1}{2} \cdot \left[\overline{\Sigma}_{jk}^{-1} - \overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}\right]$$

$$\frac{d(a^{T}X^{-1}b)}{dX} = -X^{T}ab^{T}X^{T}$$

$$\frac{d[det(X)]}{dX} = det(X) \cdot X^{-T}$$

$$and \Sigma_{jk} is symmetric here$$

$$\Rightarrow \sum_{\tau=1}^{T} \left\{\gamma_{\tau}(j,k)\left(-\frac{1}{2}\right) \cdot \left[\overline{\Sigma}_{jk}^{-1} - \overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}\right]\right\} = 0$$

$$\Rightarrow \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}^{-1} = \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}$$

$$\Rightarrow \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}^{-1} = \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}$$

$$\Rightarrow \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}\overline{\Sigma}_{jk}\overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1}$$

$$\Rightarrow \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}\overline{\Sigma}_{jk}\overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1}$$

$$\Rightarrow \sum_{\tau=1}^{T} \gamma_{\tau}(j,k)\overline{\Sigma}_{jk}\overline{\Sigma}_{jk}\overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1}(o_{\tau} - \overline{\mu}_{jk})(o_{\tau} - \overline{\mu}_{jk})\overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}^{-1} \overline{\Sigma}_{jk}$$

EM Applied to Continuous HMM Training (7/7)

• The new model parameter set for each mixture component and mixture weight can be expressed as:

$$\overline{\boldsymbol{\mu}}_{jk} = \frac{\sum_{t=1}^{T} \left[\frac{p(\mathbf{O}, s_t = j, k_t = k | \boldsymbol{\lambda})}{p(\mathbf{O} | \boldsymbol{\lambda})} \mathbf{o}_t \right]}{\sum_{t=1}^{T} \frac{p(\mathbf{O}, s_t = j, k_t = k | \boldsymbol{\lambda})}{p(\mathbf{O} | \boldsymbol{\lambda})}} = \frac{\sum_{t=1}^{T} \left[\gamma_t(j, k) \mathbf{o}_t \right]}{\sum_{t=1}^{T} \gamma_t(j, k)}$$

$$\overline{\boldsymbol{\Sigma}}_{jk} = \frac{\sum_{t=1}^{T} \left[\frac{p(\mathbf{O}, s_t = j, k_t = k | \boldsymbol{\lambda})}{p(\mathbf{O} | \boldsymbol{\lambda})} (\mathbf{o}_t - \overline{\boldsymbol{\mu}}_{jk}) (\mathbf{o}_t - \overline{\boldsymbol{\mu}}_{jk}) \right]}{\sum_{t=1}^{T} \frac{p(\mathbf{O}, s_t = j, k_t = k | \boldsymbol{\lambda})}{p(\mathbf{O} | \boldsymbol{\lambda})}} = \frac{\sum_{t=1}^{T} \left[\gamma_t(j, k) (\mathbf{o}_t - \overline{\boldsymbol{\mu}}_{jk}) (\mathbf{o}_t - \overline{\boldsymbol{\mu}}_{jk}) \right]}{\sum_{t=1}^{T} \frac{p(\mathbf{O}, s_t = j, k_t = k | \boldsymbol{\lambda})}{p(\mathbf{O} | \boldsymbol{\lambda})}} = \frac{\sum_{t=1}^{T} \left[\gamma_t(j, k) (\mathbf{o}_t - \overline{\boldsymbol{\mu}}_{jk}) (\mathbf{o}_t - \overline{\boldsymbol{\mu}}_{jk}) \right]}{\sum_{t=1}^{T} \gamma_t(j, k)}$$

$$\overline{c}_{jk} = \frac{\sum_{t=1}^{T} \gamma_t(j, k)}{\sum_{t=1}^{T} \sum_{t=1}^{M} \gamma_t(j, k)}$$
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