

Hidden Markov Models for Speech Recognition

Berlin Chen

Department of Computer Science & Information Engineering National Taiwan Normal University

References:

- 1. Rabiner and Juang. Fundamentals of Speech Recognition. Chapter 6
- 2. Huang et. al. Spoken Language Processing. Chapters 4, 8
- 3. Rabiner. A Tutorial on Hidden Markov Models and Selected Applications in Speech Recognition. Proceedings of the IEEE, vol. 77, No. 2, February 1989
- 4. Gales and Young. The Application of Hidden Markov Models in Speech Recognition, Chapters 1-2, 2008
- Young. HMMs and Related Speech Recognition Technologies. Chapter 27, Springer Handbook of Speech Processing, Springer, 2007
- 6. J.A. Bilmes, A Gentle Tutorial of the EM Algorithm and its Application to Parameter Estimation for Gaussian Mixture and Hidden Markov Models, U.C. Berkeley TR-97-021

Hidden Markov Model (HMM): A Brief Overview

History

- Published in papers of Baum in late 1960s and early 1970s
- Introduced to speech processing by Baker (CMU) and Jelinek (IBM) in the 1970s (discrete HMMs)
- Then extended to continuous HMMs by Bell Labs

Assumptions

- Speech signal can be characterized as a parametric random (stochastic) process
- Parameters can be estimated in a precise, well-defined manner

Three fundamental problems

- Evaluation of probability (likelihood) of a sequence of observations given a specific HMM
- Determination of a best sequence of model states
- Adjustment of model parameters so as to best account for observed signals (or discrimination purposes)

Stochastic Process

- A stochastic process is a mathematical model of a probabilistic experiment that evolves in time and generates a sequence of numeric values
 - Each numeric value in the sequence is modeled by a random variable
 - A stochastic process is just a (finite/infinite) sequence of random variables

Examples

- (a) The sequence of recorded values of a speech utterance
- (b) The sequence of daily prices of a stock
- (c) The sequence of hourly traffic loads at a node of a communication network
- (d) The sequence of radar measurements of the position of an airplane

Observable Markov Model

- Observable Markov Model (Markov Chain)
 - First-order Markov chain of N states is a triple (S, A, π)
 - **S** is a set of *N* states
 - **A** is the $N \times N$ matrix of transition probabilities between states $P(s_t=j|s_{t-1}=i, s_{t-2}=k, \ldots) \approx P(s_t=j|s_{t-1}=i) \approx A_{ij}$ First-order and time-invariant assumptions
 - π is the vector of initial state probabilities $\pi_i = P(s_1 = j)$
 - The output of the process is the set of states at each instant of time, when each state corresponds to an observable event
 - The output in any given state is not random (*deterministic!*)
 - Too simple to describe the speech signal characteristics

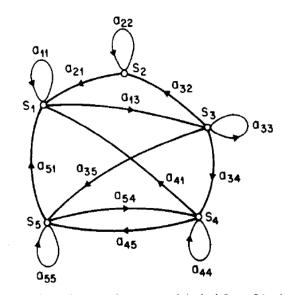
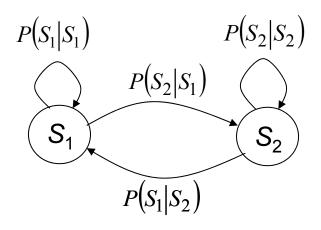
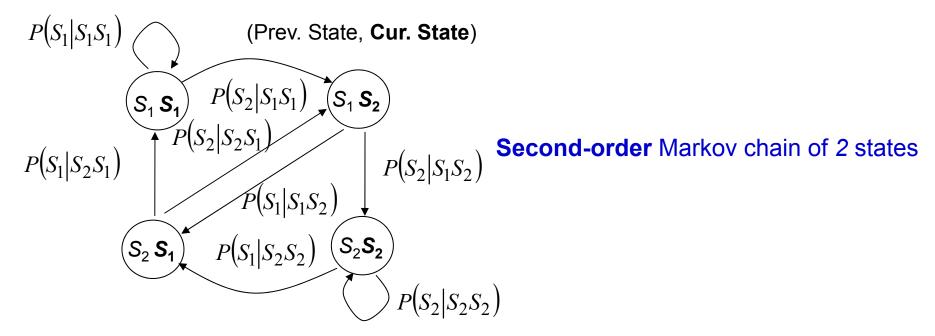


Fig. 1. A Markov chain with 5 states (labeled S_1 to S_5) with selected state transitions.



First-order Markov chain of 2 states

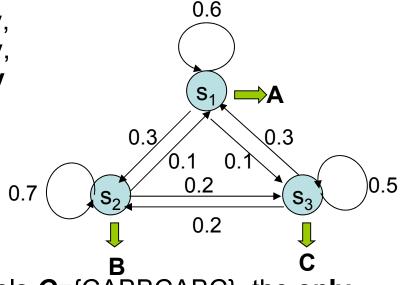


Example 1: A 3-state Markov Chain λ

State 1 generates symbol A **only**, State 2 generates symbol B **only**, and State 3 generates symbol C **only**

$$\mathbf{A} = \begin{bmatrix} 0.6 & 0.3 & 0.1 \\ 0.1 & 0.7 & 0.2 \\ 0.3 & 0.2 & 0.5 \end{bmatrix}$$

$$\pi = \begin{bmatrix} 0.4 & 0.5 & 0.1 \end{bmatrix}$$



- Given a sequence of observed symbols $O = \{CABBCABC\}$, the **only one** corresponding state sequence is $\{S_3S_1S_2S_2S_3S_1S_2S_3\}$, and the corresponding probability is

$$P(O|\lambda)$$

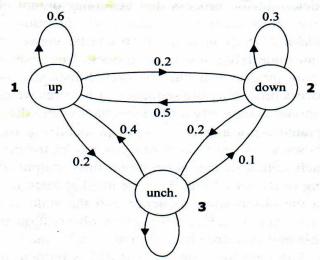
= $P(S_3)P(S_1|S_3)P(S_2|S_1)P(S_2|S_2)P(S_3|S_2)P(S_1|S_3)P(S_2|S_1)P(S_3|S_2)$
= $0.1 \times 0.3 \times 0.3 \times 0.7 \times 0.2 \times 0.3 \times 0.3 \times 0.2 = 0.00002268$

 Example 2: A three-state Markov chain for the Dow Jones Industrial average

state 1 - up (in comparison to the index of previous day)

state 2 - down (in comparison to the index of previous day)

state 3 – unchanged (in comparison to the index of previous day)



The probability of 5 consecutive up days

P(5 consecutive up days) = P(1,1,1,1,1)

$$= \pi_1 a_{11} a_{11} a_{11} a_{11} = 0.5 \times (0.6)^4 = 0.0648$$

Figure 8.1 A Markov chain for the Dow Jones Industrial average. Three states represent up, down, and unchanged, respectively.

The parameter for this Dow Jones Markov chain may include a state-transition probability matrix

$$A = \{a_{ij}\} = \begin{bmatrix} 0.6 & 0.2 & 0.2 \\ 0.5 & 0.3 & 0.2 \\ 0.4 & 0.1 & 0.5 \end{bmatrix} \qquad \pi = (\pi_i)^t = \begin{bmatrix} 0.5 \\ 0.2 \\ 0.3 \end{bmatrix}$$

and an initial state probability matrix

 Example 3: Given a Markov model, what is the mean occupancy duration of each state i

$$P_i(d)$$
 = probability mass function of duration d in state i

$$= (a_{ii})^{d-1} (1 - a_{ii})$$
 a geometric distribution

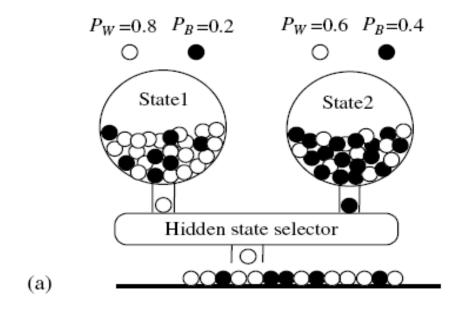
Expected number of duration in a state

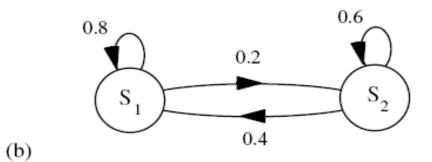
$$\overline{d}_{i} = \sum_{d=1}^{\infty} dP_{i}(d) = \sum_{d=1}^{\infty} d(a_{ii})^{d-1} (1 - a_{ii}) = (1 - a_{ii}) \frac{\partial}{\partial a_{ii}} \sum_{d=1}^{\infty} (a_{ii})^{d}$$

$$= (1 - a_{ii}) \frac{\partial}{\partial a_{ii}} \frac{1}{1 - a_{ii}} = \frac{1}{1 - a_{ii}}$$
Probability

Time (Duration)

Hidden Markov Model





(a) Illustration of a two-layered random process. (b) An HMM model of the process in (a).

- HMM, an extended version of Observable Markov Model
 - The observation is turned to be a probabilistic function (discrete or continuous) of a state instead of an one-to-one correspondence of a state
 - The model is a doubly embedded stochastic process with an underlying stochastic process that is not directly observable (hidden)
 - What is hidden? **The State Sequence!**According to the observation sequence, we are not sure which state sequence generates it!
- Elements of an HMM (the State-Output HMM) λ={S,A,B,π}
 - S is a set of N states
 - **A** is the $N \times N$ matrix of transition probabilities between states
 - B is a set of N probability functions, each describing the observation probability with respect to a state
 - $-\pi$ is the vector of initial state probabilities

- Two major assumptions
 - First order (Markov) assumption
 - The state transition depends only on the origin and destination
 - Time-invariant

$$P(s_t = j | s_{t-1} = i) \approx P(s_\tau = j | s_{\tau-1} = i) = P(j | i) = A_{i,j}$$

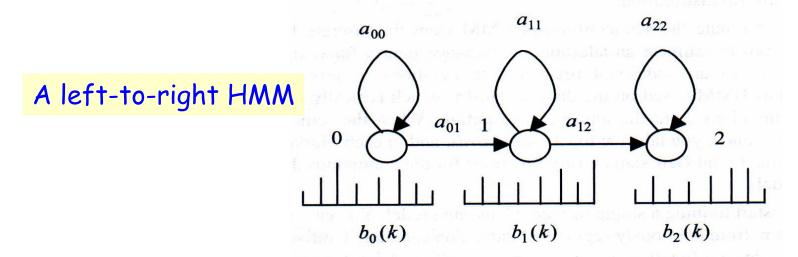
- Output-independent assumption
 - All observations are dependent on the state that generated them, not on neighboring observations

$$P(\mathbf{o}_t|s_t,\ldots,\mathbf{o}_{t-2},\mathbf{o}_{t-1},\mathbf{o}_{t+1},\mathbf{o}_{t+2}\ldots) = P(\mathbf{o}_t|s_t)$$

- Two major types of HMMs according to the observations
 - Discrete and finite observations:
 - The observations that all distinct states generate are finite in number

$$V=\{v_1, v_2, v_3, \dots, v_M\}, v_k \in R^L$$

- In this case, the set of observation probability distributions $B=\{b_j(\mathbf{v}_k)\}$, is defined as $b_j(\mathbf{v}_k)=P(\mathbf{o}_t=\mathbf{v}_k|s_t=j)$, $1 \le k \le M$, $1 \le j \le N$ \mathbf{o}_t : observation at time t, s_t : state at time t
 - \Rightarrow for state j, $b_i(\mathbf{v}_k)$ consists of only M probability values



- Two major types of HMMs according to the observations
 - Continuous and infinite observations:
 - The observations that all distinct states generate are infinite and continuous, that is, V={v| v∈R^d}
 - In this case, the set of observation probability distributions $B=\{b_j(\mathbf{v})\}$, is defined as $b_j(\mathbf{v})=f_{O|S}(\mathbf{o}_t=\mathbf{v}|s_t=j)$, $1 \le j \le N$ $\Rightarrow b_j(\mathbf{v})$ is a **continuous probability density function (pdf)** and is often a mixture of Multivariate Gaussian (Normal) Distributions

$$b_{j}(\mathbf{v}) = \sum_{k=1}^{M} w_{jk} \left(\frac{1}{(2\pi)^{\frac{d}{2}} |\mathbf{\Sigma}_{jk}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (\mathbf{v} - \mathbf{\mu}_{jk})^{t} \mathbf{\Sigma}_{jk}^{-1} (\mathbf{v} - \mathbf{\mu}_{jk}) \right) \right)$$
Mixture
Weight

Covariance
Matrix

Observation Vector

- Multivariate Gaussian Distributions
 - When $X=(x_1, x_2, ..., x_d)$ is a *d*-dimensional random vector, the multivariate Gaussian pdf has the form:

$$f(\mathbf{X} = \mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = N(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{\frac{d}{2}} |\boldsymbol{\Sigma}|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{t} \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right)$$

where μ is the L - dimensional mean vector, $\mu = E[\mathbf{x}]$

 Σ is the coverance matrix, $\Sigma = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^t] = E[\mathbf{x}\mathbf{x}^t] - \boldsymbol{\mu}\boldsymbol{\mu}^t$ and $|\Sigma|$ is the the determinant of Σ

The *i-j*th elevment
$$\sigma_{ij}$$
 of Σ , $\sigma_{ij} = E[(x_i - \mu_i)(x_j - \mu_j)] = E[x_i x_j] - \mu_i \mu_j$

- If x₁, x₂,..., x_d are independent, the covariance matrix is reduced to diagonal covariance
 - Viewed as d independent scalar Gaussian distributions
 - Model complexity is significantly reduced

Multivariate Gaussian Distributions

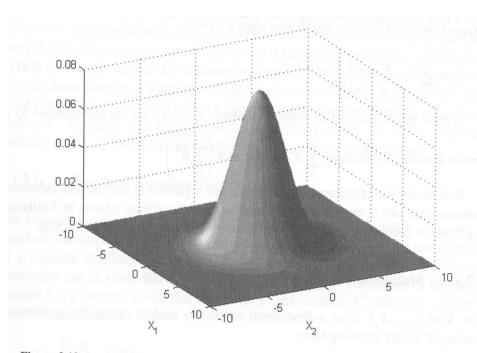


Figure 3.12 A two-dimensional multivariate Gaussian distribution with independent random variables x_1 and x_2 that have the same variance.

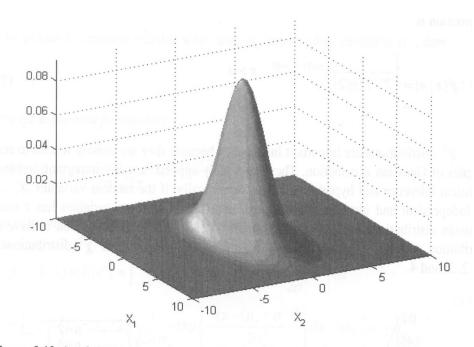
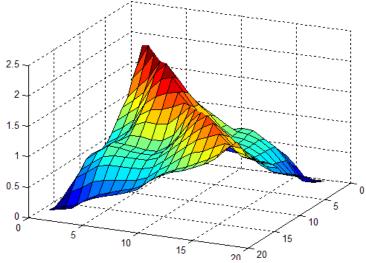
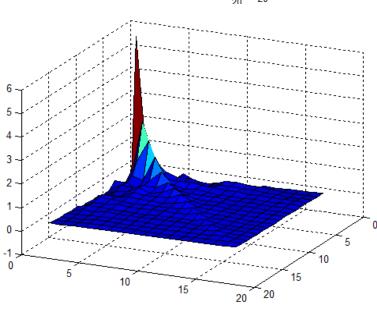


Figure 3.13 Another two-dimensional multivariate Gaussian distribution with independent random variable x_1 and x_2 which have different variances.

 Covariance matrix of the correlated feature vectors (Mel-frequency filter bank outputs)



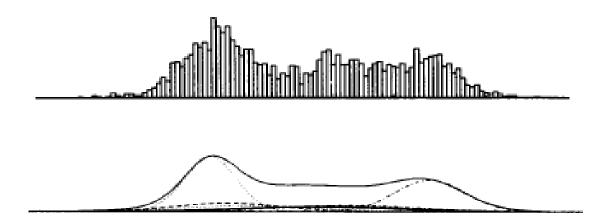
- Covariance matrix of the partially de-correlated feature vectors (MFCC without C_0)
 - MFCC: Mel-frequency cepstral coefficients



- Multivariate Mixture Gaussian Distributions (cont.)
 - More complex distributions with multiple local maxima can be approximated by Gaussian (a unimodal distribution) mixture

$$f(\mathbf{x}) = \sum_{k=1}^{M} w_k N_k(\mathbf{x}; \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \qquad \sum_{k=1}^{M} w_k = 1$$

Gaussian mixtures with enough mixture components can approximate any distribution



Example 4: a 3-state discrete HMM λ

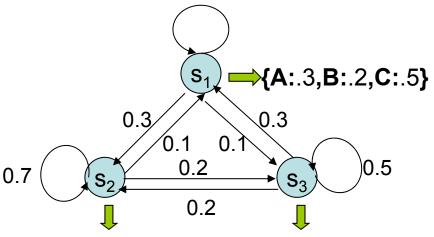
$$\mathbf{A} = \begin{bmatrix} 0.6 & 0.3 & 0.1 \\ 0.1 & 0.7 & 0.2 \\ 0.3 & 0.2 & 0.5 \end{bmatrix}$$

$$b_1(\mathbf{A}) = 0.3, b_1(\mathbf{B}) = 0.2, b_1(\mathbf{C}) = 0.5$$

$$b_2(\mathbf{A}) = 0.7, b_2(\mathbf{B}) = 0.1, b_2(\mathbf{C}) = 0.2$$

$$b_3(\mathbf{A}) = 0.3, b_3(\mathbf{B}) = 0.6, b_3(\mathbf{C}) = 0.1$$

 $\pi = \begin{bmatrix} 0.4 & 0.5 & 0.1 \end{bmatrix}$



{A:.7,B:.1,C:.2} {A:.3,B:.6,C:.1}

Given a sequence of observations O={ABC}, there are 27
 possible corresponding state sequences, and therefore the corresponding probability is

$$P(\mathbf{O}|\lambda) = \sum_{i=1}^{27} P(\mathbf{O}, \mathbf{S}_i | \lambda) = \sum_{i=1}^{27} P(\mathbf{O}|\mathbf{S}_i, \lambda) P(\mathbf{S}_i | \lambda), \quad \mathbf{S}_i : \text{state sequence}$$

$$E.g. \text{ when } \mathbf{S}_i = \{s_2 s_2 s_3\}, P(\mathbf{O}|\mathbf{S}_i, \lambda) = P(\mathbf{A}|s_2) P(\mathbf{B}|s_2) P(\mathbf{C}|s_3) = 0.7 * 0.1 * 0.1 = 0.007$$

$$P(\mathbf{S}_i | \lambda) = P(s_2) P(s_2 | s_2) P(s_3 | s_2) = 0.5 * 0.7 * 0.2 = 0.07$$
SP-Berlin Chen. 18

Notations:

- $O=\{o_1o_2o_3....o_7\}$: the observation (feature) sequence
- **S=** $\{s_1s_2s_3....s_7\}$: the state sequence
- $-\lambda$: model, for HMM, $\lambda = \{A,B,\pi\}$
- $-P(\mathbf{O}|\lambda)$: The probability of observing **O** given the model λ
- $-P(O|S,\lambda)$: The probability of observing **O** given λ and a state sequence **S** of λ
- $-P(O,S|\lambda)$: The probability of observing **O** and **S** given λ
- $-P(S|O,\lambda)$: The probability of observing S given O and λ

Useful formulas

– Bayes' Rule :

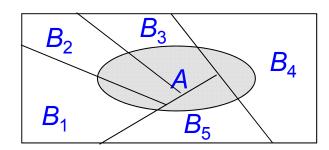
P(A|B) =
$$\frac{P(A,B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)}$$
 \Rightarrow $P(A|B,\lambda) = \frac{P(A,B|\lambda)}{P(B|\lambda)} = \frac{P(B|A,\lambda)P(A|\lambda)}{P(B|\lambda)}$ \Rightarrow \Rightarrow model describing the probability

$$P(A,B) = P(B|A)P(A) = P(A|B)P(B)$$
 chain rule

- Useful formulas (Cont.):
 - Total Probability Theorem

marginal probability
$$P(A) = \begin{cases} \sum_{all \ B} P(A,B) = \sum_{all \ B} P(A|B)P(B), & \text{if } B \text{ is disrete and disjoint} \\ \int_{B} f(A,B)dB = \int_{B} f(A|B)f(B)dB, & \text{if } B \text{ is continuous} \end{cases}$$

if x_1, x_2, \dots, x_n are independent, $\Rightarrow P(x_1, x_2, \dots, x_n) = P(x_1)P(x_2) \dots P(x_n)$



Venn Diagram

$$E_z[q(z)] = \begin{cases} \sum_k P(z=k)q(k), & z : \text{discrete} \\ \int_z f_{\mathbf{z}}(z)q(z)dz, & z : \text{continuous} \end{cases}$$
 Expectation

Three Basic Problems for HMM

- Given an observation sequence O=(o₁,o₂,....,o_T), and an HMM λ=(S,A,B,π)
 - Problem 1:

How to *efficiently* compute $P(\mathbf{O}|\lambda)$?

⇒ Evaluation problem

- Problem 2:

How to choose an optimal state sequence $S=(s_1, s_2, \ldots, s_T)$?

⇒ Decoding Problem

- Problem 3:

How to adjust the model parameter $\lambda = (A, B, \pi)$ to maximize $P(O|\lambda)$?

⇒ Learning / Training Problem

Given O and λ , find $P(O|\lambda)$ = Prob[observing O given λ]

- Direct Evaluation
 - Evaluating all possible state sequences of length T that generating observation sequence O

$$P\left(\boldsymbol{O}\mid\lambda\right) = \sum_{all\ \boldsymbol{S}} P\left(\boldsymbol{O}\ ,\boldsymbol{S}\mid\lambda\right) = \sum_{all\ \boldsymbol{S}} P\left(\boldsymbol{O}\mid\boldsymbol{S}\ ,\lambda\right) P\left(\boldsymbol{S}\mid\lambda\right)$$

- $P(S \mid \lambda)$: The probability of each path **S**
 - By Markov assumption (First-order HMM)

$$P\left(\mathbf{S} \mid \lambda\right) = P\left(s_1 \mid \lambda\right) \prod_{t=2}^{T} P\left(s_t \mid s_1^{t-1}, \lambda\right)$$
 By chain rule

$$\approx P(s_1|\lambda)\prod_{t=2}^{T} P(s_t|s_{t-1},\lambda)$$

$$= \pi_{s_1} a_{s_1 s_2} a_{s_2 s_3} \dots a_{s_{T-1} s_T}$$

By Markov assumption

- Direct Evaluation (cont.)
 - $-P(o|S,\lambda)$: The joint output probability along the path S
 - By output-independent assumption
 - The probability that a particular observation symbol/vector is emitted at time t depends only on the state s_t and is conditionally independent of the past observations

$$P\left(\boldsymbol{O} \mid \boldsymbol{S}, \boldsymbol{\lambda}\right) = P\left(\boldsymbol{o}_{1}^{T} \mid \boldsymbol{s}_{1}^{T}, \boldsymbol{\lambda}\right)$$

$$= P\left(\boldsymbol{o}_{1} \mid \boldsymbol{s}_{1}^{T}, \boldsymbol{\lambda}\right) \prod_{t=2}^{T} P\left(\boldsymbol{o}_{t} \mid \boldsymbol{o}_{1}^{t-1}, \boldsymbol{s}_{1}^{T}, \boldsymbol{\lambda}\right)$$

$$\approx \prod_{t=1}^{T} P\left(\boldsymbol{o}_{t} \mid \boldsymbol{s}_{t}, \boldsymbol{\lambda}\right) \quad \text{By output-independent assumption}$$

$$= \prod_{t=1}^{T} b_{s_{t}}\left(\boldsymbol{o}_{t}\right)$$

Direct Evaluation (Cont.)

$$P(\boldsymbol{o}_t|s_t,\lambda) = b_{s_t}(\boldsymbol{o}_t)$$

$$P(\mathbf{O}|\lambda) = \sum_{all \ S} P(\mathbf{S}|\lambda) P(\mathbf{O}|\mathbf{S},\lambda)$$

$$= \sum_{all \ S} \left(\left[\pi_{s_1} a_{s_1 s_2} a_{s_2 s_3} \dots a_{s_{T-1} s_T} \right] \left[b_{s_1} (\mathbf{o}_1) b_{s_2} (\mathbf{o}_2) \dots b_{s_T} (\mathbf{o}_T) \right] \right)$$

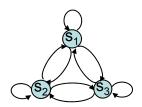
$$= \sum_{s_1, s_2, \dots, s_T} \pi_{s_1} b_{s_1} (\mathbf{o}_1) a_{s_1 s_2} b_{s_2} (\mathbf{o}_2) \dots a_{s_{T-1} s_T} b_{s_T} (\mathbf{o}_T)$$

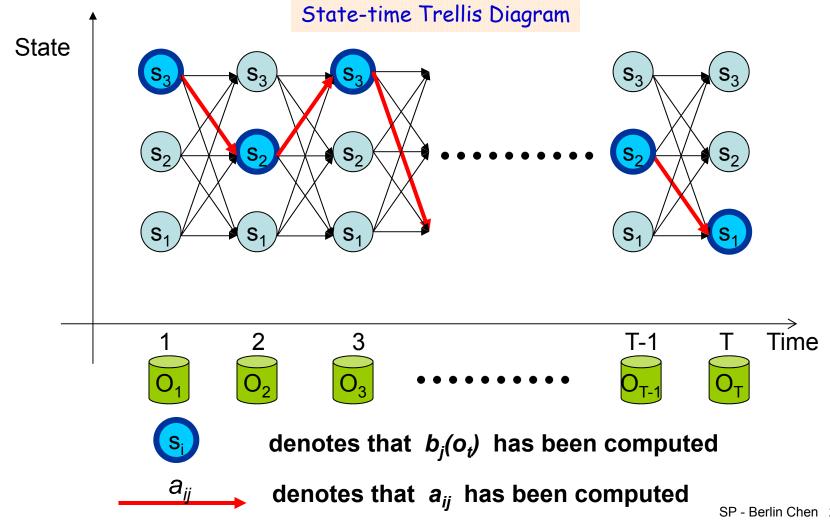
- Huge Computation Requirements: $O(N^T)$
 - Exponential computational complexity

Complexity :
$$(2T-1)N^T MUL \approx 2TN^T, N^T-1 ADD$$

- A more efficient algorithms can be used to evaluate $P(\boldsymbol{O}|\lambda)$
 - Forward/Backward Procedure/Algorithm

Direct Evaluation (Cont.)





- The Forward Procedure

- Based on the HMM assumptions, the calculation of $P(s_t|s_{t-1},\lambda)$ and $P(o_t|s_t,\lambda)$ involves only s_{t-1} , s_t and o_t , so it is possible to compute the likelihood with recursion on t
- Forward variable: $\alpha_t(i) = P(o_1 o_2 ... o_t, s_t = i | \lambda)$
 - The probability that the HMM is in state i at time t having generating partial observation $o_1 o_2 ... o_t$

- The Forward Procedure (cont.)

Algorithm

1. Initialization $\alpha_1(i) = \pi_i b_i(\boldsymbol{o}_1)$, $1 \le i \le N$

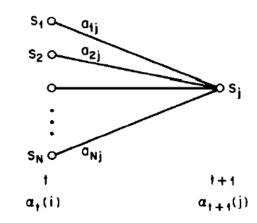
2. Induction
$$\alpha_{t+1}(j) = \left[\sum_{i=1}^{N} \alpha_t(i) a_{ij}\right] b_j(\boldsymbol{o}_{t+1}), \quad 1 \le t \le T-1, 1 \le j \le N$$

3. Termination
$$P(\mathbf{O}|\lambda) = \sum_{i=1}^{N} \alpha_{T}(i)$$

- Complexity: $O(N^2T)$

MUL :
$$N(N+1)(T-1)+N \approx N^2T$$

ADD : $(N-1)N(T-1)+(N-1)\approx N^2T$



- Based on the lattice (trellis) structure
 - Computed in a time-synchronous fashion from left-to-right, where each cell for time t is completely computed before proceeding to time t+1
- All state sequences, regardless how long previously, merge to N nodes (states) at each time instance t

- The Forward Procedure (cont.)

$$\alpha_{t}(j) = P(o_{1}o_{2}...o_{t}, s_{t} = j | \lambda)$$

$$= P(o_{1}o_{2}...o_{t} | s_{t} = j, \lambda)P(s_{t} = j | \lambda)$$

$$= P(o_{1}o_{2}...o_{t-1} | s_{t} = j, \lambda)P(o_{t} | s_{t} = j, \lambda)P(s_{t} = j | \lambda)$$

$$= P(o_{1}o_{2}...o_{t-1}, s_{t} = j | \lambda)P(o_{t} | s_{t} = j, \lambda)$$

$$= P(o_{1}o_{2}...o_{t-1}, s_{t} = j | \lambda)P(o_{t} | s_{t} = j, \lambda)$$

$$= P(o_{1}o_{2}...o_{t-1}, s_{t} = j | \lambda)P(o_{t} | s_{t} = j, \lambda)$$

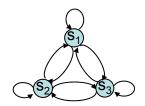
$$= P(o_{1}o_{2}...o_{t-1}, s_{t} = j | \lambda)P(o_{t} | s_{t} = j, \lambda)$$

$$= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i, s_{t} = j | \lambda)P(s_{t} = j | o_{1}o_{2}...o_{t-1}, s_{t-1} = i, \lambda)\right]b_{j}(o_{t})$$

$$= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i | \lambda)P(s_{t} = j | s_{t-1} = i, \lambda)\right]b_{j}(o_{t})$$

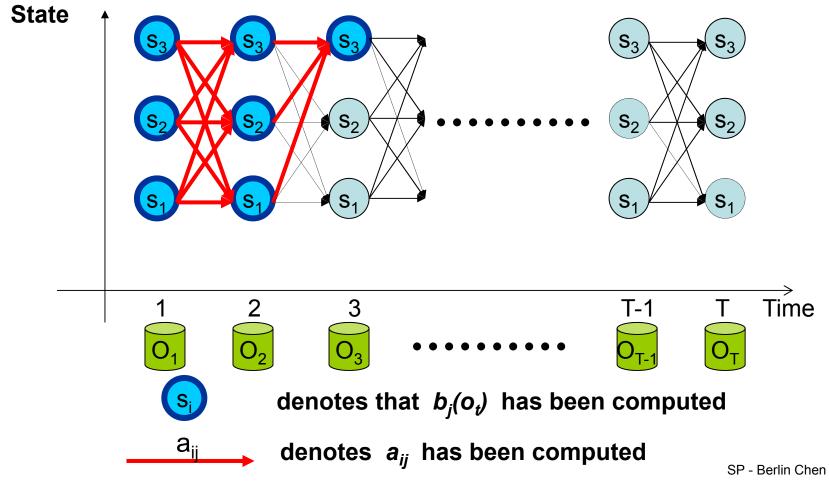
$$= \left[\sum_{i=1}^{N} P(o_{1}o_{2}...o_{t-1}, s_{t-1} = i | \lambda)P(s_{t} = j | s_{t-1} = i, \lambda)\right]b_{j}(o_{t})$$
first-order Markov assumption sp-Berlin Chen 28

- The Forward Procedure (cont.)



•
$$\alpha_3(3) = P(o_1, o_2, o_3, s_3 = 3 | \lambda)$$

= $[\alpha_2(1)^* a_{13} + \alpha_2(2)^* a_{23} + \alpha_2(3)^* a_{33}] b_3(\mathbf{o}_3)$



- The Forward Procedure (cont.)
- A three-state Hidden Markov Model for the Dow Jones Industrial average

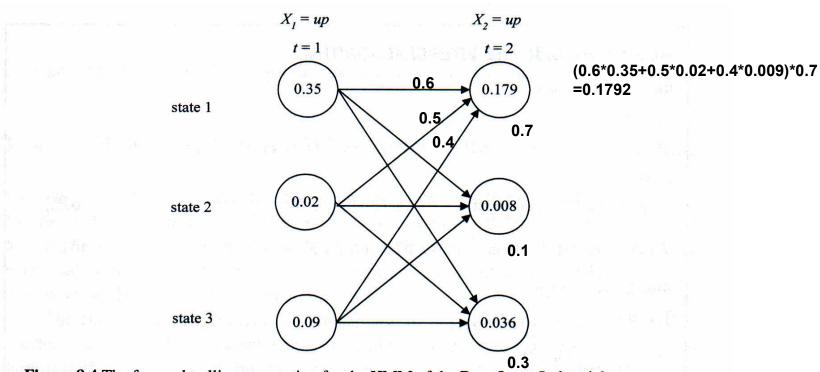


Figure 8.4 The forward trellis computation for the HMM of the Dow Jones Industrial average.

- The Backward Procedure

- Backward variable : $\beta_t(i) = P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, \dots, \mathbf{o}_T | s_t = i, \lambda)$
 - 1. Initialization : $\beta_{T}(i) = 1$, $1 \le i \le N$
 - 2. Induction: $\beta_t(i) = \sum_{j=1}^{N} a_{ij} b_j(\mathbf{o}_{t+1}) \beta_{t+1}(j), \ 1 \le t \le T-1, 1 \le i \le N$

3. Termination :
$$P(\mathbf{O}|\lambda) = \sum_{j=1}^{N} \pi_j b_j(\mathbf{o}_1) \beta_1(j)$$

Complexity MUL:
$$2N^2(T-1) + 2N \approx N^2T$$
;

ADD:
$$(N-1)N(T-1) + N \approx N^2T$$

- Backward Procedure (cont.)

• Why
$$P(\mathbf{O}, s_t = i | \lambda) = \alpha_t(i) \beta_t(i)$$
 ?
$$\alpha_t(i) \beta_t(i)$$

$$= P(\mathbf{o}_1, \mathbf{o}_2, ..., \mathbf{o}_t, s_t = i | \lambda) \cdot P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, ..., \mathbf{o}_T | s_t = i, \lambda)$$

$$= P(\mathbf{o}_1, \mathbf{o}_2, ..., \mathbf{o}_t | s_t = i, \lambda) P(s_t = i | \lambda) P(\mathbf{o}_{t+1}, \mathbf{o}_{t+2}, ..., \mathbf{o}_T | s_t = i, \lambda)$$

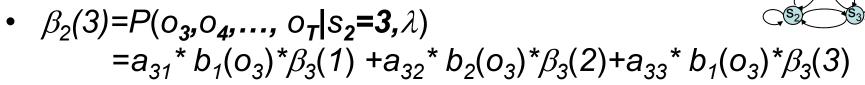
$$= P(\mathbf{o}_1, ..., \mathbf{o}_t, ..., \mathbf{o}_T | s_t = i, \lambda) P(s_t = i | \lambda)$$

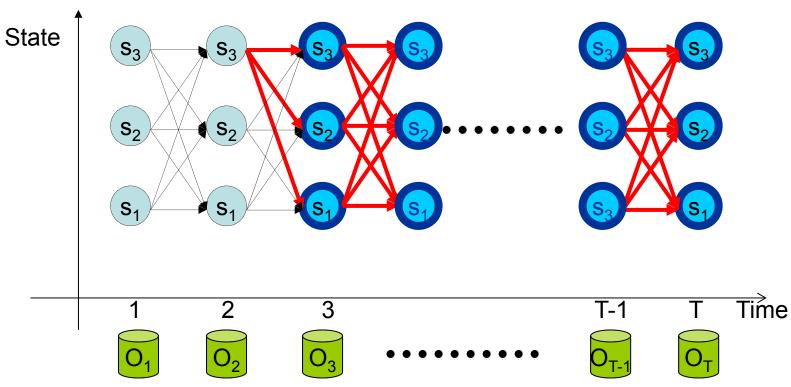
$$= P(\mathbf{o}_1, ..., \mathbf{o}_t, ..., \mathbf{o}_T, s_t = i | \lambda)$$

$$= P(\mathbf{O}, s_t = i | \lambda)$$

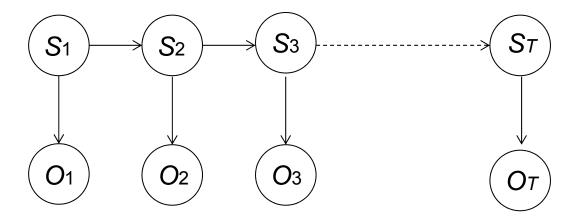
$$= P(\mathbf{O}, s_t = i | \lambda)$$
• $P(\mathbf{O} | \lambda) = \sum_{i=1}^{N} P(\mathbf{O}, s_t = i | \lambda) = \sum_{i=1}^{N} \alpha_t(i) \beta_t(i)$

- The Backward Procedure (cont.)





HMM is a Kind of Bayesian Network



How to choose an optimal state sequence $S=(s_1, s_2, \ldots, s_T)$?

 The first optimal criterion: Choose the states s_t are individually most likely at each time t

Define a posteriori probability variable $\gamma_t(i) = P(s_t = i | \mathbf{O}, \lambda)$

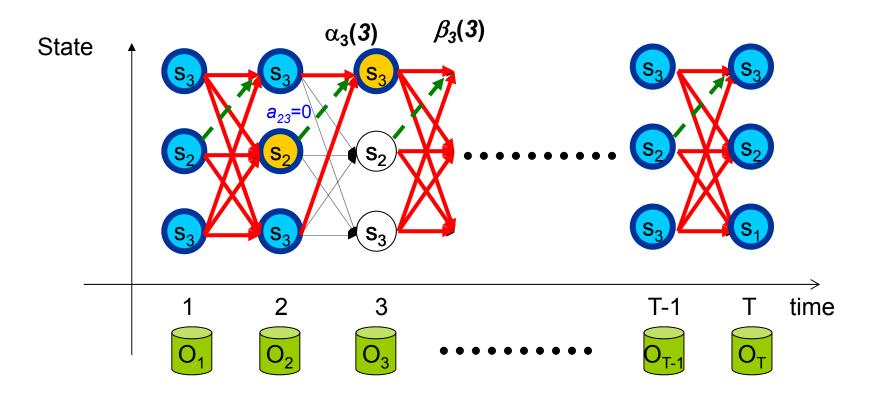
$$\gamma_{t}(i) = \frac{P(s_{t} = i, \mathbf{O}|\lambda)}{P(\mathbf{O}|\lambda)} = \frac{P(s_{t} = i, \mathbf{O}|\lambda)}{\sum_{m=1}^{N} P(s_{t} = m, \mathbf{O}|\lambda)} = \frac{\alpha_{t}(i) \beta_{t}(i)}{\sum_{m=1}^{N} \alpha_{t}(m) \beta_{t}(m)}$$

state occupation probability (count) – a soft alignment of HMM state to the observation (feature)

- Solution : $s_t^* = arg_i max [\gamma_t(i)], 1 \le t \le T$
 - Problem: maximizing the probability at each time t individually $S^* = s_1^* s_2^* ... s_T^*$ may not be a valid sequence (e.g. $a_{s_t^* s_{t+1}^*} = 0$)

\$₁ \$₂ \$₃

• $P(s_3 = 3, \mathbf{O} \mid \lambda) = \alpha_3(3) * \beta_3(3)$



- The Viterbi Algorithm
- The second optimal criterion: The Viterbi algorithm can be regarded as the dynamic programming algorithm applied to the HMM or as a modified forward algorithm
 - Instead of summing up probabilities from different paths coming to the same destination state, the Viterbi algorithm picks and remembers the best path
 - Find a single optimal state sequence $S=(s_1, s_2, \ldots, s_T)$
 - How to find the second, third, etc., optimal state sequences (difficult?)
 - The Viterbi algorithm also can be illustrated in a trellis framework similar to the one for the forward algorithm
 - State-time trellis diagram

- The Viterbi Algorithm (cont.)

Algorithm

Find a best state sequence $S = (s_1, s_2, ..., s_T)$ for a given observation $O = (o_1, o_2, ..., o_T)$?

Define a new variable

$$\delta_{t}(i) = \max_{S_{1}, S_{2}, ..., S_{t-1}} P[s_{1}, s_{2}, ..., s_{t-1}, s_{t} = i, \boldsymbol{o}_{1}, \boldsymbol{o}_{2}, ..., \boldsymbol{o}_{t} | \lambda]$$

= the best score along a single path at time *t*, which accounts for the first *t* observation and ends in state *i*

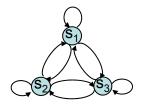
By induction
$$\therefore \delta_{t+1}(j) = \left[\max_{1 \le i \le N} \delta_t(i) a_{ij}\right] b_j(\mathbf{o}_{t+1})$$

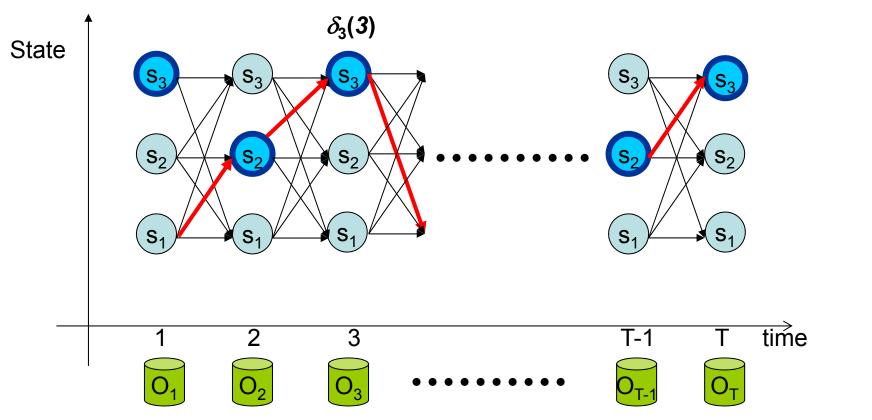
$$\psi_{t+1}(j) = \arg\max_{1 \le i \le N} \delta_t(i) a_{ij} \quad \text{ For backtracing}$$

We can backtrace from $s_T^* = \arg \max_{1 \le i \le N} \delta_T(i)$

- Complexity: $O(N^2T)$

- The Viterbi Algorithm (cont.)





- The Viterbi Algorithm (cont.)

 A three-state Hidden Markov Model for the Dow Jones Industrial average

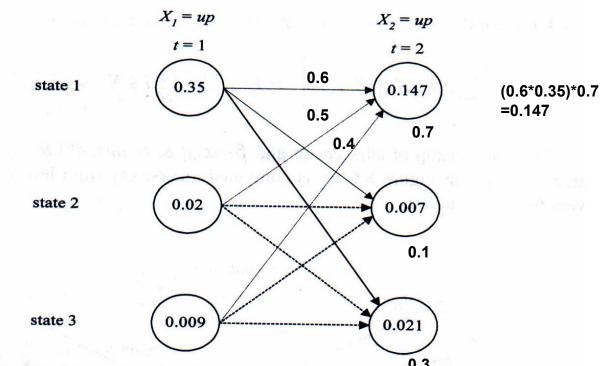


Figure 8.5 The Viterbi trellis computation for the HMM of the Dow Jones Industrial average.

- The Viterbi Algorithm (cont.)

Algorithm in the logarithmic form

Find a best state sequence $S = (s_1, s_2, ..., s_T)$ for a given observation $O = (o_1, o_2, ..., o_T)$?

Define a new variable

$$\delta_{t}(i) = \max_{s_{1}, s_{2}, ..., s_{t-1}} \log P[s_{1}, s_{2}, ..., s_{t-1}, s_{t} = i, \boldsymbol{o}_{1}, \boldsymbol{o}_{2}, ..., \boldsymbol{o}_{t} | \lambda]$$

= the best score along a single path at time *t*, which accounts for the first *t* observation and ends in state *i*

By induction :
$$\delta_{t+1}(j) = \left[\max_{1 \le i \le N} \left(\delta_t(i) + \log a_{ij}\right)\right] + \log b_j(\mathbf{o}_{t+1})$$

$$\psi_{t+1}(j) = \arg\max_{1 \le i \le N} \left(\delta_t(i) + \log a_{ij}\right) \dots \text{ For backtracing}$$

We can backtrace from $s_T^* = \arg \max_{1 \le i \le N} \delta_T(i)$

Homework 1

 A three-state Hidden Markov Model for the Dow Jones Industrial average

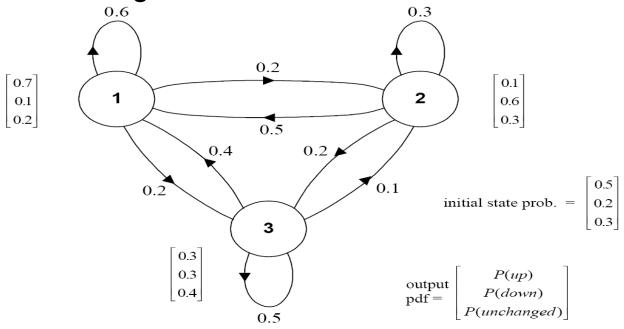


Figure 8.2 A hidden Markov model for the Dow Jones Industrial average. The three states no longer have deterministic meanings as the Markov chain illustrated in Figure 8.1.

- Find the probability: $P(up, up, unchanged, down, unchanged, down, up | \lambda)$
- Fnd the optimal state sequence of the model which generates the observation sequence: (up, up, unchanged, down, unchanged, down, up)

Probability Addition in F-B Algorithm

• In Forward-backward algorithm, operations usually implemented in logarithmic domain

 $\begin{array}{c|c}
P_1 & P_1 + P_2 \\
\log P_2 & \log(P_1 + P_2)
\end{array}$

• Assume that we want to add P_1 and P_2

if
$$P_1 \ge P_2$$

 $\log_b(P_1 + P_2) = \log_b P_1 + \log_b(1 + b^{\log_b P_2 - \log_b P_1})$
else
 $\log_b(P_1 + P_2) = \log_b P_2 + \log_b(1 + b^{\log_b P_1 - \log_b P_2})$

The values of $\log_b (1+b^x)$ can be saved in in a table to speedup the operations

Probability Addition in F-B Algorithm (cont.)

An example code

```
#define LZERO (-1.0E10) // ~log(0)
#define LSMALL (-0.5E10) // log values < LSMALL are set to LZERO
#define minLogExp -log(-LZERO) // ~=-23
double LogAdd(double x, double y)
double temp, diff, z;
 if (x<y)
   temp = x; x = y; y = temp;
 diff = y-x; //notice that diff <= 0
 if (diff<minLogExp) // if y' is far smaller than x'
   return (x<LSMALL)? LZERO:x;
  else
   z = \exp(diff);
   return x + \log(1.0 + z);
```

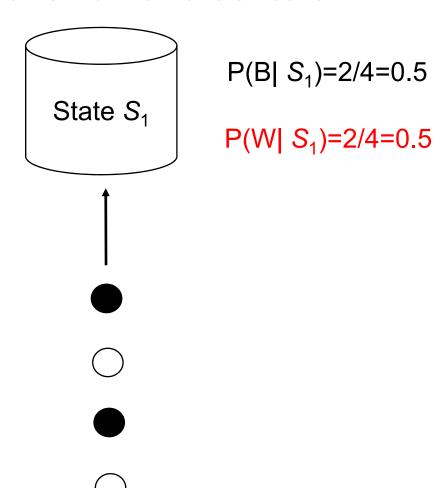
- How to adjust (re-estimate) the model parameter λ=(A,B,π) to maximize P(O₁,..., O_L|λ) or logP(O₁,..., O_L|λ)?
 - Belonging to a typical problem of "inferential statistics"
 - The most difficult of the three problems, because there is no known analytical method that maximizes the joint probability of the training data in a close form

data in a close form
$$\log P(\mathbf{O}_1, \mathbf{O}_2, ..., \mathbf{O}_L | \lambda) = \log \prod_{l=1}^L P(\mathbf{O}_l | \lambda)$$
 The "log of sum" form is difficult to deal with

- -Suppose that we have L training utterances for the HMM
- -S: a possible state sequence of the HMM
- The data is incomplete because of the hidden state sequences
- Well-solved by the Baum-Welch (known as forward-backward)
 algorithm and EM (Expectation-Maximization) algorithm
 - Iterative update and improvement
 - Based on Maximum Likelihood (ML) criterion

Maximum Likelihood (ML) Estimation: A Schematic Depiction (1/2)

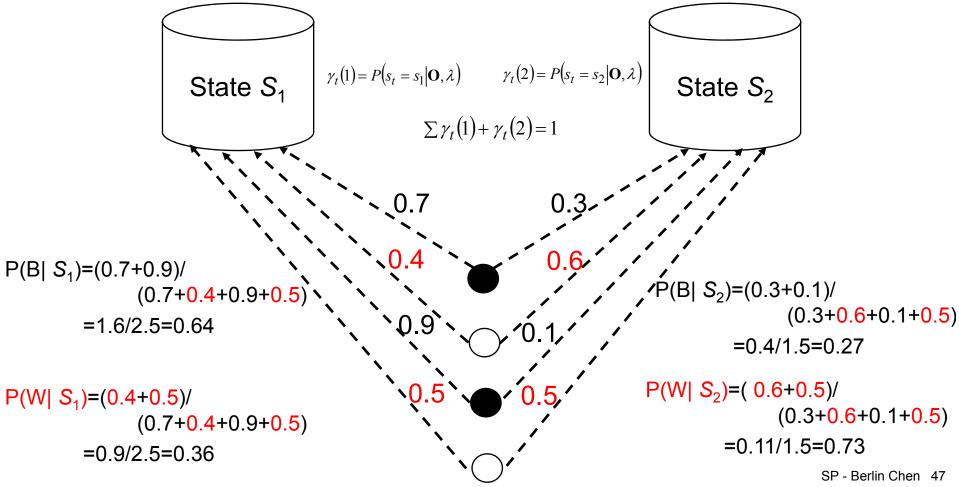
- Hard Assignment
 - Given the data follow a multinomial distribution



Maximum Likelihood (ML) Estimation: A Schematic Depiction (1/2)

- Soft Assignment
 - Given the data follow a multinomial distribution

Maximize the likelihood of the data given the alignment



Relationship between the forward and backward variables

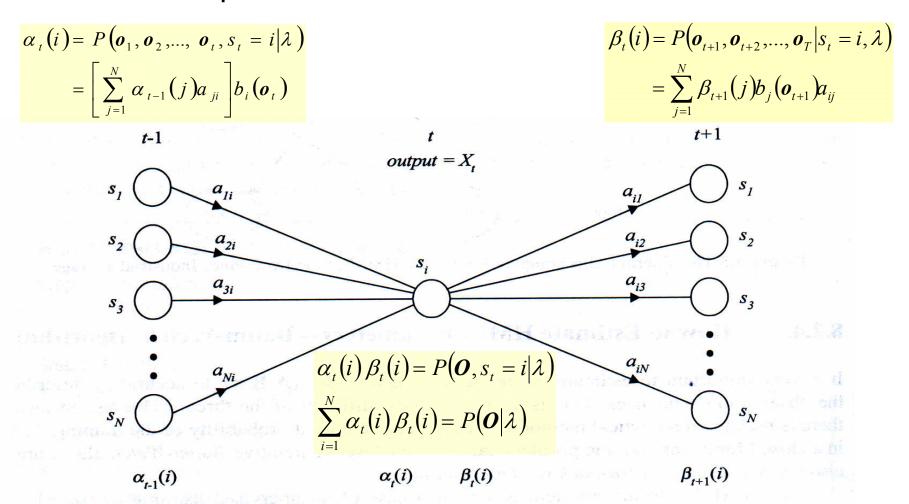
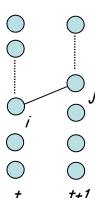


Figure 8.6 The relationship of α_{t-1} and α_t and β_t and β_{t+1} in the forward-backward algorithm.



Define a new variable:

$$\xi_t(i,j) = P(s_t = i, s_{t+1} = j | \mathbf{O}, \lambda)$$

Probability being at state i at time t and at state j at time t+1

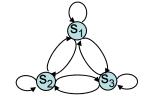
$$\xi_{t}(i,j) = \frac{P(s_{t} = i, s_{t+1} = j, \mathbf{O}|\lambda)}{P(\mathbf{O}|\lambda)}$$

$$= \frac{\alpha_{t}(i)a_{ij}b_{j}(\mathbf{o}_{t+1})\beta_{t+1}(j)}{P(\mathbf{O}|\lambda)} = \frac{\alpha_{t}(i)a_{ij}b_{j}(\mathbf{o}_{t+1})\beta_{t+1}(j)}{\sum_{m=1}^{N}\sum_{n=1}^{N}\alpha_{t}(m)a_{mn}b_{n}(\mathbf{o}_{t+1})\beta_{t+1}(n)}$$

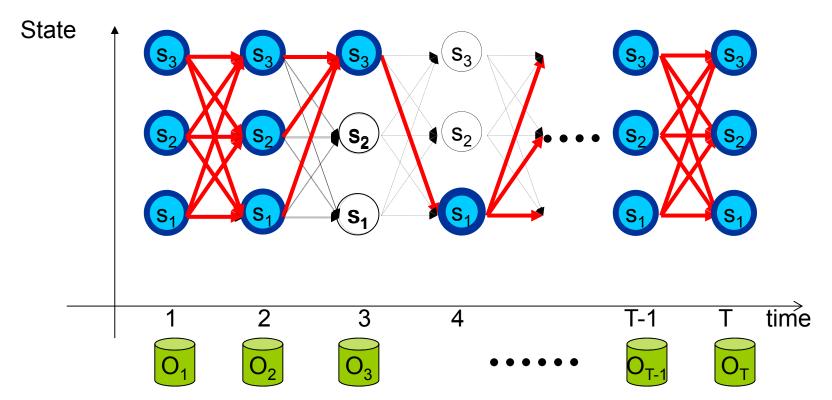
Recall the posteriori probability variable:

$$\gamma_{t}(i) = P(s_{t} = i | \mathbf{O}, \lambda)$$
Note: $\gamma_{t}(i)$ also can be represented as
$$\frac{\alpha_{t}(i)\beta_{t}(i)}{\sum_{m=1}^{N} \alpha_{t}(m)\beta_{t}(m)}$$

$$\gamma_t(i) = \sum_{j=1}^{N} P(s_t = i, s_{t+1} = j | \mathbf{O}, \lambda) = \sum_{j=1}^{N} \xi_t(i, j)$$
 (for $t < T$)



• $P(s_3 = 3, s_4 = 1, \mathbf{O} \mid \lambda) = \alpha_3(3)^* a_{31}^* b_1(o_4)^* \beta_1(4)$



- $\xi_t(i,j) = P(s_t = i, s_{t+1} = j | \mathbf{O}, \lambda)$ $\sum_{t=1}^{T-1} \xi_t(i,j) = \text{expected number of transitions from state } i \text{ to state } j \text{ in } \mathbf{O}$
- $\gamma_t(i) = P(s_t = i | \mathbf{0}, \lambda)$ $\sum_{t=1}^{T-1} \gamma_t(i) = \sum_{t=1}^{T-1} \sum_{i=1}^{N} \xi_t(i, j) = \text{expected number of transitions from state } i \text{ in } \mathbf{0}$
- A set of reasonable re-estimation formula for $\{A, \pi\}$ is

 $\overline{\pi}_i$ = expected frequency (number of times) in state i at time t = 1= $\gamma_1(i)$

$$\overline{a}_{ij} = \frac{\text{expected number of transitio n from state } i \text{ to state } j}{\text{expected number of transitio n from state } i} = \frac{\sum\limits_{t=1}^{T-1} \zeta_t(i,j)}{\sum\limits_{t=1}^{T-1} \gamma_t(i)}$$

Formulae for Single Training Utterance

- A set of reasonable re-estimation formula for {B} is
 - For discrete and finite observation $b_i(\mathbf{v}_k) = P(\mathbf{o}_t = \mathbf{v}_k | \mathbf{s}_t = j)$

$$\overline{b}_{j}(\mathbf{v}_{k}) = \overline{P}(\mathbf{o} = \mathbf{v}_{k} | s = j) = \frac{\text{expected number of times in state } j \text{ and observing symbol } \mathbf{v}_{k}}{\text{expected number of times in state } j} = \frac{\sum_{t=1}^{T} \gamma_{t}(j)}{\sum_{t=1}^{T} \gamma_{t}(j)}$$

- For continuous and infinite observation $b_i(\mathbf{v}) = f_{O|S}(\mathbf{o}_t = \mathbf{v}|\mathbf{s}_t = j)$,

$$\overline{b}_{j}(\mathbf{v}) = \sum_{k=1}^{M} \overline{c}_{jk} N(\mathbf{v}; \overline{\boldsymbol{\mu}}_{jk}, \overline{\boldsymbol{\Sigma}}_{jk}) = \sum_{k=1}^{M} \overline{c}_{jk} \left(\frac{1}{(\sqrt{2\pi})^{L} |\overline{\boldsymbol{\Sigma}}_{jk}|^{1/2}} \exp\left(-\frac{1}{2} (\mathbf{v} - \overline{\boldsymbol{\mu}}_{jk})^{t} \overline{\boldsymbol{\Sigma}}_{jk}^{-1} (\mathbf{v} - \overline{\boldsymbol{\mu}}_{jk})\right) \right)$$

Modeled as a mixture of multivariate Gaussian distributions

$$p(A|B) = \frac{p(A,B)}{P(B)}$$

- For continuous and infinite observation (Cont.)
 - Define a new variable $\gamma_t(j,k)$
 - $-\gamma_t(j,k)$ is the probability of being in state j at time t with the k-th mixture component accounting for \mathbf{o}_t

$$\gamma_{t}(j,k) = P(s_{t} = j, m_{t} = k | \mathbf{O}, \lambda)$$

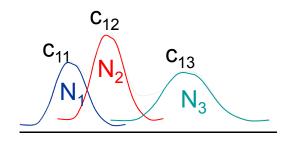
$$= P(s_{t} = j | \mathbf{O}, \lambda) P(m_{t} = k | s_{t} = j, \mathbf{O}, \lambda)$$

$$= \gamma_{t}(j) P(m_{t} = k | s_{t} = j, \mathbf{O}, \lambda)$$

$$= \gamma_{t}(j) \frac{p(m_{t} = k, \mathbf{O} | s_{t} = j, \lambda)}{p(\mathbf{O} | s_{t} = j, \lambda)}$$

$$= \gamma_{t}(j) \frac{P(m_{t} = k | s_{t} = j, \lambda) p(\mathbf{O} | s_{t} = j, m_{t} = k, \lambda)}{p(\mathbf{O} | s_{t} = j, \lambda)}$$
..... (observation - independent assumption is applied)
$$= \gamma_{t}(j) \frac{P(m_{t} = k | s_{t} = j, \lambda) p(\mathbf{o}_{t} | s_{t} = j, m_{t} = k, \lambda)}{p(\mathbf{o}_{t} | s_{t} = j, \lambda)}$$

$$= \left[\frac{\alpha_{t}(j)\beta_{t}(j)}{\sum_{i}^{N} \alpha_{t}(s)\beta_{t}(s)}\right] \frac{c_{jk}N(\mathbf{o}_{t}; \mathbf{\mu}_{jk}, \mathbf{\Sigma}_{jk})}{\sum_{i}^{N} c_{jm}N(\mathbf{o}_{t}; \mathbf{\mu}_{jm}, \mathbf{\Sigma}_{jm})}$$



Distribution for State 1

Note:
$$\gamma_t(j) = \sum_{m=1}^{M} \gamma_t(j, m)$$

For continuous and infinite observation (Cont.)

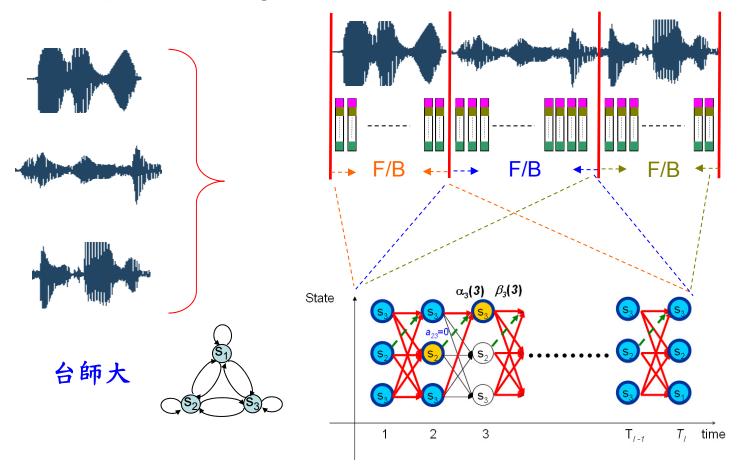
$$\overline{c}_{jk} = \frac{\text{expected number of times in state } j \text{ and mixture } k}{\text{expected number of times in state } j} = \frac{\sum\limits_{t=1}^{T} \gamma_t(j,k)}{\sum\limits_{t=1}^{T} \sum\limits_{m=1}^{M} \gamma_t(j,m)}$$

 $\overline{\mu}_{jk}$ = weighted average (mean) of observations at state j and mixture $k = \frac{\sum_{t=1}^{T} \gamma_t(j,k) \cdot \boldsymbol{o}_t}{\sum_{t=1}^{T} \gamma_t(j,k)}$

$$\overline{\Sigma}_{jk} = \text{weighted covariance of observations at state } j \text{ and mixture } k$$

$$= \frac{\sum_{t=1}^{T} \gamma_{t}(j,k) \cdot (\boldsymbol{o}_{t} - \overline{\boldsymbol{\mu}}_{jk}) (\boldsymbol{o}_{t} - \overline{\boldsymbol{\mu}}_{jk})^{t}}{\sum_{t=1}^{T} \gamma_{t}(j,k)}$$

Multiple Training Utterances



For continuous and infinite observation (Cont.)

 $\overline{\pi}_i$ = expected frequency (number of times) in state *i* at time $(t=1) = \frac{1}{L} \sum_{l=1}^{L} \gamma_1^l(i)$

$$\overline{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state } i} = \frac{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_l-1} \zeta_t^l(i,j)}{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_l-1} \gamma_t^l(i)}$$

$$\overline{c}_{jk} = \frac{\text{expected number of times in state } j \text{ and mixture } k}{\text{expected number of times in state } j} = \frac{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_l}\gamma_t^l\left(j,k\right)}{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_l}\sum\limits_{m=1}^{M}\gamma_t^l\left(j,m\right)}$$

$$\overline{\boldsymbol{\mu}}_{jk} = \text{weighted average (mean) of observations at state } j \text{ and mixture } k = \frac{\sum\limits_{l=1}^{L} \sum\limits_{t=1}^{T_l} \gamma_t^l(j,k) \cdot \boldsymbol{o}_t}{\sum\limits_{l=1}^{L} \sum\limits_{t=1}^{T_l} \gamma_t^l(j,k)}$$

$$\begin{split} \overline{\boldsymbol{\Sigma}}_{jk} &= \text{weighted covariance of observations at state } \boldsymbol{j} \text{ and mixture } \boldsymbol{k} \\ &= \frac{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_{l}}\gamma_{t}^{l}(j,k)\cdot\left(\mathbf{o}_{t} - \overline{\boldsymbol{\mu}}_{jk}\right)\!\left(\mathbf{o}_{t} - \overline{\boldsymbol{\mu}}_{jk}\right)^{t}}{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_{l}}\gamma_{t}^{l}(j,k)} \end{split}$$

For discrete and finite observation (cont.)

$$\overline{\pi}_i$$
 = expected frequency (number of times) in state *i* at time $(t = 1) = \frac{1}{L} \sum_{l=1}^{L} \gamma_1^l(i)$

$$\overline{a}_{ij} = \frac{\text{expected number of transition from state } i \text{ to state } j}{\text{expected number of transition from state} i} = \frac{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_l-1}\xi_t^l(i,j)}{\sum\limits_{l=1}^{L}\sum\limits_{t=1}^{T_l-1}\gamma_t^l(i)}$$

$$\overline{b}_{j}(\mathbf{v}_{k}) = \overline{P}(\mathbf{o} = \mathbf{v}_{k} | s = j) = \frac{\text{expected number of times in state } j \text{ and observing symbol } \mathbf{v}_{k}}{\text{expected number of times in state } j} = \frac{\sum\limits_{l=1}^{L} \sum\limits_{t=1}^{T_{l}} \gamma_{t}^{l}(j)}{\sum\limits_{l=1}^{L} \sum\limits_{t=1}^{T_{l}} \gamma_{t}^{l}(j)}$$

Formulae for Multiple (L) Training Utterances

Semicontinuous HMMs

- The HMM state mixture density functions are tied together across all the models to form a set of shared kernels
 - The semicontinuous or tied-mixture HMM

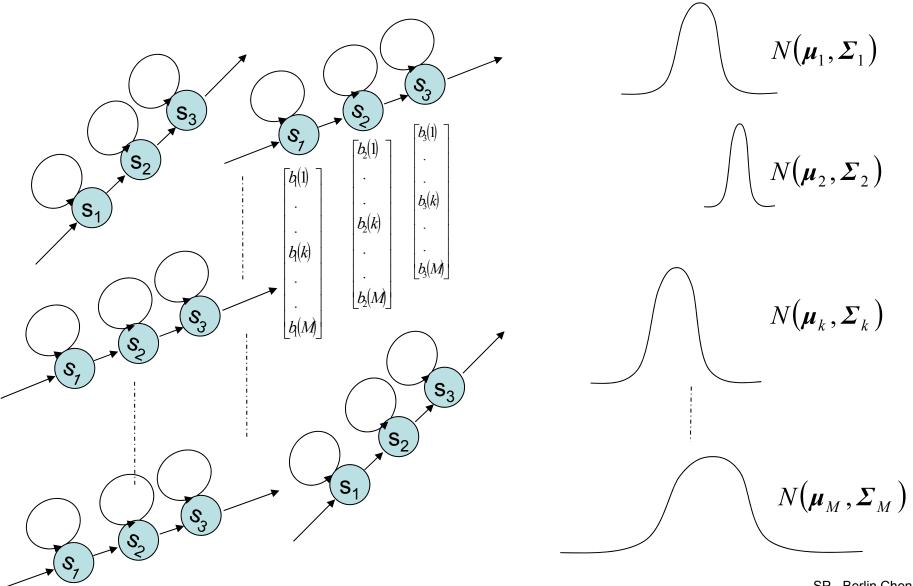
$$b_{j}(\boldsymbol{o}) = \sum_{k=1}^{M} b_{j}(k) f(\boldsymbol{o}|v_{k}) = \sum_{k=1}^{M} b_{j}(k) N(\boldsymbol{o}, \boldsymbol{\mu}_{k}, \boldsymbol{\Sigma}_{k})$$

state output Probability of state *j*

k-th mixture weight k-th mixture density function or k-th codeword t of state j (shared across HMMs, M is very large)
 (discrete, model-dependent)

- A combination of the discrete HMM and the continuous HMM
 - A combination of discrete model-dependent weight coefficients and continuous model-independent codebook probability density functions
- Because M is large, we can simply use the L most significant values $f(o|v_k)$
 - Experience showed that L is $1\sim3\%$ of M is adequate
- Partial tying of $f(o|v_k)$ for different phonetic class

Semicontinuous HMMs (cont.)



HMM Topology

- Speech is time-evolving non-stationary signal
 - Each HMM state has the ability to capture some quasi-stationary segment in the non-stationary speech signal
 - A *left-to-right* topology is a natural candidate to model the speech signal (also called the "beads-on-a-string" model)

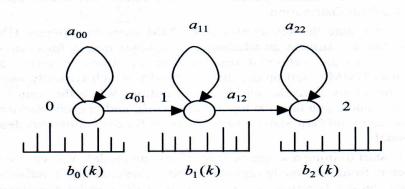


Figure 8.8 A typical hidden Markov model used to model phonemes. There are three states (0-2) and each state has an associated output probability distribution.

 It is general to represent a phone using 3~5 states (English) and a syllable using 6~8 states (Mandarin Chinese)

Initialization of HMM

-S₁-S₂-S₃-

- A good initialization of HMM training : Segmental K-Means Segmentation into States
 - Assume that we have a training set of observations and an initial estimate of all model parameters
 - Step 1 : The set of training observation sequences is segmented into states, based on the initial model (finding the optimal state sequence by Viterbi Algorithm)
 - Step 2:
 - For discrete density HMM (using M-codeword codebook)

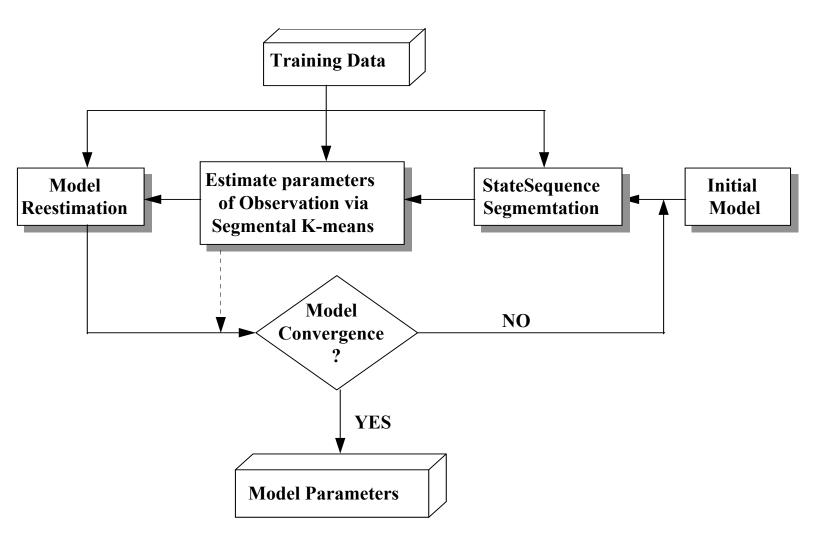
```
\overline{b}_{j}(k) = \frac{\text{the number of vectors with codebook index } k \text{ in state } j}{\text{the number of vectors in state } j}
```

• For continuous density HMM (M Gaussian mixtures per state)

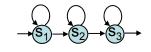
```
\Rightarrow cluster the observation vectors within each state j into a set of M clusters \overline{w}_{jm} = number of vectors classified in cluster m of state j divided by the number of vectors in state j \overline{\mu}_{jm} = sample mean of the vectors classified in cluster m of state j \overline{\Sigma}_{jm} = sample covariance matrix of the vectors classified in cluster m of state j
```

Step 3: Evaluate the model score
 If the difference between the previous and current model scores is greater than a threshold, go back to Step 1, otherwise stop, the initial model is generated

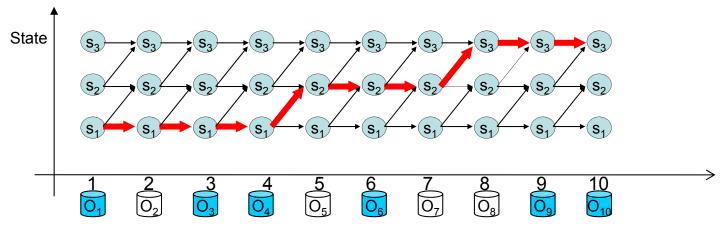
Initialization of HMM (cont.)



Initialization of HMM (cont.)



- An example for discrete HMM
 - 3 states and 2 codeword



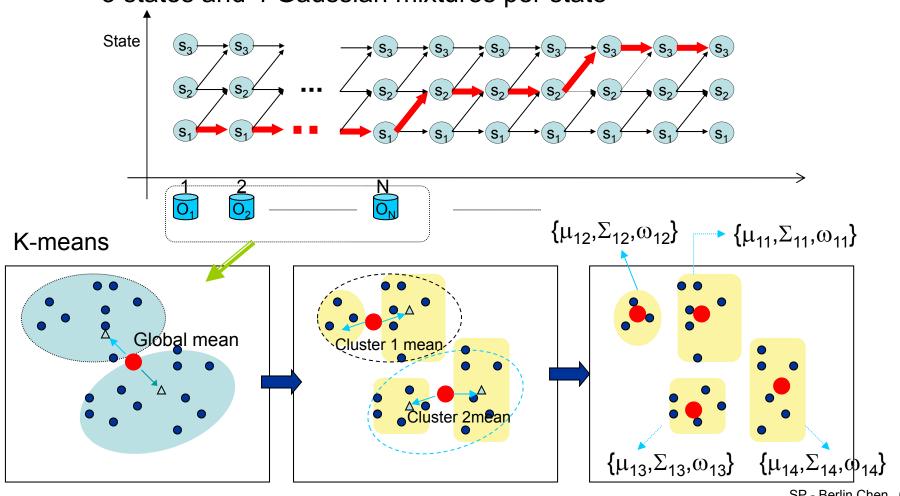
- $b_1(\mathbf{v}_1)=3/4$, $b_1(\mathbf{v}_2)=1/4$
- $b_2(\mathbf{v}_1)=1/3$, $b_2(\mathbf{v}_2)=2/3$
- $b_3(\mathbf{v}_1)=2/3$, $b_3(\mathbf{v}_2)=1/3$



$$\mathbf{v}_2$$

Initialization of HMM (cont.)

- An example for Continuous HMM
 - 3 states and 4 Gaussian mixtures per state



Known Limitations of HMMs (1/3)

- The assumptions of conventional HMMs in Speech Processing
 - The state duration follows an exponential distribution
 - Don't provide adequate representation of the temporal structure of speech

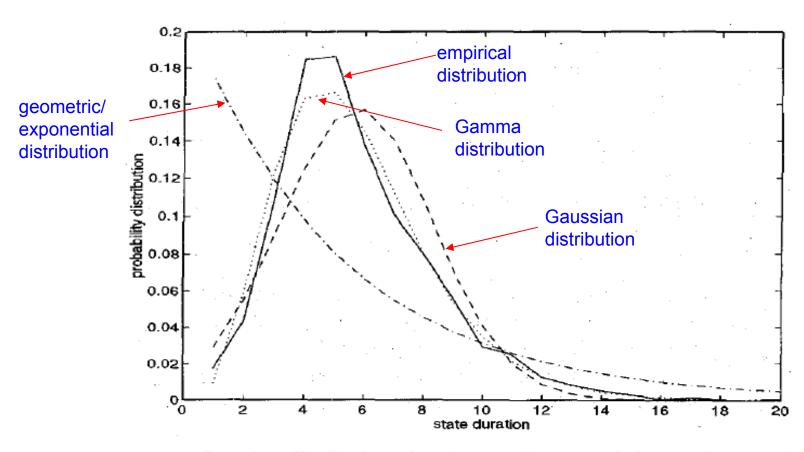
 $d_i(t) = a_{ii}^{t-1} \left(1 - a_{ii}\right)$

- First-order (Markov) assumption: the state transition depends only on the origin and destination
- Output-independent assumption: all observation frames are dependent on the state that generated them, not on neighboring observation frames

Researchers have proposed a number of techniques to address these limitations, albeit these solution have not significantly improved speech recognition accuracy for practical applications.

Known Limitations of HMMs (2/3)

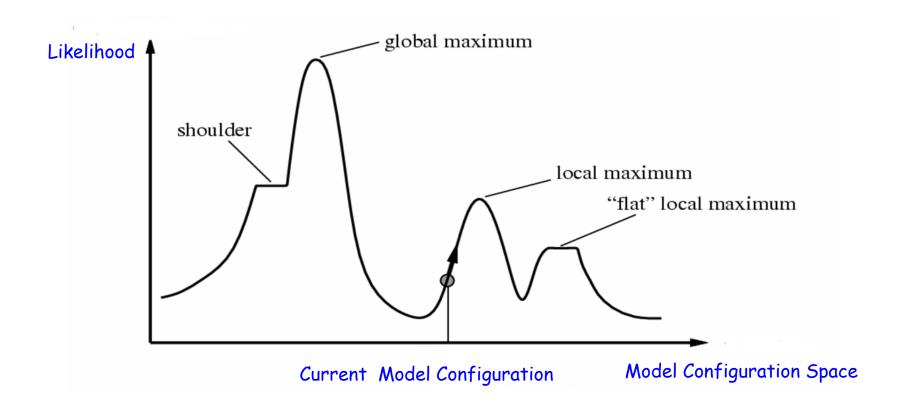
Duration modeling



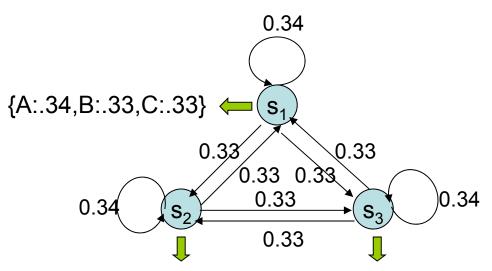
Duration distributions for the seventh state of the word "seven:" empirical distribution (solid line); Gauss fit (dashed line); gamma fit (dotted line); and (d) geometric fit (dash-dot line).

Known Limitations of HMMs (3/3)

 The HMM parameters trained by the Baum-Welch algorithm (or EM algorithm) were only locally optimized



Homework-2 (1/2)



{A:.33,B:.34,C:.33} {A:.33,B:.33,C:.34}

TrainSet 1:

- 1. ABBCABCAABC
- 2. ABCABC
- 3. ABCA ABC
- 4. BBABCAB
- 5. BCAABCCAB
- 6. CACCABCA
- 7. CABCABCA
- 8. CABCA
- 9. CABCA

TrainSet 2:

- 1. BBBCCBC
- 2. CCBABB
- 3. AACCBBB
- 4. BBABBAC
- 5. CCA ABBAB
- 6. BBBCCBAA
- 7. ABBBBABA
- 8. CCCCC
- 9. BBAAA

Homework-2 (2/2)

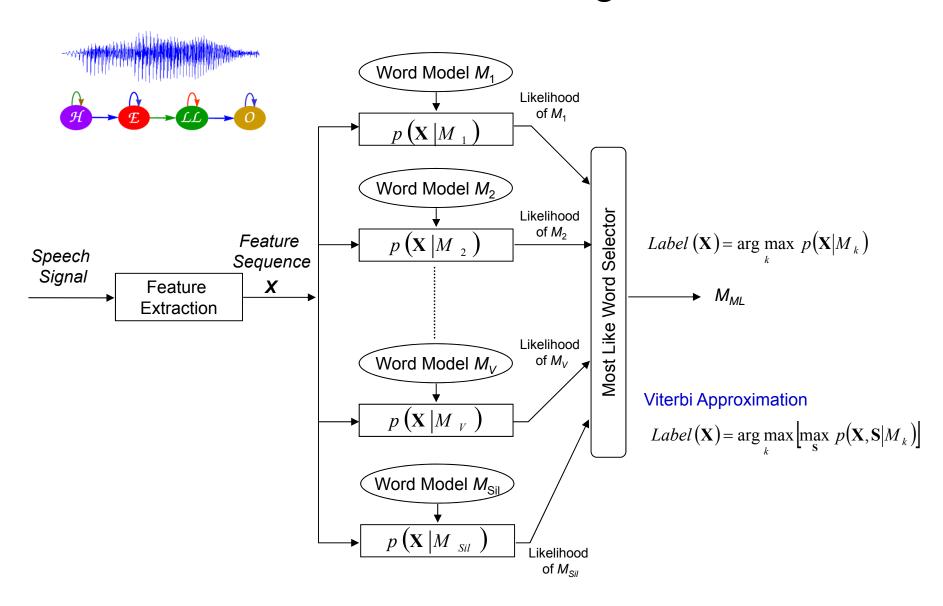
- P1. Please specify the model parameters after the first and 50th iterations of Baum-Welch training
- P2. Please show the recognition results by using the above training sequences as the testing data (The so-called inside testing).

 *You have to perform the recognition task with the HMMs trained from the first and 50th iterations of Baum-Welch training, respectively
- P3. Which class do the following testing sequences belong to?

 ABCABCCAB

 AABABCCCCBBB
- P4. What are the results if Observable Markov Models were instead used in P1, P2 and P3?

Isolated Word Recognition



Measures of ASR Performance (1/8)

- Evaluating the performance of automatic speech recognition (ASR) systems is critical, and the Word Recognition Error Rate (WER) is one of the most important measures
- There are typically three types of word recognition errors
 - Substitution
 - An incorrect word was substituted for the correct word
 - Deletion
 - A correct word was omitted in the recognized sentence
 - Insertion
 - An extra word was added in the recognized sentence
- How to determine the minimum error rate?

Measures of ASR Performance (2/8)

- Calculate the WER by aligning the correct word string against the recognized word string
 - A maximum substring matching problem
 - Can be handled by dynamic programming deleted
- Example:

WER+

=100%

WAR

Correct: "the effect is clear"

Recognized: "effect is not clear"

matched inserted matched

- Error analysis: one deletion and one insertion
- Measures: word error rate (WER), word correction rate (WCR),
 word accuracy rate (WAR)
 Might be higher than 100%

Word Error Rate =
$$100\%$$
 $\frac{\text{Sub.} + \text{Del.} + \text{Ins. words}}{\text{No. of words in the correct sentence}} = \frac{2}{4} = 50\%$

Word Correction Rate = 100% Matched words words in the correct sentence = $\frac{3}{4}$ = 75%

Word Accuracy Rate = 100% $\frac{\text{Matched - Ins. words}}{\text{No. of words in the correct sentence}} = \frac{3-1}{4} = 50\%$

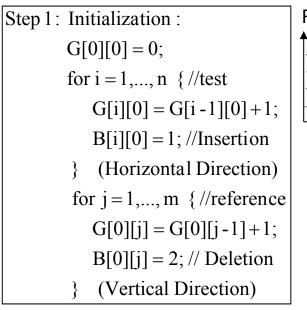
Measures of ASR Performance (3/8)

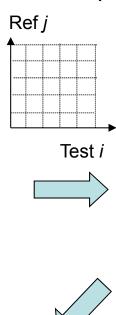
A Dynamic Programming Algorithm (Textbook)

```
ALGORITHM 9.1: ALGORITHM TO MEASURE THE WORD ERROR RATE
             Step 1: Initialization R[0,0] = 0 R[i, j] = \infty if (i < 0) or (j < 0) B[0,0] = 0
             Step 2: Iteration
             for i = 1, ..., n { //denotes for the word length of the correct/reference sentence
                 for j = 1, ..., m { //denotes for the word length of the recognized/test sentence
                                             R[i-1, j]+1 (deletion)
    minimum word
                                         R[i-1, j-1] (match)/hit
    error alignment
                                        R[i-1, j-1]+1 (substitution)
    at the a grid [i,i]
                        R[i, j] = \min R[i, j-1]+1 (insertion)
                                 1 if deletion
kinds of
alignment
                                    if substitution
                                                                                                             Ref i
            Step 3: Backtracking and termination
                   word error rate = 100\% \times \frac{R(n,m)}{}
                   optimal backward path = (s_1, s_2, ..., 0)
                   where s_1 = B[n,m], s_t = \begin{bmatrix} B[i-1,j] \text{ if } s_{t-1} = 1 \\ B[i,j-1] \text{ if } s_{t-1} = 2 \\ B[i-1,j-1] \text{ if } s_{t-1} = 3 \text{ or } 4 \end{bmatrix} for t = 2,... until s_t = 0
```

Measures of ASR Performance (4/8)

Algorithm (by Berlin Chen)





```
Step 2: Iteration: \\ for i = 1,...,n ~ {//test} \\ for j = 1,...,m ~ {//reference} \\ G[i][j] = min \begin{bmatrix} G[i-1][j]+1 & (Insertion) \\ G[i][j-1]+1 & (Delection) \\ G[i-1][j-1]+1 & (if LR[j]!=LT[i], Substitution) \\ G[i-1][j-1] & (if LR[j]=LT[i], Match) \end{bmatrix} \\ B[i][j] = \begin{cases} 1; //Insertion, & (Horizontal Direction) \\ 2; //Deletion, & (Vertical Direction) \\ 3; //Substitution & (Diagonal Direction) \\ 4; //match & (Diagonal Direction) \end{cases} \\ \} //for j, reference \\ \} //for i, test
```

```
Step 3: Measure and Backtrace : Word Error Rate = 100\% \times \frac{G[n][m]}{m} Word Accuracy Rate = 100\% - W Word Error Rate Optimal backtrace path = (B[n][m] \rightarrow ..... \rightarrow B[0][0]) if B[i][j] = 1 print " LT[i]";//Insertion, then go left else if B[i][j] = 2 print "LR[j] ";//Deletion, then go down else print "LR[j] LR[i]";//Hit/Matc h or Substituti on, then go down diagonally
```

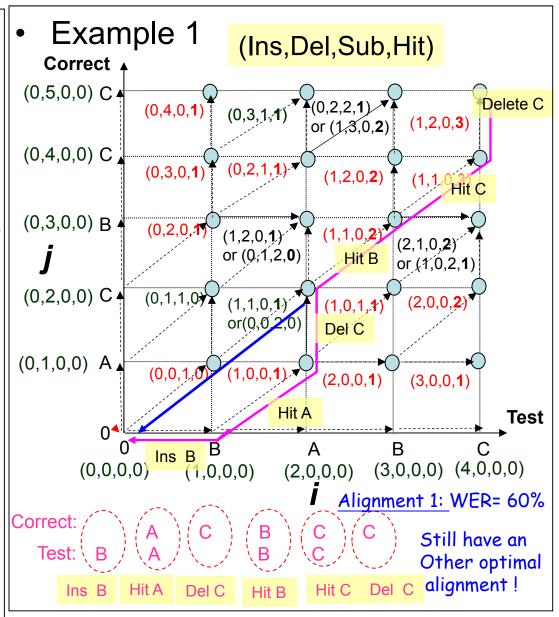
Measures of ASR Performance (5/8)

A Dynamic Programming Algorithm

 Initialization Ins. (n,m)m Correct/Reference Word Del. Sequence *m*-1 for (j=1;j<=m;j++){ //reference Ins. (*i,j*) grid[0][j] = grid[0][j-1];(*i*-1,*i*) grid[0][j].dir = VERT;Del grid[0][j].score (i-1,j-1)+= DelPen; (i,j-1)grid[0][j].del ++; 4 3Del. 3 2Del. 2 Del. 1Del. 1 HTK 5 3 n-1 n 1Ins. 2Ins. 3Ins. Recognized/test Word for (i=1;i<=n;i++) { // test grid[0][0].score = grid[0][0].insSequence grid[i][0] = grid[i-1][0];= grid[0][0].del = 0;grid[i][0].dir = HOR;grid[0][0].sub = grid[0][0].hit = 0;grid[i][0].score +=InsPen; grid[0][0].dir = NIL;grid[i][0].ins ++; SP - Berlin Chen 75

Measures of ASR Performance (6/8)

```
Program
     for (i=1:i<=n:i++) //test
      { gridi = grid[i]; gridi1 = grid[i-1];
         for (j=1;j<=m;j++) //reference
               h = gridi1[j].score +insPen;
              d = gridi1[j-1].score;
              if (IRef[i] != ITest[i])
                d += subPen:
              v = gridi[j-1].score + delPen;
              if (d \le h \& \& d \le v) {/* DIAG = hit or sub */}
                gridi[j] = gridi1[j-1]; //structure assignment
HTK
                gridi[i].score = d;
                gridi[j].dir = DIAG;
                if (IRef[i] == ITest[i])
                                        ++gridi[j].hit;
                else ++aridi[i].sub;
                else if (h<v) {
                                    /* HOR = ins */
                gridi[j] = gridi1[j];
                                      //structure assignment
                gridi[j].score = h;
                gridi[j].dir = HOR;
                ++ gridi[i].ins;
              else {
                                    /* VERT = del */
                gridi[j] = gridi[j-1];
                                      //structure assignment
                gridi[j].score = v;
                gridi[j].dir = VERT;
                ++gridi[j].del; }
                     /* for i */
       } /* for i */
```

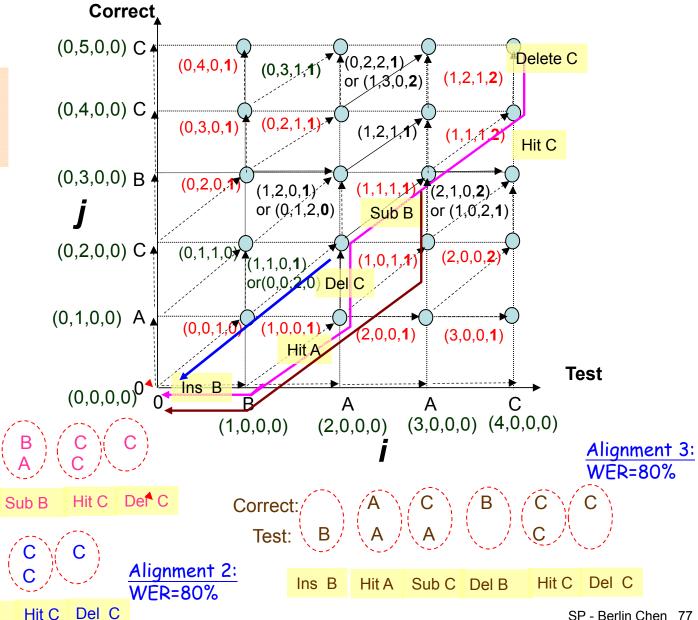


Measures of ASR Performance (7/8)



Note: the penalties for substitution, deletion and insertion errors are all set to be 1 here

(Ins,Del,Sub,Hit)



Alignment 1: WER= 80%

Correct:

Test:

Correct: C В Test:

Sub A Sub C Sub B

Measures of ASR Performance (8/8)

 Two common settings of different penalties for substitution, deletion, and insertion errors

```
/* HTK error penalties */
subPen = 10;
delPen = 7;
insPen = 7;

/* NIST error penalties*/
subPenNIST = 4;
delPenNIST = 3;
insPenNIST = 3;
```

Homework 3

Measures of ASR Performance

Reference

100000 100000 桃 100000 100000 芝 100000 100000 颱 100000 100000 風 100000 100000 重 100000 100000 創 100000 100000 花 100000 100000 蓮 100000 100000 光 100000 100000 復 100000 100000 鄉 100000 100000 大 100000 100000 興 100000 100000 村 100000 100000 死 100000 100000 傷 100000 100000 慘 100000 100000 重 100000 100000 感 100000 100000 觸 100000 100000 最 100000 100000 多

ASR Output

```
100000 100000 桃
100000 100000 芝
100000 100000 颱
100000 100000 風
100000 100000 重
100000 100000 創
100000 100000 花
100000 100000 蓮
100000 100000 光
100000 100000 復
100000 100000 鄉
100000 100000 打
100000 100000 新
100000 100000 村
100000 100000 次
100000 100000 傷
100000 100000 殘
100000 100000 周
100000 100000 感
100000 100000 觸
100000 100000 最
100000 100000 多
```

•

Homework 3

- 506 BN stories of ASR outputs
 - Report the CER (character error rate) of the first one, 100, 200, and 506 stories
 - The result should show the number of substitution, deletion and insertion errors

Overall Results
SENT: %Correct=0.00 [H=0, S=1, N=1] WORD: %Corr=81.52, Acc=81.52 [H=75, D=4, S=13, I=0, N=92]
Overall Results
SENT: %Correct=0.00 [H=0, S=100, N=100] WORD: %Corr=87.66, Acc=86.83 [H=10832, D=177, S=1348, I=102, N=12357]
Overall Results
SENT: %Correct=0.00 [H=0, S=200, N=200] WORD: %Corr=87.91, Acc=87.18 [H=22657, D=293, S=2824, I=186, N=25774]
Overall Results
SENT: %Correct=0.00 [H=0, S=506, N=506] WORD: %Corr=86.83, Acc=86.06 [H=57144, D=829, S=7839, I=504, N=65812]

Symbols for Mathematical Operations

A α alpha
B β beta
Γ γ gamma
E ε epsilon
Δ δ delta
Z ζ zeta
Η η eta
Θ θ theta

I ι iota
Κ κ kappa
na Λ λ lambda
on Μ μ mu
Ν ν nu
Ξ ξ xi
Ο ο omicron
η π pi

P ρ rho
Σ σ sigma
Τ τ tau
Y υ upsilon
Φ φ phi
Χ χ chi
Ψ ψ psi
Ω ω omega